

The  $\Lambda$ -Fleming-Viot process and a connection with  
Wright-Fisher diffusion

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A  $d$ -dimensional  $\Lambda$ -Fleming-Viot process  $\{\mathbf{X}(t)\}_{t \geq 0}$  representing frequencies of  $d$  types of individuals in a population has a generator described by

$$\mathcal{L}g(\mathbf{x}) = \int_0^1 \sum_{i=1}^d x_i \left( g(\mathbf{x}(1-y) + y\mathbf{e}_i) - g(\mathbf{x}) \right) \frac{F(dy)}{y^2}.$$

The population is **partitioned** at events of change by choosing type  $i \in [d]$  to reproduce with probability  $x_i$ , then rescaling the population with additional offspring  $y$  of type  $i$  to be  $\mathbf{x}(1-y) + y\mathbf{e}_i$  at rate  $y^{-2}F(dy)$ .

## Examples

Eldon and Wakeley (2006). A model where  $F$  has a single point of increase in  $(0, 1]$  with a possible atom at 0.

A natural class that arises from discrete models is when  $F$  has a  $\text{Beta}(\alpha, \beta)$  density, particularly a  $\text{Beta}(2 - \alpha, \alpha)$  density coming from a discrete model where the offspring distribution tails are asymptotic to a power law of index  $\alpha$ . Birkner, Blath, Capaldo, Etheridge, Möhle, Schweinsberg, Wakolbinger (2005) give a connection to stable processes.

Birkner and Blath (2009) describe the  $\Lambda$ -Fleming-Viot process and discrete models whose limit gives rise to it.

If  $F$  has a single atom at 0, then  $\{\mathbf{X}_t\}_{t \geq 0}$  is the  $d$ -dimensional Wright-Fisher diffusion process with generator

$$\mathcal{L} = \frac{1}{2} \sum_{i,j=1}^d x_i(\delta_{ij} - x_j) \frac{\partial^2}{\partial x_i \partial x_j}$$

$X_1(t)$  is a one-dimensional Wright-Fisher diffusion process with generator

$$\mathcal{L} = \frac{1}{2} x(1-x) \frac{\partial^2}{\partial x^2}$$

The  $\Lambda$ -coalescent process is a random tree process back in time which has multiple merger rates for a specific  $k$  lineages coalescing while  $n$  edges in the tree of

$$\lambda_{nk} = \int_0^1 x^k (1-x)^{n-k} \frac{F(dx)}{x^2}, \quad k \geq 2$$

After coalescence there are  $n - k + 1$  edges in the tree.

This process was introduced by Pitman (1999), Sagitov (1999) and has been extensively studied. Berestycki (2009); recent results in the  $\Lambda$ -coalescent.

There is a connection between continuous state branching processes and the  $\Lambda$ -coalescent. The connection is through the Laplace exponent

$$\psi(q) = \int_0^1 (e^{-qy} - 1 + qy) y^{-2} F(dy)$$

Bertoin and Le Gall (2006) showed that the  $\Lambda$ -coalescent comes down from infinity under the same condition that the continuous state branching process becomes extinct in finite time, that is when

$$\int_1^\infty \frac{dq}{\psi(q)} < \infty$$

## Some papers on the $\Lambda$ -coalescent

Berestycki (2009) Recent progress in coalescent theory.

Berestycki, Berestycki, and Limic (2012) Asymptotic sampling formulae for  $\Lambda$ -coalescents.

Berestycki, Berestycki, and Limic (2012) A small-time coupling between  $\Lambda$ -coalescents and branching processes.

Bertoin and Le Gall (2003) Stochastic flows associated to coalescent processes.

Bertoin and Le Gall (2006) J. Bertoin and J.-F. Le Gall (2006). Stochastic flows associated to coalescent processes III: Limit theorems.

Birkner, Blath, Capaldo, Etheridge, Möhle, Schweinsberg, Wakolbinger, (2005) Alpha-stable branching and Beta-coalescents.

Birkner and Blath (2009) Measure-Valued diffusions, general coalescents and population genetic inference.

## A Wright-Fisher generator connection

### Theorem

Let  $\mathcal{L}$  be the  $\Lambda$ -Fleming-Viot generator,  $V$  be a uniform random variable on  $[0, 1]$ ,  $U$  a random variable on  $[0, 1]$  with density  $2u$ ,  $0 < u < 1$  and  $W = YU$ , where  $Y$  has distribution  $F$  and  $V, U, Y$  are independent. Denote the second derivatives of a function  $g(\mathbf{x})$  by  $g_{ij}(\mathbf{x})$ .

Then

$$\mathcal{L}g(\mathbf{x}) = \frac{1}{2} \sum_{i,j=1}^d x_i(\delta_{ij} - x_j) \mathbb{E} \left[ g_{ij}(\mathbf{x}(1 - W) + WV\mathbf{e}_i) \right]$$

where expectation  $\mathbb{E}$  is taken over  $V, W$ .



Wright-Fisher generator

$$\mathcal{L}g(\mathbf{x}) = \frac{1}{2} \sum_{i,j=1}^d x_i(\delta_{ij} - x_j)g_{ij}(\mathbf{x})$$

$\Lambda$ -Fleming-Viot generator

$$\mathcal{L}g(\mathbf{x}) = \frac{1}{2} \sum_{i,j=1}^d x_i(\delta_{ij} - x_j)\mathbb{E}\left[g_{ij}(\mathbf{x}(1 - W) + WV\mathbf{e}_i)\right]$$

Method of proof

$$\mathcal{L}g(\mathbf{x}) = \int_0^1 \sum_{i=1}^d x_i \left( g(\mathbf{x}(1-y) + y\mathbf{e}_i) - g(\mathbf{x}) \right) \frac{F(dy)}{y^2}$$

$$\mathcal{L}g(\mathbf{x}) = \frac{1}{2} \sum_{i,j=1}^d x_i (\delta_{ij} - x_j) \mathbb{E} \left[ g_{ij}(\mathbf{x}(1-W) + WV\mathbf{e}_i) \right]$$

Show that the generators have the same answer acting on

$$g(\mathbf{x}) = \exp \left\{ \sum_{i=1}^d \eta_i x_i \right\}, \quad \boldsymbol{\eta} \in \mathbb{R}^d$$

## 1-dimensional generator

Wright-Fisher diffusion generator

$$\mathcal{L}g(x) = \frac{1}{2}x(1-x)g''(x)$$

$\Lambda$ -Fleming-Viot process generator

$$\mathcal{L}g(x) = \frac{1}{2}x(1-x)\mathbb{E}\left[g''(x(1-W) + WV)\right]$$

or

$$\mathcal{L}g(x) = \frac{1}{2}x(1-x)\mathbb{E}\left[\frac{g'(x(1-W) + W) - g'(x(1-W))}{W}\right]$$

The Laplace transform of  $W$  is related to the Laplace exponent by

$$\mathbb{E}\left[e^{-\eta W}\right] = 2 \int_0^1 \frac{e^{-\eta y} - 1 + \eta y}{(y\eta)^2} F(dy)$$

$W = UY$  is continuous in  $(0, 1)$  with a possible atom at 0.

$$P(W = 0) = P(Y = 0)$$

## Adding mutation

The generator has an additional term added of

$$\frac{\theta}{2} \sum_{i=1}^d \left( \sum_{j=1}^d p_{ji} x_j - x_i \right) \frac{\partial}{\partial x_i}$$

If mutation is **parent independent**  $\theta p_{ij} = \theta_j$ , not depending on  $i$ , and the additional term is

$$\frac{1}{2} \sum_{i=1}^d \left( \sum_{j=1}^d \theta_j x_j - \theta x_i \right) \frac{\partial}{\partial x_i}$$

## Eigenstructure of the $\Lambda$ -Fleming-Viot process

### Theorem

Let  $\{\lambda_{\mathbf{n}}\}$ ,  $\{P_{\mathbf{n}}(\mathbf{x})\}$  be the eigenvalues and eigenvectors of  $\mathcal{L}$ , the generator which includes mutation, satisfying

$$\mathcal{L}P_{\mathbf{n}}(\mathbf{x}) = -\lambda_{\mathbf{n}}P_{\mathbf{n}}(\mathbf{x})$$

Denote the  $d - 1$  eigenvalues of the mutation matrix  $P$  which have modulus less than 1 by  $\{\phi_k\}_{k=1}^{d-1}$ .

The eigenvalues of  $\mathcal{L}$  are

$$\lambda_{\mathbf{n}} = \frac{1}{2}n(n-1)\mathbb{E}\left[(1-W)^{n-2}\right] + \frac{\theta}{2}\sum_{k=1}^{d-1}(1-\phi_k)n_k$$

## Polynomial Eigenvectors

Denote the  $d - 1$  eigenvalues of  $P$  which have modulus less than 1 by  $\{\phi_k\}_{k=1}^{d-1}$  corresponding to eigenvectors which are rows of a  $d - 1 \times d$  matrix  $R$  satisfying

$$\sum_{i=1}^d r_{ki} p_{ji} = \phi_k r_{kj}, \quad k = 1, \dots, d - 1.$$

Define a  $d - 1$  dimensional vector  $\boldsymbol{\xi} = R\mathbf{x}$ .

The polynomials  $P_n(\mathbf{x})$  are polynomials in the  $d - 1$  terms in  $\boldsymbol{\xi} = (\xi_1, \dots, \xi_{d-1})$  whose only leading term of degree  $n$  is

$$\prod_{j=1}^{d-1} \xi_j^{n_j}$$

In the **parent independent** model of mutation

$$\lambda_{\mathbf{n}} = \lambda_n = \frac{1}{2}n \left\{ (n-1)\mathbb{E} \left[ (1-W)^{n-2} \right] + \theta \right\}$$

repeated  $\binom{n+d-2}{n}$  times with non-unique polynomial eigenfunctions within the same degree  $n$ .

The non-unit eigenvalues of the mutation matrix with identical rows are zero.

**Wright-Fisher diffusion, general mutation structure.**

$$\lambda_{\mathbf{n}} = \frac{1}{2}n(n-1) + \frac{\theta}{2} \sum_{k=1}^{d-1} (1-\phi_k)n_k$$



## $\Lambda$ -coalescent eigenvalues and rates

$$\frac{1}{2}n(n-1)\mathbb{E}\left[(1-W)^{n-2}\right] = \sum_{k=2}^{\infty} \binom{n}{k} \lambda_{nk}$$

which is the total jump rate away from  $n$  individuals.

These are the eigenvalues in the  $\Lambda$ -coalescent tree.

The individual rates can be expressed as

$$\begin{aligned}\binom{n}{k} \lambda_{nk} &= \binom{n}{k} \int_0^1 y^k (1-y)^{n-k} \frac{F(dy)}{y^2} \\ &= \frac{n}{2} \mathbb{E} \left[ \frac{P_k(n, W) - P_{k-1}(n, W)}{W^2} \right],\end{aligned}$$

where

$$P_k(n, w) = \binom{n-1}{k-1} (1-w)^{n-k} w^k$$

is a negative binomial probability of a waiting time of  $n$  trials to obtain  $k$  successes, where the success probability is  $w$ .

Two types

The generator is specified by

$$\mathcal{L}g(x) = \frac{1}{2}x(1-x)\mathbb{E}\left[g''(x(1-W) + WV)\right] + \frac{1}{2}(\theta_1 - \theta x)g'(x)$$

The eigenvalues are

$$\lambda_n = \frac{1}{2}n \left\{ (n-1)\mathbb{E}\left[(1-W)^{n-2}\right] + \theta \right\}$$

and the eigenvectors are polynomials satisfying

$$\mathcal{L}P_n(x) = -\lambda_n P_n(x), \quad n \geq 1.$$

## Polynomial eigenvectors

$$\begin{aligned} & \frac{1}{2}x(1-x)\mathbb{E}\left[\frac{P'_n(x(1-W)+W) - P'_n(x(1-W))}{W}\right] \\ & + \frac{1}{2}(\theta_1 - \theta x)P'_n(x) \\ & = \frac{1}{2}n\left[(n-1)\mathbb{E}\left[(1-W)^{n-2}\right] + \theta\right]P_n(x) \end{aligned}$$

The monic polynomial  $P_n(x)$  is uniquely defined by recursion of its coefficients.

Stationary distribution  $\psi(x)$

$$\int_0^1 \mathcal{L}g(x)\psi(x)dx = 0$$

$$\sigma^2(x) = x(1-x), \quad \mu(x) = \theta_1 - \theta x$$

$$k(x) = \mathbb{E}\left[(1-W)^{-2}g(x(1-W) + VW)\right]$$

An equation for the stationary distribution

$$0 = \int_0^1 \left[ k(x) \frac{1}{2} \frac{d^2}{dx^2} [\sigma^2(x)\psi(x)] - g(x) \frac{d}{dx} [\mu(x)\psi(x)] \right] dx \\ + k(x) \frac{d}{dx} \left[ \frac{1}{2} \sigma^2(x)\psi(x) \right] \Big|_0^1 + g(x)\mu(x)\psi(x) \Big|_0^1$$

In a diffusion process  $k(x) = g(x)$  and the boundary terms vanish. Then there is a solution found by solving

$$\frac{1}{2} \frac{d^2}{dx^2} [\sigma^2(x)\psi(x)] - \frac{d}{dx} [\mu(x)\psi(x)] = 0$$

however  $k(x) \neq g(x)$  so we do not have an equation like this.

Green's function,  $\gamma(x)$

Solve, for a given function  $g(x)$

$$\mathcal{L}\gamma(x) = -g(x), \quad \gamma(0) = \gamma(1) = 0.$$

Then

$$\gamma(x) = \int_0^1 G(x, \xi)g(\xi)d\xi$$

A non-linear equation, equivalent to

$$\frac{1}{2}x(1-x)\mathbb{E}\left[\gamma''(x(1-W) + VW)\right] = -g(x)$$

## Green's function solution

Define

$$k(x) = \mathbb{E} \left[ (1 - W)^{-2} \gamma(x(1 - W) + VW) \right]$$

then

$$k''(x) = -2 \frac{g(x)}{x(1-x)}$$

with a solution

$$\begin{aligned} k(x) = & k(0)(1-x) + k(1)x \\ & + (1-x) \int_0^x \frac{2g(\eta)}{(1-\eta)} d\eta + x \int_x^1 \frac{2g(\eta)}{\eta} d\eta \end{aligned}$$



## Mean time to absorption

If  $g(x) = 1$ ,  $x \in (0, 1)$  then  $\gamma(x)$  is the mean time to absorption at 0 or 1 when  $X(0) = x$ .

$$k(x) = k(0)(1 - x) + k(1)x + (1 - x) \int_0^x \frac{2}{(1 - \eta)} d\eta + x \int_x^1 \frac{2}{\eta} d\eta$$

There is a non-linear equation to solve of

$$k(x) = k(0)(1 - x) + k(1)x - 2(1 - x) \log(1 - x) - 2x \log x$$

where

$$k(x) = \mathbb{E} \left[ (1 - W)^{-2} \gamma(x(1 - W) + VW) \right]$$

## Stationary distribution, $\Lambda$ -Fleming process with mutation

Let  $Z$  be a random variable with the size-biased distribution of  $X$ ,  $Z_*$  a size-biased  $Z$  random variable and  $Z^*$  a size-biased random variable with respect to  $1 - Z$ .

Let  $B$  be Bernoulli random variable, independent of the other random variables in the following equation such that

$$P(B = 1) = \frac{\theta - \theta_1}{\theta(\theta_1 + 1)}$$

An interesting distributional identity

$$VZ_* \stackrel{\mathcal{D}}{=} (1 - B)VZ + B(Z^*(1 - W) + WV)$$

## The frequency spectrum in the infinitely-many-alleles model

Take a limit from a  $d$ -allele model with  $\theta_i = \theta/d$ ,  $i \in [d]$ . The limit is a point process  $\{X_i\}_{i=1}^{\infty}$ . The 1-dimensional frequency spectrum  $h(x)$  is a non-negative measure such that for suitable functions  $k$  on  $[0,1]$  in the stationary distribution

$$\mathbb{E} \left[ \sum_{i=1}^{\infty} k(X_i) \right] = \int_0^1 k(x) h(x) dx.$$

Symmetry in the  $d$ -allele model shows that

$$\int_0^1 k(x) h(x) dx = \lim_{d \rightarrow \infty} d \mathbb{E} \left[ k(X_1) \right].$$

The classical Wright-Fisher diffusion gives rise to the Poisson-Dirichlet process with a frequency spectrum of

$$h(x) = \theta x^{-1} (1-x)^{\theta-1}, \quad 0 < x < 1.$$

## $\Lambda$ Fleming-Viot frequency spectrum

Let  $Z$  have a density

$$f(z) = zh(z), \quad 0 < z < 1$$

Interesting identity

$$VZ_* \stackrel{\mathcal{D}}{=} Z^*(1 - W) + WV$$

where  $Z_*$  is size-biased with respect to  $Z$ , and  $Z^*$  is size-biasing with respect to  $1 - Z$ .

The constant  $\theta$  appears in the identity through scaling in the size-biased distributions.

Limit distribution of excess life in a renewal process

$$P(VZ_* > \eta) = \int_{\eta}^1 \frac{P(Z > z)}{\mathbb{E}[Z]} dz$$

## Typed dual $\Lambda$ -coalescent

The  $\Lambda$ -Fleming-Viot process is dual to the system of coalescing lineages  $\{L(t)\}_{t \geq 0}$  which takes values in  $\mathbb{Z}_+^d$  and for which the transition rates are, for  $i, j \in [d]$  and  $l \geq 2$ ,

$$q_\Lambda(\xi, \xi - e_i(l-1)) = \int_{[0,1]} \binom{|\xi|}{l} y^l (1-y)^{|\xi|-l} \frac{F(dy)}{y^2} \\ \times \frac{\xi_i + 1 - l}{|\xi| + 1 - l} \frac{\mathcal{M}(\xi - e_i(l-1))}{\mathcal{M}(\xi)}$$
$$q_\Lambda(\xi, \xi + e_i - e_j) = \mu_{ij}(\xi_i + 1 - \delta_{ij}) \frac{\mathcal{M}(\xi + e_i - e_j)}{\mathcal{M}(\xi)}$$

The process is constructed as a moment dual from the generator.

## A different dual process

Define a sequence of monic polynomials  $\{g_n(x)\}$  by the generator equation

$$\begin{aligned} & \frac{1}{2}x(1-x)\mathbb{E}g_n''(x(1-W) + VW) + \frac{1}{2}(\theta_1 - \theta x)g_n'(x) \\ &= \binom{n}{2}\mathbb{E}(1-W)^{n-2}[g_{n-1}(x) - g_n(x)] + n\frac{1}{2}[\theta_1 g_{n-1}(x) - \theta g_n(x)] \end{aligned}$$

The defining equation mimics the Wright-Fisher diffusion acting on test functions  $g_n(x) = x^n$

$$\begin{aligned} & \frac{1}{2}x(1-x)\frac{d^2}{dx^2}x^n + \frac{1}{2}(\theta_1 - \theta x)\frac{d}{dx}x^n \\ &= \binom{n}{2}(x^{n-1} - x^n) + \frac{1}{2}n(\theta_1 x^{n-1} - \theta x^n) \end{aligned}$$

## Jacobi polynomial analogues

The eigenfunctions are polynomials  $\{P_n(x)\}$  satisfying

$$\mathcal{L}P_n(x) = \lambda_n P_n(x)$$

$$P_n(x) = g_n(x) + \sum_{r=0}^{n-1} c_{nr} g_r(x)$$

The coefficients are

$$c_{nr} = \frac{\lambda_{r+1}^\circ \cdots \lambda_n^\circ}{(\lambda_r - \lambda_n) \cdots (\lambda_{n-1} - \lambda_n)}$$

$$\text{where } \lambda_n = \frac{n}{2} \left[ (n-1) \mathbb{E}(1-W)^{n-2} + \theta \right]$$
$$\lambda_n^\circ = \frac{n}{2} \left[ (n-1) \mathbb{E}(1-W)^{n-2} + \theta_1 \right]$$

In the stationary distribution

$$\mathbb{E}\left[g_n(X)\right] = \omega_n$$

where  $\omega_n$  is a **Beta moment analog**

$$\omega_n = \frac{\prod_{j=1}^n \left( (j-1)\mathbb{E}(1-W)^{j-2} + \theta_1 \right)}{\prod_{j=1}^n \left( (j-1)\mathbb{E}(1-W)^{j-2} + \theta \right)}$$

Let

$$h_n = \frac{g_n}{\omega_n}$$

so

$$\mathbb{E}\left[h_n(X)\right] = 1$$



## Dual Generator

$$\mathcal{L}h_n = \lambda_n [h_{n-1} - h_n]$$

## Dual equation

$$\mathbb{E}_{X(0)=x} [h_n(X(t))] = \mathbb{E}_{N(0)=n} [h_{N(t)}(x)]$$

where  $\{N(t), t \geq 0\}$  is a death process with rates

$$\lambda_n = \frac{1}{2}n \left( (n-1)\mathbb{E}[(1-W)^{n-2}] + \theta \right)$$

The process  $\{N(t), t \geq 0\}$  comes down from infinity if and only if

$$\sum_{n=2}^{\infty} \lambda_n^{-1} < \infty$$

which implies the  $\Lambda$ -coalescent comes down from infinity.

The transition functions are then

$$P(N(t) = j \mid N(0) = i) \\ = \sum_{k=j}^i e^{-\lambda_k t} (-1)^{k-j} \frac{\prod_{\{l: j \leq l \leq k+1\}} \lambda_l}{\prod_{\{l: j \leq l \leq k+1, l \neq k\}} (\lambda_l - \lambda_k)}$$

$P(N(t) = j \mid N(0) = \infty)$  is well defined if  $N(t)$  comes down from infinity.

In the Kingman coalescent the death rates are  $n(n - 1 + \theta)/2$  and  $N(t)$  describes the number of non-mutant lineages at time  $t$  back in the population.

Wright-Fisher diffusion. Fixation probability with selection.  
Two types.

Let  $P(x)$  be the probability that the 1st type fixes, starting from an initial frequency of  $x$ .  $P(x)$  is the solution of

$$\mathcal{L}^\sigma P(x) = \frac{1}{2}x(1-x)P''(x) - \sigma x(1-x)P'(x) = 0$$

with  $P(0) = 0$ ,  $P(1) = 1$ . The solution of this differential equation is

$$P(x) = \frac{e^{2\sigma x} - 1}{e^{2\sigma} - 1}$$

$\Lambda$ -Fleming Viot process. Fixation probability with selection.

Let  $P(x)$  be the probability that the 1st type fixes, starting from an initial frequency of  $x$ .  $P(x)$  is the solution of

$$\mathcal{L}^\sigma P(x) = \frac{1}{2}x(1-x)\mathbb{E}\left[P''(x(1-W)+WV)\right] - \sigma x(1-x)P'(x) = 0$$

with  $P(0) = 0$ ,  $P(1) = 1$ .

Der, Epstein and Plotkin (2011,2012). For some measures  $F$  and  $\sigma$  it is possible that  $P(x) = 1$  or  $0$  for all  $x \in (0, 1)$ . If  $\sigma > 0$ , fixation is certain if

$$\sigma > - \int_0^1 \frac{\log(1-y)}{y^2} F(dy)$$

A computational solution for  $P(x)$  when fixation or loss is not certain from  $x \in (0, 1)$ .

Define a sequence of polynomials  $\{h_n(x)\}_{n=0}^{\infty}$  for a choice of pre-specified constants  $\{h_n(0)\}$  as solutions of

$$\mathbb{E} \left[ \frac{h_n(x(1-W) + W) - h_n(x(1-W))}{W} \right] = nh_{n-1}(x)$$

where the leading coefficient in  $h_n(x)$  is

$$\frac{1}{\prod_{j=1}^{n-1} \mathbb{E}[(1-W)^j]}$$

Polynomial solution for  $P(x)$

$$P(x) = \left(e^{2\sigma} - 1\right)^{-1} \sum_{n=1}^{\infty} \frac{(2\sigma)^n}{n!} H_n(x),$$

where  $\{H_n(x)\}$  are polynomials derived from

$$H_n(x) = \int_0^x n h_{n-1}(\xi) d\xi$$

and the constants  $\{h_n(0)\}$  are chosen so that

$$\int_0^1 n h_{n-1}(\xi) d\xi = 1$$

The coefficients of  $H_n(x)$  are well defined by a recurrence relationship with the coefficients of  $H_{n-1}(x)$ .