

Vladimir Vatutin  
Steklov Mathematical Institute  
Moscow, Russia

---

## Branching processes evolving in asynchronous environment

Vladimir Vatutin (Steklov Mathematical Institute, Moscow)

joint work with

Quansheng Liu (Universite de Bretagne Sud, Vannes)

France, 2011

## Galton-Watson processes in random environment

- Offspring generating functions  $f_n(s) := \mathbf{E}s^{\xi^{(n)}}$  in generations  $n = 0, 1, \dots$  are **RANDOM** and I.I.D.

## Galton-Watson processes in random environment

- Offspring generating functions  $f_n(s) := \mathbf{E}s^{\xi^{(n)}}$  in generations  $n = 0, 1, \dots$  are **RANDOM** and **I.I.D.**  
 $\Rightarrow f'_n(1) := \mathbf{E}\xi^{(n)}$  are **I.I.D. RANDOM** variables;

## Galton-Watson processes in random environment

- Offspring generating functions  $f_n(s) := \mathbf{E}s^{\xi^{(n)}}$  in generations  $n = 0, 1, \dots$  are **RANDOM** and **I.I.D.**  
 $\Rightarrow f'_n(1) := \mathbf{E}\xi^{(n)}$  are **I.I.D. RANDOM** variables;

- 

$$Z_{n+1} = \sum_{j=1}^{Z_n} \xi_j^{(n)} \quad \text{or} \quad \mathbf{E}[s^{Z_{n+1}} \mid Z_n; f_0, f_1, \dots] = (f_n(s))^{Z_n}.$$

## Galton-Watson processes in random environment

- Offspring generating functions  $f_n(s) := \mathbf{E}s^{\xi^{(n)}}$  in generations  $n = 0, 1, \dots$  are **RANDOM** and I.I.D.

$\Rightarrow f'_n(1) := \mathbf{E}\xi^{(n)}$  are **I.I.D. RANDOM** variables;

- 

$$Z_{n+1} = \sum_{j=1}^{Z_n} \xi_j^{(n)} \quad \text{or} \quad \mathbf{E}\left[s^{Z_{n+1}} \mid Z_n; f_0, f_1, \dots\right] = \left(f_n(s)\right)^{Z_n}.$$

- $\xi_j^{(n)} \stackrel{d}{=} \xi^{(n)}$  are i.i.d. **given**  $f_0, f_1, \dots$

## Classification of Galton-Watson processes in **i.i.d.** random environment

$$\mathbf{E} \left[ Z_n \middle| f_0, f_1, \dots \right] =: \mathbf{E}_f \left[ Z_n \right]$$

## Classification of Galton-Watson processes in **i.i.d.** random environment

$$\mathbf{E}\left[Z_n \middle| f_0, f_1, \dots\right] =: \mathbf{E}_{\textcolor{red}{f}}\left[Z_{\textcolor{blue}{n}}\right] = \prod_{j=0}^{n-1} f'_j(1)$$

## Classification of Galton-Watson processes in **i.i.d.** random environment

$$\begin{aligned}\mathbf{E} \left[ Z_n \middle| f_0, f_1, \dots \right] &=: \mathbf{E}_{\textcolor{red}{f}} \left[ Z_n \right] = \prod_{j=0}^{n-1} f'_j(1) \\ &= e^{\sum_{j=0}^{n-1} \log f'_j(1)} = e^{S_n}\end{aligned}$$

- $X_i := \log f'_{i-1}(1)$
- $S_0 = 0, \quad S_n = X_1 + X_2 + \dots + X_n, \quad n \geq 1.$

$$\mathbf{E}_{\textcolor{red}{f}}[Z_{\textcolor{blue}{n}}] = e^{S_n}$$

### Associated random walk

- $X_i := \log f'_{i-1}(1)$
- $S_0 = 0, \quad S_n = X_1 + X_2 + \cdots + X_n, \quad n \geq 1.$

**Classification** (Afanasyev, Geiger, Kersting, V.(2005))

$$\mathbf{E}_{\textcolor{red}{f}}[Z_n] = e^{S_n}$$

### Associated random walk

- $X_i := \log f'_{i-1}(1)$
- $S_0 = 0, \quad S_n = X_1 + X_2 + \cdots + X_n, \quad n \geq 1.$

**Classification** ([Afanasyev, Geiger, Kersting, V. \(2005\)](#)): A BPRE is called

- **super**critical if  $\lim_{n \rightarrow \infty} S_n = +\infty$  with probability 1;

$$\mathbf{E}_{\textcolor{red}{f}}[Z_n] = e^{S_n}$$

### Associated random walk

- $X_i := \log f'_{i-1}(1)$
- $S_0 = 0, \quad S_n = X_1 + X_2 + \cdots + X_n, \quad n \geq 1.$

**Classification** ([Afanasyev, Geiger, Kersting, V. \(2005\)](#)): A BPRE is called

- **super**critical if  $\lim_{n \rightarrow \infty} S_n = +\infty$  with probability 1;
- **sub**critical if  $\lim_{n \rightarrow \infty} S_n = -\infty$  with probability 1;

$$\mathbf{E}_{\textcolor{red}{f}}[Z_n] = e^{S_n}$$

### Associated random walk

- $X_i := \log f'_{i-1}(1)$
- $S_0 = 0, \quad S_n = X_1 + X_2 + \cdots + X_n, \quad n \geq 1.$

**Classification** ([Afanasyev, Geiger, Kersting, V. \(2005\)](#)): A BPRE is called

- **super**critical if  $\lim_{n \rightarrow \infty} S_n = +\infty$  with probability 1;
- **sub**critical if  $\lim_{n \rightarrow \infty} S_n = -\infty$  with probability 1;
- critical if  $\limsup_{n \rightarrow \infty} S_n = +\infty$  and  $\liminf_{n \rightarrow \infty} S_n = -\infty$  (both with probability 1);

Critical case includes

$$\mathbf{E} \log f'(1) = \mathbf{E}X = 0, \quad 0 < \mathbf{E}X^2 < \infty$$

**Quenched approach:** Let

$$T := \min\{k : Z_k = 0\}.$$

We study

$$\mathbf{P}_f(Z_n > 0) = \mathbf{P}_f(T > n) = \mathbf{P}(T > n | f_0, f_1, \dots),$$

as a random variable on the space of realizations of the environment  $f_0, f_1, \dots$

**Quenched approach:** Let

$$T := \min\{k : Z_k = 0\}.$$

We study

$$\mathbf{P}_f(Z_n > 0) = \mathbf{P}_f(T > n) = \mathbf{P}(T > n | f_0, f_1, \dots),$$

as a **random variable** on the space of realizations of the environment  $f_0, f_1, \dots$  and

$$\mathbf{P}_f(Z_n \in dx | T > n)$$

as a **random conditional law.**

## Basic assumptions

The distribution of  $X = \log f'(1)$  is nonlattice and belongs (without centering) to an  $\alpha$ -stable law with  $\alpha \in (0, 2]$  and

$$\rho = \mathbf{P}(X > 0) \in (0, 1).$$

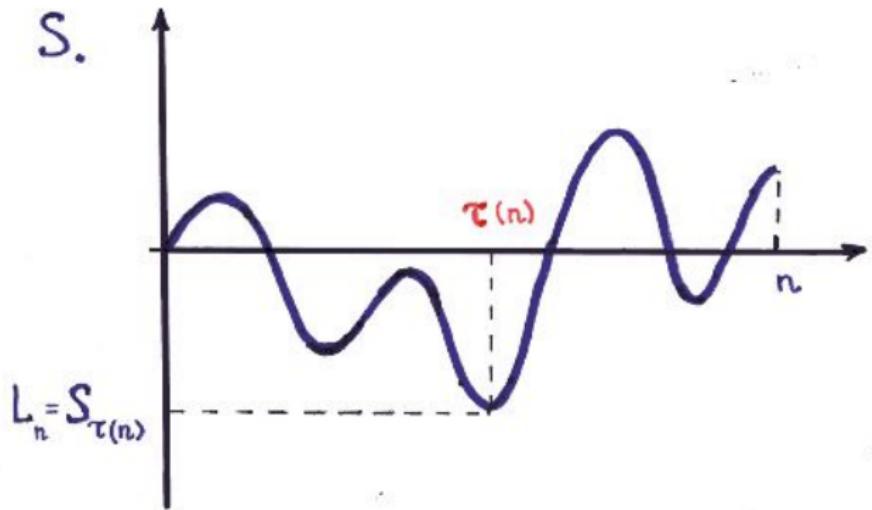
Set

$$M_n := \max(S_1, \dots, S_n), \quad L_n := \min(S_1, \dots, S_n)$$

and let

$$\mu(n) := \max \{1 \leq i \leq n : S_i = M_n\},$$

$$\tau(n) := \min \{0 \leq i \leq n : S_i = \min(0, L_n)\}.$$



Set

$$M_n := \max(S_1, \dots, S_n), \quad L_n := \min(S_1, \dots, S_n)$$

and let

$$\mu(n) := \max \{1 \leq i \leq n : S_i = M_n\},$$

$$\tau(n) := \min \{0 \leq i \leq n : S_i = \min(0, L_n)\}.$$

It is known that if  $X$  belongs (without centering) to an  $\alpha$ -stable law with  $\rho \in (0, 1)$  then for any  $t \in (0, 1]$ , as  $n \rightarrow \infty$

$$\frac{\tau(nt)}{n} \xrightarrow{d} \tau_t, \quad \frac{\mu(nt)}{n} \xrightarrow{d} \mu_t,$$

where  $t^{-1}\tau_t$  and  $t^{-1}\mu_t$  have the **generalized** arcsine distributions on  $[0, 1]$  with densities

$$\frac{C}{x^\rho(1-x)^{1-\rho}}, \quad \frac{C}{x^{1-\rho}(1-x)^\rho},$$

respectively.

Let

$$\gamma_0 := 0, \quad \gamma_{j+1} := \min(n > \gamma_j : S_n < S_{\gamma_j})$$

and

$$\Gamma_0 := 0, \quad \Gamma_{j+1} := \min(n > \Gamma_j : S_n > S_{\Gamma_j}), \quad j \geq 0,$$

be the strict descending and ascending ladder epochs of the random walk  $\{S_n, n \geq 0\}$ .

Let

$$\gamma_0 := 0, \quad \gamma_{j+1} := \min(n > \gamma_j : S_n < S_{\gamma_j})$$

and

$$\Gamma_0 := 0, \quad \Gamma_{j+1} := \min(n > \Gamma_j : S_n > S_{\Gamma_j}), \quad j \geq 0,$$

be the strict descending and ascending ladder epochs of the random walk  $\{S_n, n \geq 0\}$ . Set

$$U(x) = 1 + \sum_{j=1}^{\infty} \mathbf{P}(S_{\Gamma_j} < x), \quad x \geq 0, \quad U(x) = 0, \quad x < 0,$$

$$V(x) = \sum_{j=0}^{\infty} \mathbf{P}(S_{\gamma_j} \geq -x), \quad x \geq 0, \quad V(x) = 0, \quad x < 0.$$

Consider the filtration  $\mathcal{F} = (\mathcal{F}_n)$ , where  $\mathcal{F}_n = \sigma(f_0, \dots, f_{n-1}, Z_0, \dots, Z_n)$ .

Consider the filtration  $\mathcal{F} = (\mathcal{F}_n)$ , where  $\mathcal{F}_n = \sigma(f_0, \dots, f_{n-1}, Z_0, \dots, Z_n)$ .

For any bounded,  $\mathcal{F}_n$ -measurable random variable  $R_n$  we construct probability measures  $\mathbf{P}_x^-, x \leq 0$ , and  $\mathbf{P}_y^+, y \geq 0$ , fulfilling for each  $n$  the equalities

$$\mathbf{E}_x^- [R_n] := \frac{1}{U(-x)} \mathbf{E}_x [R_n U(-S_n); M_n < 0]$$

and

$$\mathbf{E}_y^+ [R_n] := \frac{1}{V(y)} \mathbf{E}_y [R_n V(S_n); L_n \geq 0].$$

Consider the filtration  $\mathcal{F} = (\mathcal{F}_n)$ , where  $\mathcal{F}_n = \sigma(f_0, \dots, f_{n-1}, Z_0, \dots, Z_n)$ .

For any bounded,  $\mathcal{F}_n$ -measurable random variable  $R_n$  we construct probability measures  $\mathbf{P}_x^-, x \leq 0$ , and  $\mathbf{P}_y^+, y \geq 0$ , fulfilling for each  $n$  the equalities

$$\mathbf{E}_x^- [R_n] := \frac{1}{U(-x)} \mathbf{E}_x [R_n U(-S_n); M_n < 0]$$

and

$$\mathbf{E}_y^+ [R_n] := \frac{1}{V(y)} \mathbf{E}_y [R_n V(S_n); L_n \geq 0].$$

The measures specified in this way are consistent in  $n$ .

In particular, the measure  $\mathbf{P}_0^-$  "corresponds" to the measure generated by the random walk conditioned to stay **negative**, while the measure  $\mathbf{P}_0^+$  "corresponds" to the measure generated by the random walk conditioned to stay **nonnegative**.

**Basic assumption:**

- **FROM NOW ON!!!**

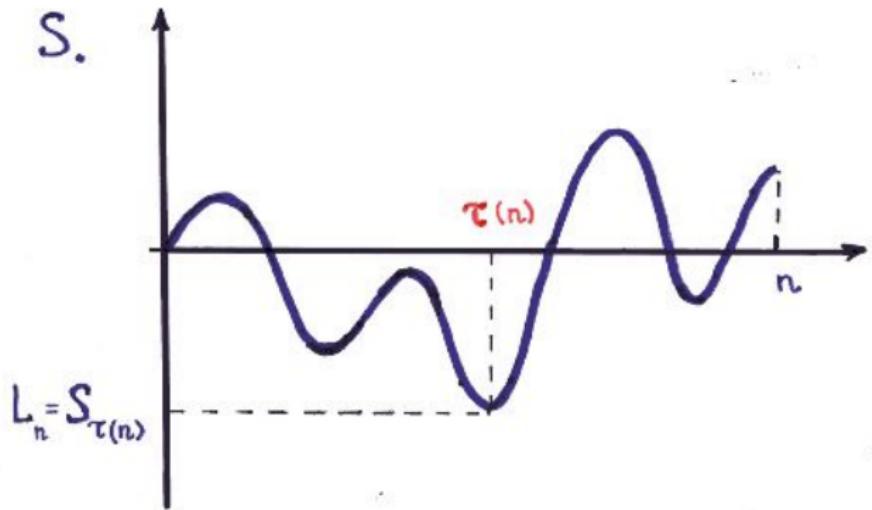
$$f_n(s) = \frac{q_n}{1 - p_n s}, \quad n = 0, 1, 2, \dots$$

(V. + Dyakonova)

In the critical case and **for the quenched approach** we have the following convergence in distribution

$$\frac{\mathbf{P}_{\textcolor{red}{f}}(Z_{\textcolor{blue}{n}} > 0)}{e^{S_{\tau(n)}}} \xrightarrow{d} \zeta$$

where  $\zeta \in (0, 1]$  with probability 1.



## (V. + Dyakonova)

In the critical case and **for the quenched approach** we have the following convergence in distribution

$$\frac{\mathbf{P}_f(Z_n > 0)}{e^{S_{\tau(n)}}} \xrightarrow{d} \zeta$$

where  $\zeta \in (0, 1]$  with probability 1.

$\Rightarrow$  if  $\sigma^2 = \text{Var}X < \infty$

$$\mathbf{P}_f(Z_n > 0) \approx \zeta e^{S_{\tau(n)}} = \zeta e^{\sigma\sqrt{n} \times \frac{1}{\sigma\sqrt{n}} S_{\tau(n)}} \approx \zeta e^{-|g|\sigma\sqrt{n}}$$

where  $g$  is a Gaussian random variable:

$$\mathbf{P}(g < x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy.$$

Let now  $S_0^\pm = 0$  and

$$(S_0^+, S_1^+, \dots, S_n^+, \dots)$$

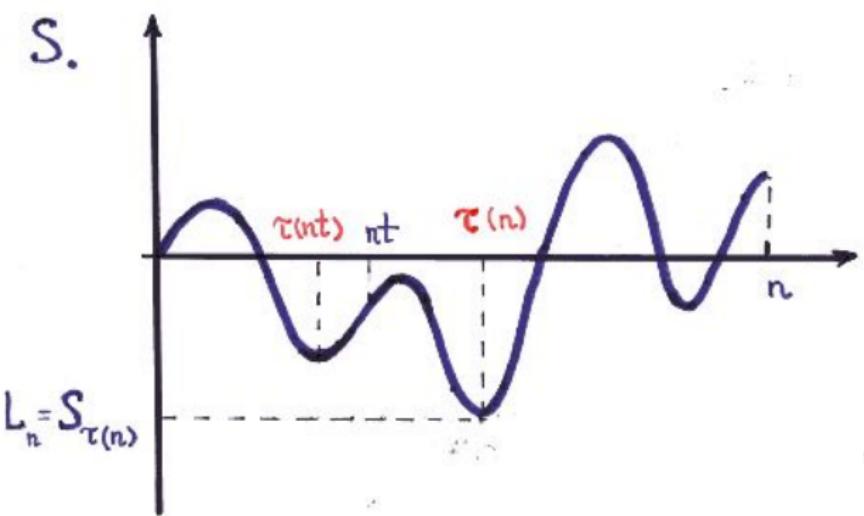
be a random walk conditioned to stay positive and

$$(S_1^-, \dots, S_n^-, \dots)$$

be a random walk conditioned to stay negative. Then

$$\frac{1}{\zeta} \stackrel{d}{=} \sum_{k=0}^{\infty} e^{-S_k^+} + \sum_{k=1}^{\infty} e^{S_k^-}.$$

**THE NUMBER OF PARTICLES BEFORE THE POINT OF  
GLOBAL MINIMUM OF  $S_k, 0 \leq k \leq n$**



## THE NUMBER OF PARTICLES **BEFORE** THE POINT OF GLOBAL MINIMUM OF $S_k, 0 \leq k \leq n$

Observe that

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

## THE NUMBER OF PARTICLES BEFORE THE POINT OF GLOBAL MINIMUM OF $S_k, 0 \leq k \leq n$

Observe that

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

**Theorem 1. (V. + Dyakonova)**

For any  $t \in (0, 1)$ , any  $x > 0$  as  $n \rightarrow \infty$

$$\mathbf{P}_f \left( \frac{Z_{nt}}{\mathbf{E}_f [Z_{nt} | Z_{nt} > 0]} \leq x \middle| T > n; nt < \tau(n) \right) \xrightarrow{d} \int_0^x ye^{-y} dy.$$

## THE NUMBER OF PARTICLES BEFORE THE POINT OF GLOBAL MINIMUM OF $S_k, 0 \leq k \leq n$

Observe that

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

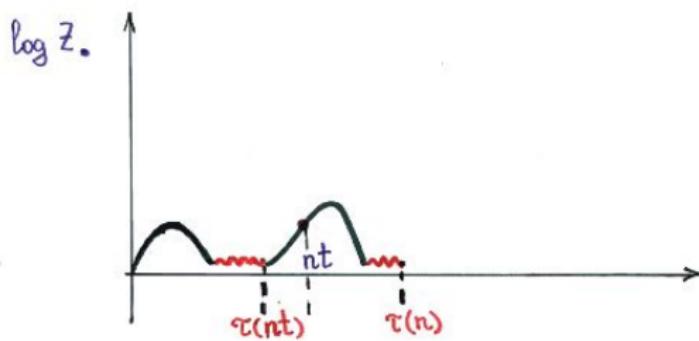
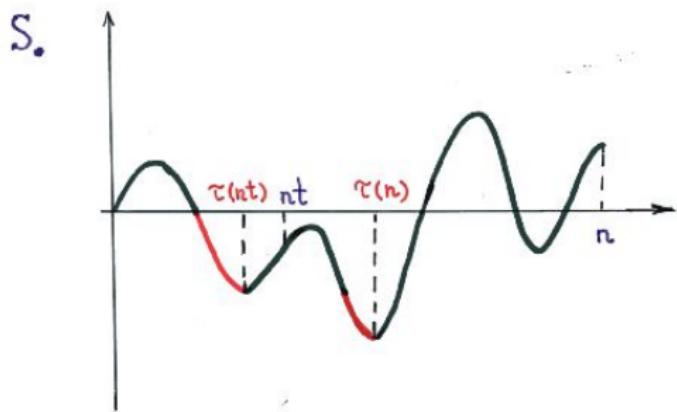
### Theorem 1. (V. + Dyakonova)

For any  $t \in (0, 1)$ , any  $x > 0$  as  $n \rightarrow \infty$

$$\mathbf{P}_f \left( \frac{Z_{nt}}{\mathbf{E}_f [Z_{nt} | Z_{nt} > 0]} \leq x \middle| T > n; nt < \tau(n) \right) \xrightarrow{d} \int_0^x ye^{-y} dy.$$

Non rigorously: if  $T > n, nt < \tau(n) \Rightarrow$

$$\ln Z_{nt} \asymp S_{nt} - S_{\tau(nt)}, n \rightarrow \infty.$$



### Theorem 1. (V. + Dyakonova)

For any  $t \in (0, 1)$ , any  $x > 0$  as  $n \rightarrow \infty$

$$\mathbf{P}_f\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt} > 0]} \leq x \middle| T > n; nt < \tau(n)\right) \xrightarrow{d} \int_0^x ye^{-y} dy.$$

Non rigorously: if  $T > n$ ,  $nt < \tau(n) \Rightarrow$

$$\ln Z_{nt} \asymp S_{nt} - S_{\tau(nt)}, \quad n \rightarrow \infty.$$

**BOTTLENECKS!!!** earlier the time-point  $\tau(n)$  of the global minimum of  $S_k$  within the interval  $[0, n]$ .

**WHAT HAPPENS WITH THE PROCESS AFTER THE POINT OF  
GLOBAL MINIMUM??**

## THE NUMBER OF PARTICLES AFTER THE POINT OF GLOBAL MINIMUM OF $S_k, 0 \leq k \leq n$

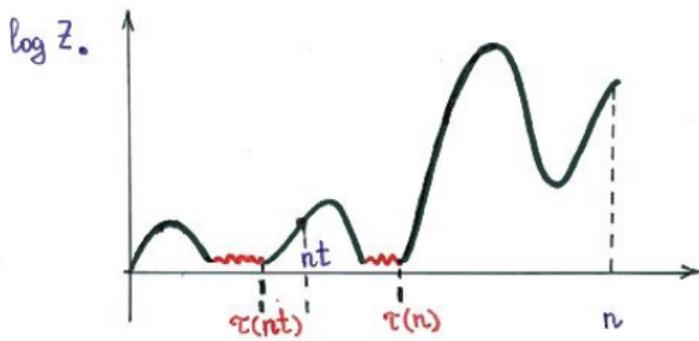
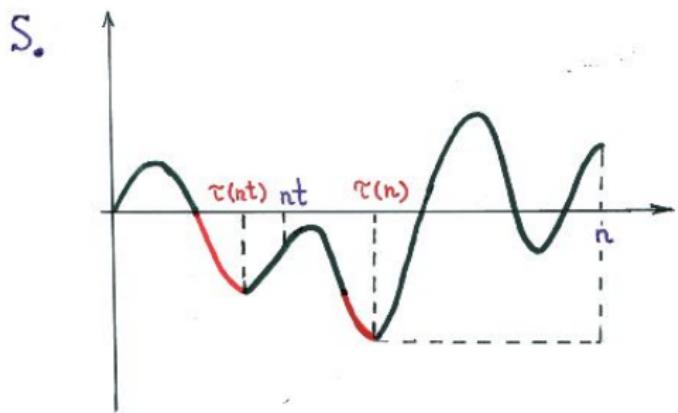
Theorem 2. (V. + Dyakonova)

For any  $t \in (0, 1)$ , any  $x > 0$  as  $n \rightarrow \infty$

$$\mathbf{P}_f\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt} > 0]} \leq x \middle| T > n, \tau(n) \leq nt\right) \xrightarrow{p} 1 - e^{-x}.$$

Non rigorously: if  $T > n, \tau(n) \leq nt \Rightarrow$

$$\ln Z_{nt} \asymp S_{nt} - S_{\tau(n)}.$$



## THE NUMBER OF PARTICLES AFTER THE POINT OF GLOBAL MINIMUM OF $S_k, 0 \leq k \leq n$

Theorem 2. (V. + Dyakonova)

For any  $t \in (0, 1)$ , any  $x > 0$  as  $n \rightarrow \infty$

$$\mathbf{P}_f\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt} > 0]} \leq x \middle| T > n; \tau(n) \leq nt\right) \xrightarrow{d} 1 - e^{-x}.$$

Non rigorously: if  $T > n, \tau(n) \leq nt \Rightarrow$

$$\ln Z_{nt} \asymp S_{nt} - S_{\tau(n)}.$$

$\Rightarrow$  NO tragic bottlenecks AFTER the moment  $\tau(n)$ .

What happens at bottlenecks if  $T > n$ ??

## What happens at bottlenecks if $T > n$ ??

If the probability generating functions are fractional linear then as  $n \rightarrow \infty$

$$\mathbf{P}_f \left( Z_{\tau(nt)} = k \mid T > n; nt < \tau(n) \right) \xrightarrow{d} kq^2 p^{k-1}, \quad k = 1, 2, \dots$$

and

$$\mathbf{P}_f \left( Z_{\tau(n)} = k \mid T > n \right) \xrightarrow{d} \tilde{q}\tilde{p}^{k-1}, \quad k = 1, 2, \dots$$

where  $q + p = 1$ ,  $\tilde{q} + \tilde{p} = 1$  and  $p, \tilde{p}$  are positive random variables.

## BPRE with two types.

Let  $\xi^{(1)}, \xi^{(2)}, \dots$ ; are i.i.d. random variables with (random) probability generating functions

$$f_n(s) := \mathbf{E} \left[ s^{\xi^{(n)}} | f_n \right], \quad 0 \leq s \leq 1$$

$\eta^{(1)}, \eta^{(2)}, \dots$  are i.i.d. random variables with (random) probability generating functions

$$g_n(s) := \mathbf{E} \left[ s^{\eta^{(n)}} | g_n \right], \quad 0 \leq s \leq 1.$$

## BPRE with two types.

Let  $\xi^{(1)}, \xi^{(2)}, \dots$ ; are i.i.d. random variables with (random) probability generating functions

$$f_n(s) := \mathbf{E} \left[ s^{\xi^{(n)}} | f_n \right], \quad 0 \leq s \leq 1$$

$\eta^{(1)}, \eta^{(2)}, \dots$  are i.i.d. random variables with (random) probability generating functions

$$g_n(s) := \mathbf{E} \left[ s^{\eta^{(n)}} | g_n \right], \quad 0 \leq s \leq 1.$$

Define a **pure decomposable two-types** BPRE  $(Z_n, Y_n)$  as

$$Z_n := \sum_{j=1}^{Z_{n-1}} \xi_j^{(n)}, \quad Y_n := \sum_{j=1}^{Y_{n-1}} \eta_j^{(n)}$$

where, given the environment,  $Z_n$  and  $Y_n$  are ordinary inhomogeneous Galton-Watson branching processes.

## BPRE with two types.

Define a **pure decomposable two-types** BPRE  $(Z_n, Y_n)$  as

$$Z_n := \sum_{j=1}^{Z_{n-1}} \xi_j^{(n)}, \quad Y_n := \sum_{j=1}^{Y_{n-1}} \eta_j^{(n)}$$

where, given the environment,  $Z_n$  and  $Y_n$  are ordinary inhomogeneous Galton-Watson branching processes.

Let

$$X_k := \log \mathbf{E} [\xi^{(k)} | f_k], \quad W_k := \log \mathbf{E} [\eta^{(k)} | g_k].$$

BPRE with two types.

Define a **pure decomposable two-types** BPRE  $(Z_n, Y_n)$  as

$$Z_n := \sum_{j=1}^{Z_{n-1}} \xi_j^{(n)}, \quad Y_n := \sum_{j=1}^{Y_{n-1}} \eta_j^{(n)}$$

where, given the environment,  $Z_n$  and  $Y_n$  are ordinary inhomogeneous Galton-Watson branching processes.

Let

$$X_k := \log \mathbf{E} [\xi^{(k)} | f_k], \quad W_k := \log \mathbf{E} [\eta^{(k)} | g_k].$$

**BASIC ASSUMPTION: Asynchronous environments (Predator ( $Z$ )-Pray ( $Y$ ) competition).** For any  $k = 0, 1, 2, \dots$

$$W_{k+1} = -X_k \quad \mathbf{P}-\text{a.s.}$$

Introduce two random walks

$$S_0 := 0, \quad S_n := \sum_{k=1}^n X_k, \quad R_0 := 0, \quad R_n := \sum_{k=2}^{n+1} W_k.$$

Clearly,  $S_n = -R_n, n = 0, 1, \dots$

Introduce two random walks

$$S_0 := 0, \quad S_n := \sum_{k=1}^n X_k, \quad R_0 := 0, \quad R_n := \sum_{k=2}^{n+1} W_k.$$

Clearly,  $S_n = -R_n, n = 0, 1, \dots$

### CRITICALITY.

The random walk  $\{S_n, n \geq 0\}$  is oscillating.

$\Rightarrow$  The branching processes  $Z(n), Y(n)$  are critical.

Denote

$$\zeta_n := \frac{\mathbf{P}_f(Z_n > 0)}{e^{S_{\tau(n)}}}, \quad \theta_n := \frac{\mathbf{P}_g(Y_n > 0)}{e^{R_{\tau(n)}}} = e^{S_{\mu(n)}} \mathbf{P}_g(Y_n > 0)$$

Denote

$$\zeta_n := \frac{\mathbf{P}_f(Z_n > 0)}{e^{S_{\tau(n)}}}, \quad \theta_n := \frac{\mathbf{P}_g(Y_n > 0)}{e^{R_{\tau(n)}}} = e^{S_{\mu(n)}} \mathbf{P}_g(Y_n > 0)$$

### Theorem

(V.+Liu). As  $n \rightarrow \infty$

$$(\zeta_n, \theta_n) \xrightarrow{d} (\zeta, \theta),$$

where the random variables  $\zeta$  and  $\theta$  are **independent** and take values in  $(0, 1]$  with probability 1.

Denote

$$\zeta_n := \frac{\mathbf{P}_f(Z_n > 0)}{e^{S_{\tau(n)}}}, \quad \theta_n := \frac{\mathbf{P}_g(Y_n > 0)}{e^{R_{\tau(n)}}} = e^{S_{\mu(n)}} \mathbf{P}_g(Y_n > 0)$$

### Theorem

(V.+Liu). As  $n \rightarrow \infty$

$$(\zeta_n, \theta_n) \xrightarrow{d} (\zeta, \theta),$$

where the random variables  $\zeta$  and  $\theta$  are **independent** and take values in  $(0, 1]$  with probability 1.

⇒

$$\mathbf{P}_f(Z_n > 0) \asymp e^{S_{\tau(n)}}, \quad \mathbf{P}_g(Y_n > 0) \asymp e^{R_{\tau(n)}} = e^{-S_{\mu(n)}}$$

NOW

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

NOW

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

$$\mathbf{E}_g [Y_{nt} | Y_{nt} > 0] = \frac{e^{R_{nt}}}{\mathbf{P}_g (Y_{nt} > 0)} \asymp e^{R_{nt} - R_{\tau(nt)}} = e^{S_{\mu(nt)} - S_{nt}}$$

NOW

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

$$\mathbf{E}_g [Y_{nt} | Y_{nt} > 0] = \frac{e^{R_{nt}}}{\mathbf{P}_g (Y_{nt} > 0)} \asymp e^{R_{nt} - R_{\tau(nt)}} = e^{S_{\mu(nt)} - S_{nt}}$$

For  $t \in (0, 1]$  consider

$$\mathcal{L}_{f,g} \left( \frac{Z_{nt}}{\mathbf{E}_f [Z_{nt} | Z_{nt} > 0]} \leq z, \frac{Y_{nt}}{\mathbf{E}_g [Y_{nt} | Y_{nt} > 0]} \leq y \middle| Z_n > 0, Y_n > 0 \right)$$

NOW

$$\mathbf{E}_f [Z_{nt} | Z_{nt} > 0] = \frac{e^{S_{nt}}}{\mathbf{P}_f (Z_{nt} > 0)} \asymp e^{S_{nt} - S_{\tau(nt)}}$$

$$\mathbf{E}_g [Y_{nt} | Y_{nt} > 0] = \frac{e^{R_{nt}}}{\mathbf{P}_g (Y_{nt} > 0)} \asymp e^{R_{nt} - R_{\tau(nt)}} = e^{S_{\mu(nt)} - S_{nt}}$$

For  $t \in (0, 1]$  consider

$$\begin{aligned} & \mathcal{L}_{f,g} \left( \frac{Z_{nt}}{\mathbf{E}_f [Z_{nt} | Z_{nt} > 0]} \leq z, \frac{Y_{nt}}{\mathbf{E}_g [Y_{nt} | Y_{nt} > 0]} \leq y \middle| Z_n > 0, Y_n > 0 \right) \\ &= \mathcal{L}_f \left( \frac{Z_{nt}}{\mathbf{E}_f [Z_{nt} | Z_{nt} > 0]} \leq z \middle| Z_n > 0 \right) \times \mathcal{L}_g \left( \frac{Y_{nt}}{\mathbf{E}_g [Y_{nt} | Y_{nt} > 0]} \leq y \middle| Y_n > 0 \right) \end{aligned}$$

For  $t \in (0, 1]$  we have **(V.+Liu)**

$$\begin{aligned} & \mathcal{L}_{f,g}\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt}>0]} \leq z, \frac{Y_{nt}}{\mathbf{E}_g[Y_{nt}|Y_{nt}>0]} \leq y \middle| Z_n>0, Y_n>0\right) \\ & \xrightarrow{p} (1-e^{-z}, 1-e^{-y}) I\{\max\{\tau_1, \mu_1\} < t\} \end{aligned}$$

For  $t \in (0, 1]$  we have **(V.+Liu)**

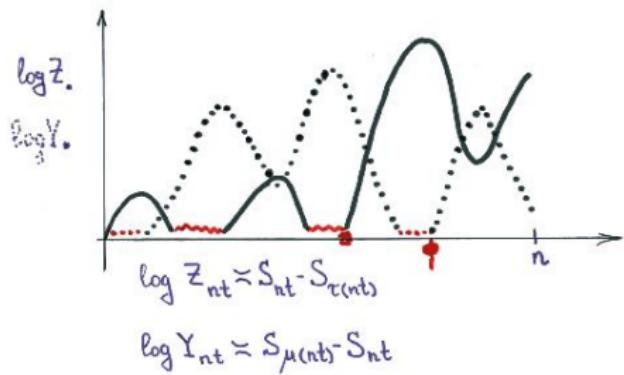
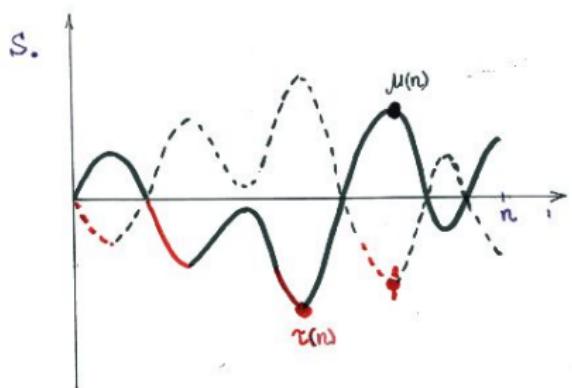
$$\begin{aligned} & \mathcal{L}_{f,g}\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt}>0]} \leq z, \frac{Y_{nt}}{\mathbf{E}_g[Y_{nt}|Y_{nt}>0]} \leq y \middle| Z_n > 0, Y_n > 0\right) \\ & \quad \xrightarrow{p} (1 - e^{-z}, 1 - e^{-y}) I\{\max\{\tau_1, \mu_1\} < t\} \\ & \quad + \left(1 - e^{-z}, \int_0^y we^{-w} dw\right) I\{\tau_1 < t < \mu_1\} \end{aligned}$$

For  $t \in (0, 1]$  we have (V.+Liu)

$$\begin{aligned} & \mathcal{L}_{f,g}\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt}>0]} \leq z, \frac{Y_{nt}}{\mathbf{E}_g[Y_{nt}|Y_{nt}>0]} \leq y \mid Z_n > 0, Y_n > 0\right) \\ & \stackrel{p}{\rightarrow} (1 - e^{-z}, 1 - e^{-y}) I\{\max\{\tau_1, \mu_1\} < t\} \\ & + \left(1 - e^{-z}, \int_0^y we^{-w} dw\right) I\{\tau_1 < t < \mu_1\} \\ & + \left(\int_0^z we^{-w} dw, 1 - e^{-y}\right) I\{\mu_1 < t < \tau_1\} \end{aligned}$$

For  $t \in (0, 1]$  we have (**V.+Liu**)

$$\begin{aligned} & \mathcal{L}_{f,g}\left(\frac{Z_{nt}}{\mathbf{E}_f[Z_{nt}|Z_{nt}>0]} \leq z, \frac{Y_{nt}}{\mathbf{E}_g[Y_{nt}|Y_{nt}>0]} \leq y \middle| Z_n > 0, Y_n > 0\right) \\ & \stackrel{p}{\rightarrow} (1 - e^{-z}, 1 - e^{-y}) I\{\max\{\tau_1, \mu_1\} < t\} \\ & + \left(1 - e^{-z}, \int_0^y we^{-w}dw\right) I\{\tau_1 < t < \mu_1\} \\ & + \left(\int_0^z we^{-w}dw, 1 - e^{-y}\right) I\{\mu_1 < t < \tau_1\} \\ & + \left(\int_0^z we^{-w}dw, \int_0^y we^{-w}dw\right) I\{\min\{\tau_1, \mu_1\} > t\} \end{aligned}$$



WHAT HAPPENS AT THE POINTS OF **LOCAL MINIMA** OF THE  
ASSOCIATED RANDOM WALK ON THE INTERVAL

$[0, n]$

??

For  $k = 1, 2, \dots$  and  $y > 0$  let

$$\mathcal{L}_{f,g} \left( Z_{\tau(nt)} = k, \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Z_n > 0, Y_n > 0 \right)$$

For  $k = 1, 2, \dots$  and  $y > 0$  let

$$\begin{aligned} & \mathcal{L}_{f,g} \left( Z_{\tau(nt)} = k, \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Z_n > 0, Y_n > 0 \right) \\ &= \mathcal{L}_f (Z_{\tau(nt)} = k | Z_n > 0) \times \mathcal{L}_g \left( \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Y_n > 0 \right) \end{aligned}$$

For  $k = 1, 2, \dots$  and  $y > 0$  we have **(V.+Liu)**

$$\begin{aligned} & \mathcal{L}_{f,g} \left( Z_{\tau(nt)} = k, \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Z_n > 0, Y_n > 0 \right) \\ & \xrightarrow{p} (\tilde{q}\tilde{p}^{k-1}, 1 - e^{-y}) I \{ \mu_1 < \tau_t = \tau_1 \} \end{aligned}$$

For  $k = 1, 2, \dots$  and  $y > 0$  we have **(V.+Liu)**

$$\begin{aligned} & \mathcal{L}_{f,g} \left( Z_{\tau(nt)} = k, \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Z_n > 0, Y_n > 0 \right) \\ & \xrightarrow{p} \left( \tilde{q}\tilde{p}^{k-1}, 1 - e^{-y} \right) I \{ \mu_1 < \tau_t = \tau_1 \} \\ & + \left( \tilde{q}\tilde{p}^{k-1}, \int_0^y w e^{-w} dw \right) I \{ \mu_1 > \tau_t = \tau_1 \} \end{aligned}$$

For  $k = 1, 2, \dots$  and  $y > 0$  we have **(V.+Liu)**

$$\begin{aligned} & \mathcal{L}_{f,g} \left( Z_{\tau(nt)} = k, \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Z_n > 0, Y_n > 0 \right) \\ & \xrightarrow{p} \left( \tilde{q}\tilde{p}^{k-1}, 1 - e^{-y} \right) I \{ \mu_1 < \tau_t = \tau_1 \} \\ & + \left( \tilde{q}\tilde{p}^{k-1}, \int_0^y w e^{-w} dw \right) I \{ \mu_1 > \tau_t = \tau_1 \} \\ & + \left( k\tilde{q}^2\tilde{p}^{k-1}, 1 - e^{-y} \right) I \{ \mu_1 < \tau_t < \tau_1 \} \end{aligned}$$

For  $k = 1, 2, \dots$  and  $y > 0$  we have **(V.+Liu)**

$$\begin{aligned} & \mathcal{L}_{f,g} \left( Z_{\tau(nt)} = k, \frac{Y_{\tau(nt)}}{\mathbf{E}_g [Y_{\tau(nt)} | Y_{\tau(nt)} > 0]} \leq y | Z_n > 0, Y_n > 0 \right) \\ & \stackrel{p}{\rightarrow} \left( \tilde{q}\tilde{p}^{k-1}, 1 - e^{-y} \right) I \{ \mu_1 < \tau_t = \tau_1 \} \\ & + \left( \tilde{q}\tilde{p}^{k-1}, \int_0^y w e^{-w} dw \right) I \{ \mu_1 > \tau_t = \tau_1 \} \\ & + \left( k\tilde{q}^2\tilde{p}^{k-1}, 1 - e^{-y} \right) I \{ \mu_1 < \tau_t < \tau_1 \} \\ & + \left( k\tilde{q}^2\tilde{p}^{k-1}, \int_0^y w e^{-w} dy \right) I \{ \tau_t < \min \{ \tau_1, \mu_1 \} \} \end{aligned}$$

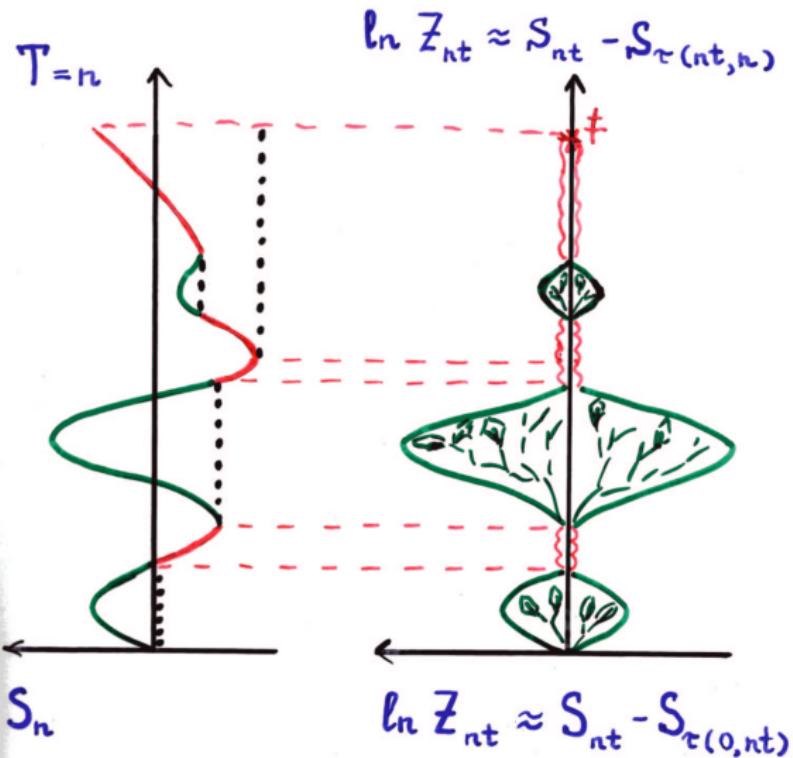
Variations:

1)

$$Y_n > 0, \quad Z_{n-1} = 0, \quad Z_n = 0;$$

2) Correlation  $\in (-1, 1)$ ;

3) Several types of individuals ( $> 2$ ).



THANKS!