

Conditional Density Method in the Computation of the Delta with Application to Power Market

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- In this presentation, we are concerned about the derivative of expectation functionals of the form

$$F(x) = \mathbb{E} [f(x + Y)] , \quad (1)$$

which can be represented as

$$F'(x) = \mathbb{E} [f(x + Y)\pi] . \quad (2)$$

- The random variable π is called a sensitivity weight.

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Density method

- The density method was introduced by Broadie and Glasserman (1996).
- If Y has a differentiable density p_Y , then

$$\begin{aligned} F'(x) &= \frac{d}{dx} \int_{\mathbb{R}} f(y) p_Y(y-x) dy \\ &= \int_{\mathbb{R}} f(y) (-1) p'_Y(y-x) dy \\ &= \mathbb{E} [f(x+Y) (-\partial_y \ln p_Y(Y))] , \end{aligned}$$

- Note no differentiability of f .

- Another case which is important in our analysis will be the situation where we have two independent random variables Y and Z .

Consider the functional

$$F(x) = \mathbb{E} [f(x + Y + Z)] .$$

- In this case we can use the density method with respect to $Y + Z$ which leads to

$$F'(x) = \mathbb{E} [f(x + Y + Z)(-\partial_{y+z} \ln p_{Y+Z}(Y + Z))] .$$

Conditional density method

Conditioning on Z , leads to

$$\begin{aligned} F'(x) &= \frac{d}{dx} \mathbb{E} \left[\mathbb{E}[x + Y + Z | Z] \right] \\ &= \frac{d}{dx} \int_{\mathbb{R}} \mathbb{E}[f(y - Z)] p_Y(y - x) dy \\ &= \mathbb{E} [f(x + Y + Z)(-\partial_y \ln p_Y(Y))] . \end{aligned}$$

Otherwise, conditioning on Y , leads to

$$F'(x) = \mathbb{E} [f(x + Y + Z)(-\partial_z \ln p_Z(Z))] .$$

We essentially consider a Lévy process $(X_t)_{t \geq 0}$ of the form

$$X_t = at + bW_t + N_t + \lim_{\varepsilon \rightarrow 0} \tilde{N}_t^\varepsilon, \quad (3)$$

where the convergence is almost sure and uniform in $t \in [0, T]$.

Here $a, b \in \mathbb{R}$,

$$N_t = \sum_{s \in [0, t]} \Delta X_s \mathbf{1}_{\{|\Delta X_s| \geq 1\}}$$

is a compound Poisson process,

$$\tilde{N}_t^\varepsilon = \sum_{s \in [0, t]} \Delta X_s \mathbf{1}_{\{\varepsilon \leq |\Delta X_s| < 1\}} - t \int_{\varepsilon \leq |z| < 1} zQ(dz)$$

is the compensated Poisson process and Q is the Lévy measure of $(X_t)_{t \geq 0}$.

The components $(W_t)_{t \geq 0}$, $(N_t)_{t \geq 0}$ and $(\tilde{N}_t^\varepsilon)_{t \geq 0}$ are independent of each other.

Proposition

Consider a process $(X_t^\varepsilon)_{t \geq 0}$ defined by

$$X_t^\varepsilon = at + bW_t + \sigma(\varepsilon)B_t + N_t + \tilde{N}_t^\varepsilon, \quad (4)$$

where $(B_t)_{t \geq 0}$ is a Brownian motion independent of the other components. Let $(X_t)_{t \geq 0}$ be as defined in (3). Then

$$\lim_{\varepsilon \rightarrow 0} X_t^\varepsilon = X_t \quad \text{the limit is in } L^1 \text{ and } \mathbb{P} - \text{a.s.}$$

Lemma

Under the condition $u\hat{f}(u) \in L^1(\mathbb{R})$ we have

$$\lim_{\varepsilon \rightarrow 0} \frac{\partial}{\partial x} \mathbb{E}[f(x + X_t^\varepsilon)] = \frac{\partial}{\partial x} \mathbb{E}[f(x + X_t)].$$

Example: NIG-distribution

Take $T=1$. We assume that $X(1)$ is an NIG-Lévy process, with parameters α , β , δ and μ , the density is

$$p_{\text{NIG}}(x; \alpha, \beta, \delta, \mu) = \frac{\alpha\delta}{\pi} e^{\delta\sqrt{\alpha^2 - \beta^2} + \beta(x - \mu)} \frac{K_1\left(\alpha\sqrt{\delta^2 + (x - \mu)^2}\right)}{\sqrt{\delta^2 + (x - \mu)^2}}.$$

Here, K_λ is the modified Bessel function of the second order with parameter λ .

Notice that there is no Brownian motion component in the NIG process.

We apply the density method to find a sensitivity weight π :

$$\pi = -\beta + \frac{X(1) - \mu}{\delta^2 + (X(1) - \mu)^2} \left\{ 1 - \frac{\alpha \sqrt{\delta^2 + (X(1) - \mu)^2} K_1' \left(\alpha \sqrt{\delta^2 + (x(1) - \mu)^2} \right)}{K_1 \left(\alpha \sqrt{\delta^2 + (x(1) - \mu)^2} \right)} \right\}.$$

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- The computation of the two modified Bessel functions K_1 and K_2 in order to calculate the sensitivity weight π makes the Monte Carlo simulation technique very inefficient.
- An alternative will then be to use an approximation of $X(t)$ and derive the sensitivity weight with respect to this process. In this case we find from the calculations above

$$\pi = \frac{B(1)}{\sigma(\varepsilon)}.$$

- This will be obviously much simpler to compute.

- We consider a multi-factor spot price process $S(t)$ as follows

$$S(t) = g(t, X_1(t), \dots, X_n(t)).$$

- We consider a payoff function $h : \mathbb{R} \rightarrow \mathbb{R}$ and we suppose that there exist a function f and a differentiable function ζ such that

$$h(S(T)) = f(X_1(T) + \zeta(S(0)), X_2(T), \dots, X_n(T)). \quad (5)$$

- The arbitrage-free price of a european option written on the spot price S is defined as

$$C(S(0)) = e^{-rT} \mathbb{E} [f(X_1(T) + \zeta(S(0)), X_2(T), \dots, X_n(T))] , \quad (6)$$

where r is the risk-free instantaneous interest rate.

- The aim is to study the delta in such a model.

Denote by p_1 the density of $X_1(T)$, which we suppose to be known.

Proposition

Assume that there exists an integrable function u on \mathbb{R} such that

$$|\mathbb{E}[f(x, X_2(T), \dots, X_n(T))]| p_1(x - \zeta(S(0))) \leq u(x). \quad (7)$$

Then

$$\frac{\partial C}{\partial S(0)} = -\zeta'(S(0)) e^{-rT} \mathbb{E}[h(S(T)) \partial \ln p_1(X_1(T))].$$

Example 1

We start with considering the two-factor model of Schwartz and Smith (2000) applied as a model for the oil price dynamics. It takes the form

$$S(t) = S(0) \exp(X(t) + Y(t)), \quad (8)$$

where

$$dX(t) = (\theta - \alpha X(t)) dt + \sigma dW(t), \quad (9)$$

and

$$dY(t) = \mu dt + \eta d\widetilde{W}(t), \quad (10)$$

with independent \widetilde{W} and W .

We represent the payoff as

$$h(S(T)) = h(\exp(\ln(S(0)) + X(T) + Y(T))).$$

Thus

$$f(x + \zeta(S(0)), y) = h(\exp(\ln(S(0)) + x + y))$$

with $\zeta(s) = \ln s$. The mean of $X(t)$ is

$$X(0) \exp(-\alpha t) + \theta(1 - \exp(-\alpha t))/\alpha$$

and the variance is

$$\sigma^2/2\alpha(1 - \exp(-2\alpha t)).$$

Therefore, it holds that

$$\partial \ln p_1(x) = -\frac{1}{\frac{\sigma^2}{2\alpha}(1 - e^{-2\alpha T})} \left(x - X(0)e^{-\alpha T} - \frac{\theta}{\alpha}(1 - e^{-\alpha T}) \right).$$

Hence, the delta is

$$\frac{\partial C}{\partial S(0)} = \frac{e^{-rT} 2\alpha}{S(0)\sigma^2(1 - e^{-2\alpha T})} \mathbb{E} \left[h(S(T)) \left(X_1(T) - X(0)e^{-\alpha T} - \frac{\theta}{\alpha}(1 - e^{-\alpha T}) \right) \right].$$

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Example 2

- We consider the additive model for the electricity spot price proposed by Benth, Kallsen, and Meyer-Brandis (2007) given by

$$S(t) = X(t) + Y(t), \quad S(0) > 0.$$

- The process $Y(t)$ is given by

$$Y(t) = -\lambda_2 Y(t)dt + dL_2(t), \quad Y(0) = 0,$$

where L_2 is a compound Poisson process.

- The process $X(t)$ is a $\Gamma(a, b)$ -OU process. Namely, it is a Lévy process following the dynamics

$$dX(t) = -\lambda_1 X(t)dt + dL_1(t), \quad X(0) = S(0),$$

where $L_1(t)$ is a subordinator, admitting a stationary distribution which is here $\Gamma(a, b)$.

- The solution takes the form

$$X(t) = e^{-\lambda_1 t} X(0) + \int_0^t e^{\lambda_1(s-t)} dL_1(s),$$

- We observe that the stochastic integral

$$\int_0^t e^{\lambda_1(s-t)} dL_1(s) = X(t) - e^{-\lambda_1 t} X(0)$$

has an asymptotic distribution being $\Gamma(a, b)$ when t goes to infinity, since $e^{-\lambda_1 t} X(0)$ goes to 0 when t goes to infinity.

- Denoting by Z a random variable which is $\Gamma(a, b)$ -distributed, we consider

$$\tilde{S}(t) = e^{-\lambda_1 t} X(0) + Z + Y(t),$$

which is asymptotically equal in distribution to $S(t)$.

- Consider the payoff of a call option

$$h(S(T)) = \max(S(T) - K, 0).$$

- Therefore, the expression for the delta is given by

$$\frac{\partial C}{\partial S(0)} \approx e^{-rT} \mathbb{E} \left[h(e^{-\lambda_1 T} X(0) + Z + Y(T)) e^{-\lambda_1 T} \left(b - \frac{a-1}{Z} \right) \right].$$

- To make a numerical example which is relevant for energy markets, we note that Benth, Kiesel, and Nazarova (2009) showed empirically that the spot model fitted the Phelix Base electricity price index at the European Power Exchange (EEX) very well. We used the data from that paper.

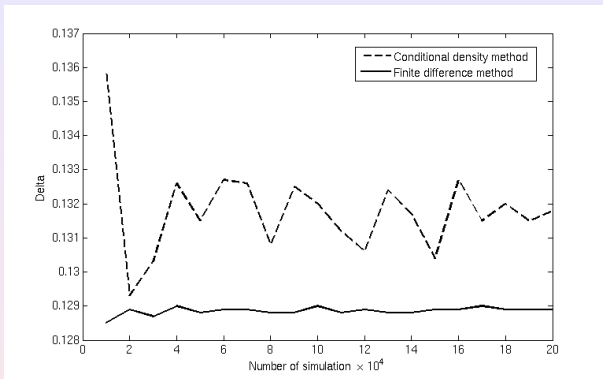


Figure: Simulation of a delta for a call option

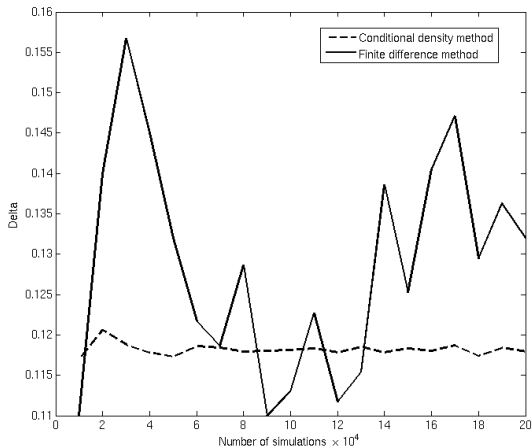







Figure: Simulation of a delta for a digital option

References

-  Benth, F. E., Kiesel, R., and Nazarova, A. (2009).
A critical empirical study of two electricity spot price models.
Manuscript.
-  Benth, F. E., Di Nunno, G., and Khedher, A. (2009).
Lévy-models robustness and sensitivity.
To appear in Volume XXV of the series QP-PQ, Quantum
Probability and White Noise Analysis, H. Ouerdiane and A.
Barhoumi (eds.), World Scientific.
-  Benth, F. E., Kallsen, J., and Meyer-Brandis, T. (2007).
A non-Gaussian Ornstein-Uhlenbeck process for electricity spot
price modeling and derivatives pricing.
Appl. Math. Finance, **14**, pp. 153-169.

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ur-logo

-  Broadie, M., and Glasserman, P. (1996).
Estimating security price derivatives using simulation.
Manag. Science, **42**, pp. 169–285.
-  Schwartz, E.S. and Smith, J. E. (2000).
Short-term variations and long-term dynamics in commodity prices.
Manag. Science, **46**(7), pp. 893–911.

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