## **Derivative Free Optimization**

**Optimization and AMS Masters - University Paris Saclay** 

Exercices - Class 3

Anne Auger anne.auger@inria.fr https://www.lri.fr/~auger/teaching.html

## I On linear convergence

For a deterministic sequence  $x_t$  the linear convergence towards a point  $x^*$  is defined as: The sequence  $(x_t)_t$  convergences linearly towards  $x^*$  if there exists  $\mu \in (0, 1)$  such that

$$\lim_{t \to \infty} \frac{\|x_{t+1} - x^*\|}{\|x_t - x^*\|} = \mu \tag{1}$$

The constant  $\mu$  is then the convergence rate.

We consider a sequence  $(x_t)_t$  that converges linearly towards  $x^*$ .

1. Prove that (1) is equivalent to

$$\lim_{t \to \infty} \ln \frac{\|x_{t+1} - x^*\|}{\|x_t - x^*\|} = \ln \mu$$
(2)

2. Prove that (2) implies

$$\lim_{t \to \infty} \frac{1}{t} \sum_{k=0}^{t-1} \ln \frac{\|x_{k+1} - x^*\|}{\|x_k - x^*\|} = \ln \mu$$
(3)

3. Prove that (3) is equivalent

$$\lim_{t \to \infty} \frac{1}{t} \ln \frac{\|x_t - x^*\|}{\|x_0 - x^*\|} = \ln \mu$$
(4)

For a sequence of random variables  $(x_t)_t$ . We can define linear convergence by considering the expected log progress, that is the sequence converges linearly if

$$\lim_{t \to \infty} E\left[ \ln \frac{\|x_{t+1} - x^*\|}{\|x_t - x^*\|} \right] = \ln \mu \; ,$$

Remark that in general

$$E\left[\ln\frac{\|x_{t+1} - x^*\|}{\|x_t - x^*\|}\right] \neq \ln E\left[\frac{\|x_{t+1} - x^*\|}{\|x_t - x^*\|}\right]$$

and thus defining linear convergence via  $\lim_{t} E\left[\frac{\|x_{t+1}-x^*\|}{\|x_t-x^*\|}\right]$  would not be equivalent contrary to the deterministic case.

If we want to define the almost sure linear convergence we cannot use directly (1) or (2) as  $\frac{\|x_{t+1}-x^*\|}{\|x_t-x^*\|}$  or  $\ln \frac{\|x_{t+1}-x^*\|}{\|x_t-x^*\|}$  will not convergence almost surely to a constant. We therefore have to resort to (5) and define the almost sure linear convergence of a sequence of random variables as

$$\lim_{t \to \infty} \frac{1}{t} \ln \frac{\|x_t - x^*\|}{\|x_0 - x^*\|} = \ln \mu \text{ a.s.}$$
(5)

This will is illustrated as the log-distance to the optimum decreases to minus infinity as  $\ln \mu \times t$ , that is you observe asymptotically a line if you plot the convergence using a log-scale for the *y*-axis.

## **II** Order statistics - Effect of selection

We want to illustrate the effect of selection on the distribution of candidate solutions in a stochastic algorithm. More precisely we consider a  $(1, \lambda)$ -ES algorithm whose state is given by  $X_t \in \mathbb{R}^n$ . At each iteration  $t, \lambda$  candidate solutions are sampled according to

$$X_{t+1}^{i} = X_t + U_{t+1}^{i}$$

with  $(U_{t+1}^i)_{1 \leq i \leq \lambda}$  i.i.d. and  $U_{t+1}^i \sim \mathcal{N}(0, I_d)$ . Those candidate are evaluated on the function  $f : \mathbb{R}^n \to \mathbb{R}$  to be minimized and then ranked according the their f values:

$$f(X_{t+1}^{1:\lambda}) \le \ldots \le f(X_{t+1}^{\lambda:\lambda})$$

where  $i:\lambda$  denotes the index of the  $i^{\text{th}}$  best candidate solution. The best candidate solution is then selected that is

$$X_{t+1} = X_{t+1}^{1:\lambda}$$
 .

We will compute for the linear function  $f(x) = x_1$  to be minimized the conditional distribution of  $X_{t+1}^{1;\lambda}$ (i.e. after selection) and compare it to the distribution of  $X_i^{t+1}$  (i.e. before selection).

1. What is the distribution of  $X_{t+1}^i$  conditional to  $X_t$ ? Deduce the density of each coordinate of  $X_{t+1}^i$ .

We remind that given  $\lambda$  random variables independent and identically distributed  $Y_1, Y_2, \ldots, Y_{\lambda}$ , the order statistics  $Y_{(1)}, Y_{(2)}, \ldots, Y_{(\lambda)}$  are random variables defined by sorting the realizations of  $Y_1, Y_2, \ldots, Y_{\lambda}$  in increasing order. We consider that each random variable  $Y_i$  admits a density f(x) and we denote F(x) the cumulative distribution function, that is  $F(x) = \Pr(Y \leq x)$ .

- 2. Compute the cumulative distribution of  $Y_{(1)}$  and deduce the density of  $Y_{(1)}$ .
- 3. Let  $U_{t+1}^{1:\lambda}$  be the random vector such that

$$X_{t+1}^{1:\lambda} = X_t + U_{t+1}^{1:\lambda}$$

Express for the minimization of the linear function  $f(x) = x_1$ , the first coordinate of  $U_{t+1}^{1:\lambda}$  as an order statistic.

4. Deduce the conditional distribution and conditional density of the random vector  $X_{t+1}^{1:\lambda}$ .

## III Cumulative Step-size Adaptation (CSA)

In this exercice, we want to understand the normalization constants in the CSA algorithm and how they implement the idea explained during the class. The pseudo-code of the  $(\mu/\mu, \lambda)$ -ES with CSA step-size adaption is given in the following.

[Objective: minimize  $f : \mathbb{R}^n \to \mathbb{R}$ ]

1. Initialize  $\sigma_0 > 0$ ,  $\mathbf{m}_0 \in \mathbb{R}^n$ ,  $\mathbf{p}_0 = 0$ , t = 02. set  $w_1 \ge w_2 \ge \dots w_\mu \ge 0$  with  $\sum w_i = 1$ ;  $\mu_{\text{eff}} = 1/\sum w_i^2$ ,  $0 < c_\sigma < 1$  (typically  $c_\sigma \approx 4/n$ ),  $d_\sigma > 0$ 3. while not terminate

- 4. Sample  $\lambda$  independent candidate solutions :
- 5.
- $$\begin{split} \mathbf{X}_{t+1}^{i} &= \mathbf{m}_{t} + \sigma_{t} \mathbf{y}_{t+1}^{i} \quad \text{for } i = 1 \dots \lambda \\ & \text{with } (\mathbf{y}_{t+1}^{i})_{1 \leq i \leq \lambda} \text{ i.i.d. following } \mathcal{N}(\mathbf{0}, I_{d}) \\ \text{Evaluate and rank solutions:} \\ & f(\mathbf{X}_{t+1}^{1:\lambda}) \leq \dots \leq f(\mathbf{X}_{t+1}^{\lambda:\lambda}) \\ \text{Update the mean vector:} \\ & \mu \end{split}$$
  6.
- 7.
- 8.
- 9.

$$\mathbf{m}_{t+1} = \mathbf{m}_t + \sigma_t \underbrace{\sum_{i=1}^{i:\lambda} w_i \mathbf{y}_{t+1}^{i:\lambda}}_{\mathbf{y}_{t+1}^w}$$

12. 
$$\mathbf{p}_{t+1} = (1 - c_{\sigma})\mathbf{p}_t + \sqrt{1 - (1 - c_{\sigma})^2}\sqrt{\mu_{\text{eff}}}\mathbf{y}_{t+1}^w$$

13. Update the step-size:

14. 
$$\sigma_{t+1} = \sigma_t \exp\left(\frac{c_\sigma}{d_\sigma} \left(\frac{\|p_\sigma\|}{E[\|\mathcal{N}(0, I_d)\|]} - 1\right)\right)$$

15. t=t+1

- 1. Assume that the objective function f is random, i.e. for instance  $f(X_{t+1}^i)_i$  are i.i.d. according to  $\mathcal{U}_{[0,1]}$ . What is the distribution of  $\sqrt{\mu_{\text{eff}}} \mathbf{y}_{t+1}^w$ ?
- 2. Assume that  $\mathbf{p}_t \sim \mathcal{N}(0, I_d)$  and that the selection is random, show that  $\mathbf{p}_{t+1} \sim \mathcal{N}(0, I_d)$
- 3. Deduce that under random selection

$$E\left[\ln\sigma_{t+1}|\sigma_t\right] = \ln\sigma_t$$

and then that the expected log step-size is constant.