

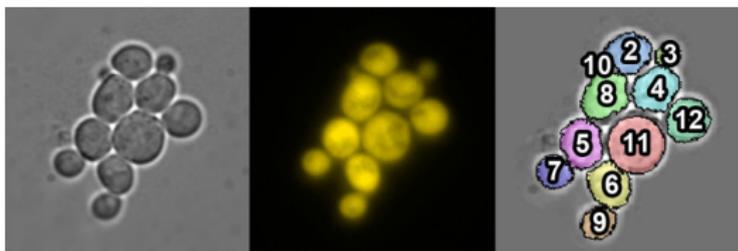
Inheritance and variability of kinetic gene expression parameters in microbial cells: Modelling and inference from lineage tree data

Aline Marguet¹, Marc Lavielle², Eugenio Cinquemani¹.

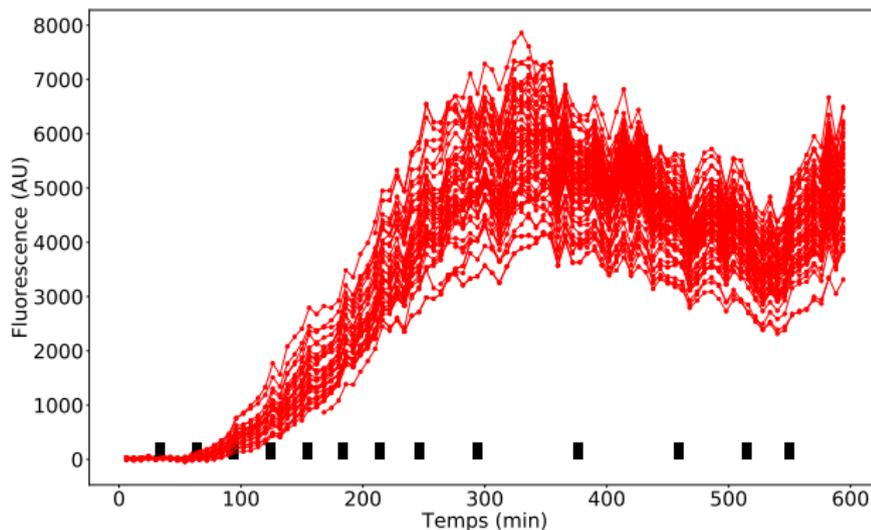
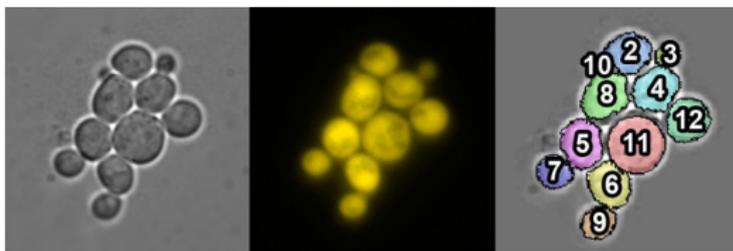
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Gene expression variability



Gene expression variability



Llamosi et al., What population reveals about cell identity: Single-cell parameter estimation of models of gene expression in yeast, 2016, *PLOS Comput. Biol.* 12(2), 2016.

Mathematical modeling of variability

Usually, cells are modeled as independent individuals.

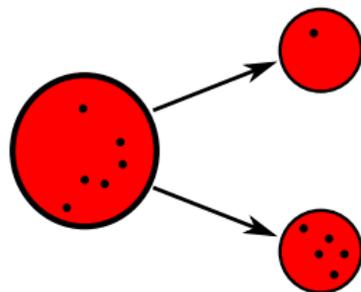
cell division \Rightarrow correlations between mother and daughter cells

Mathematical modeling of variability

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◆ random partitioning

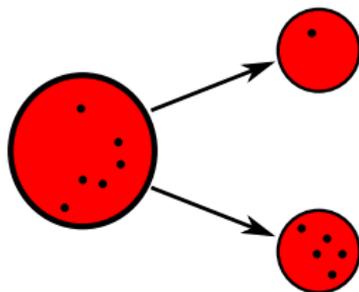


Mathematical modeling of variability

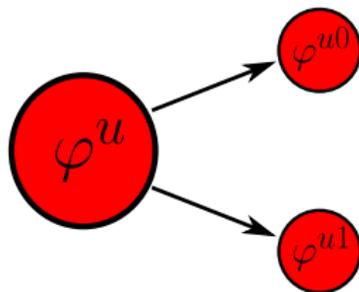
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◆ inheritance

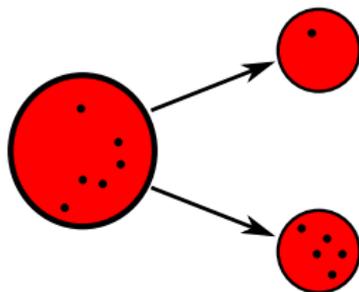


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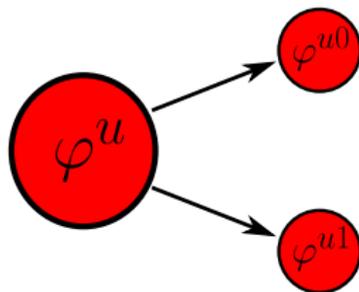
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Questions: ◆ do parameters usually treated as independent across cells show inheritance? ◆ to what extent are these parameters conserved from one generation to the next?

Gene expression modelling over a lineage tree

Identification from lineage tree data

Validation *in silico* of the ARME algorithm

Application to the study of yeast osmotic shock response

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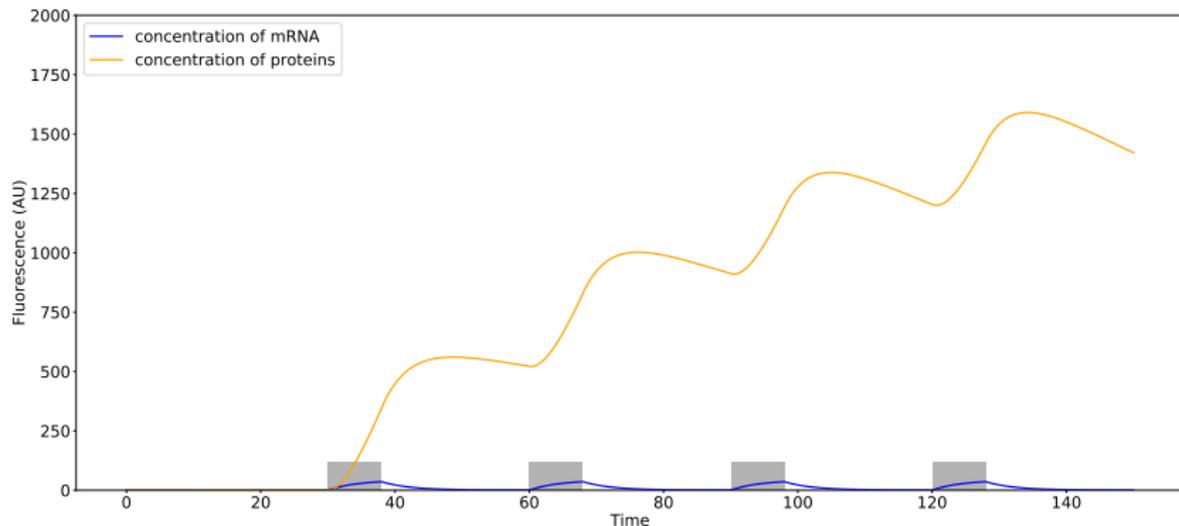
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Dynamical model of gene expression

$$\begin{cases} \dot{m}(t) &= k_m u(t) - g_m m(t) \\ \dot{p}(t) &= k_p m(t) - g_p p(t). \end{cases}$$

- ▶ $u(t)$: activity of transcription factors,
- ▶ k_m, g_m : production and decay rate of the mRNA,
- ▶ k_p, g_p : production and decay rate of the protein.



Individual parameters: mixed-effect modelling

Each cell v in the population has its own parameters

$$\psi_v = (k_m^v, g_m^v, k_p^v, g_p^v)$$

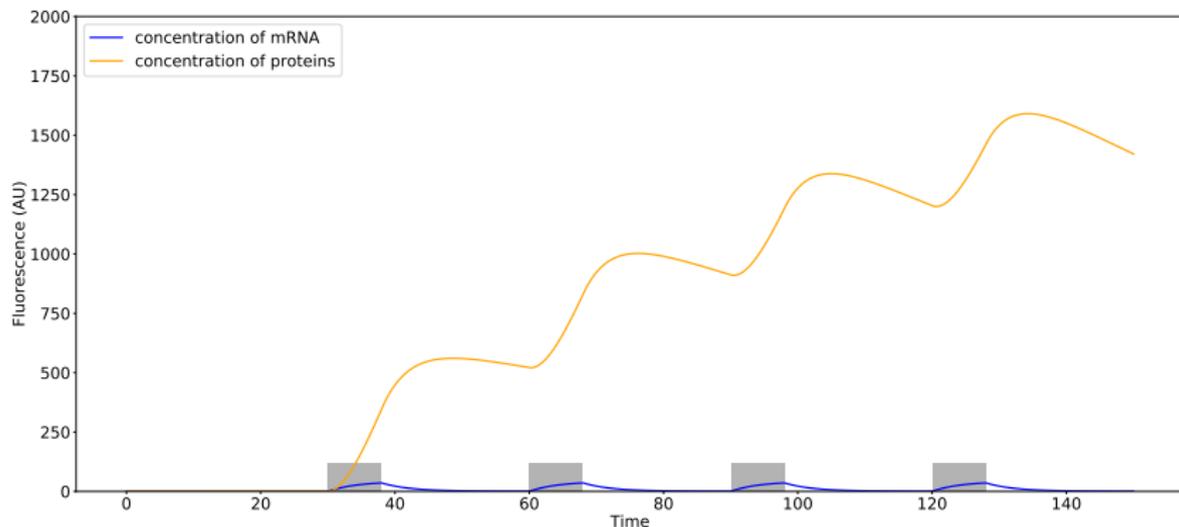
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► Variability in the response to the same stimulus:

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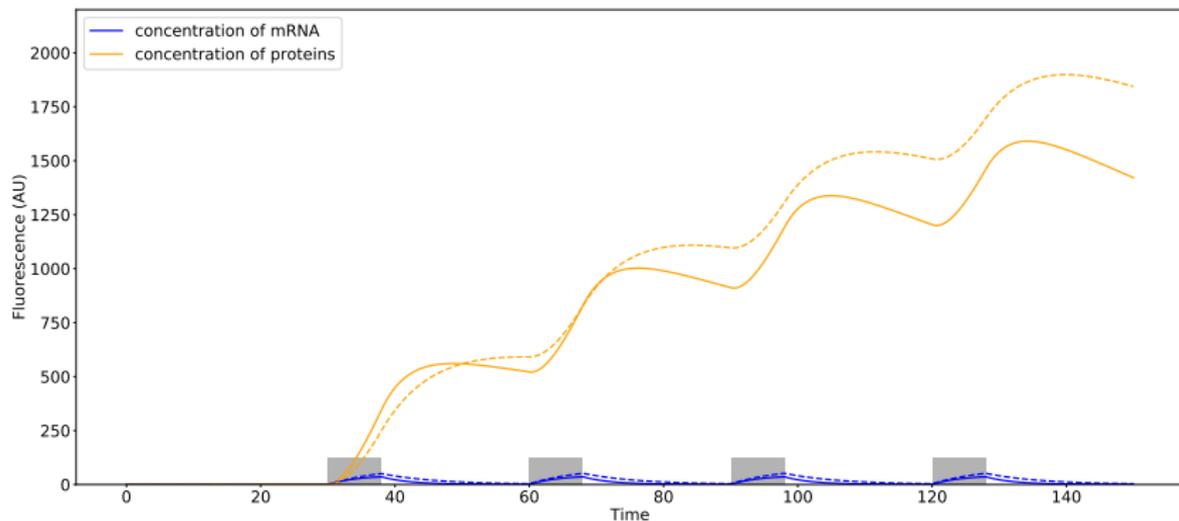
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Modelling inheritance

► Individual parameters:

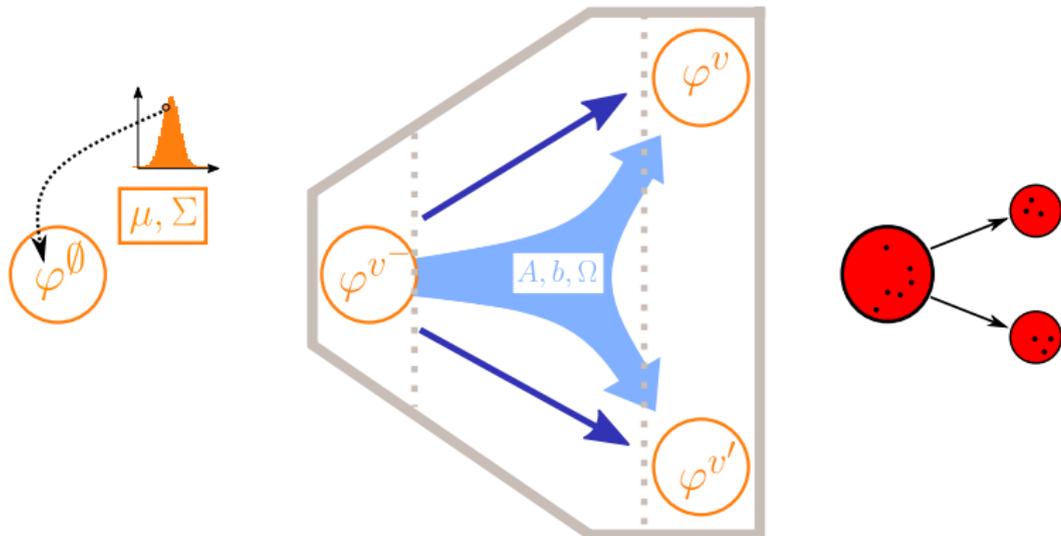
$$\psi^v = (k_m^v, g_m^v, k_p^v, g_p^v) \text{ and } \varphi^v = \log(\psi^v).$$

Modelling inheritance

- ▶ Individual parameters:

$$\psi^v = (k_m^v, g_m^v, k_p^v, g_p^v) \text{ and } \varphi^v = \log(\psi^v).$$

- ▶ Transmission mechanism:



$$\varphi^0 \sim \mathcal{N}(\mu, \Sigma) \\ \Sigma = A \Sigma A^T + \Omega$$

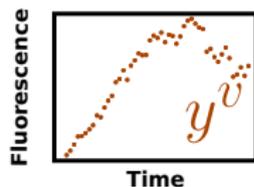
$$\varphi^v = A \varphi^{v-} + (I - A)b + \eta^v \\ \eta^v \sim \mathcal{N}(0, \Omega)$$

equal sharing
at division.

Identification problem

- ▶ Additive noise model for the fluorescence measurements:

$$Y_j^v = p(t_j^v, \psi_v) + h\varepsilon_j^v$$

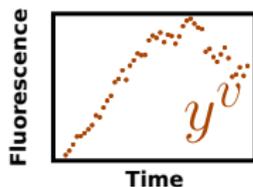


where $h \geq 0$, $(\varepsilon_j^v, j = 1, \dots, n^v, v \in V)$ are i.i.d. $\varepsilon_j^v \sim \mathcal{N}(0, 1)$.

Identification problem

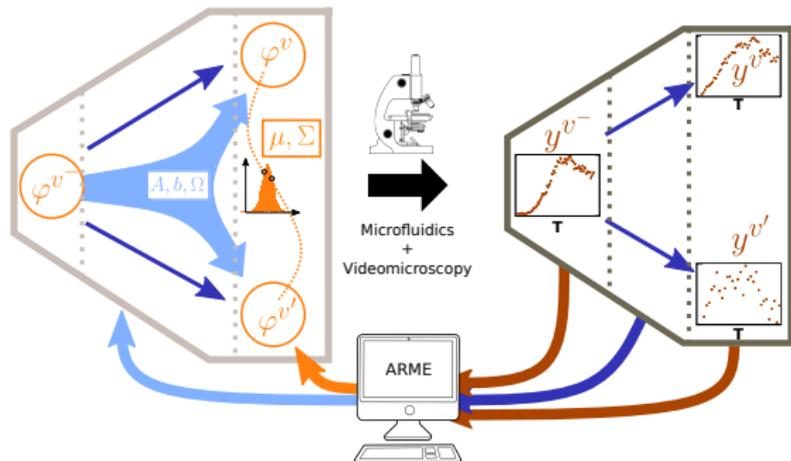
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- ▶ Goal: estimate $\theta = (A, b, \Omega, h)$ from y and lineage informations W .



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ARME algorithm : a generalization of the SAEM algorithm

1. Initialisation : $\theta_0 = \vartheta_0$.
2. For k from 0 to N ,
 - ▶ Simulation step : using MCMC methods, simulate

$$(\varphi_v^k)_{v \geq 0} \sim p((\varphi_v)_{v \geq 0} | Y, W, \theta_k).$$

- ▶ Stochastic approximation step :

$$Q_{k+1}(\theta) = Q_k(\theta) + \gamma_k (\log(p(Y, (\varphi_v^k)_{v \geq 0} | W, \theta) - Q_k(\theta)).$$

- ▶ Maximization step : $\theta_{k+1} = \operatorname{argmax}_{\theta} (Q_{k+1}(\theta))$.

⚠ For the simulation step: dependencies between individuals.



"Convergence of a stochastic approximation version of the EM algorithm." B. Delyon, M. Lavielle, E. Moulines, Ann. Statist. 27 (1999), no. 1, 94–128.

Implementation: detailed simulation step

- ◇ Using Metropolis-Hasting algorithm with several proposal law, simulate

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Update of $\varphi = (\varphi_v)_{v \geq 0}$

- ▶ at the **population level**: takes into account every correlations, very low acceptance rate.

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- ▶ at the **population level**: takes into account every correlations, very low acceptance rate.
- ▶ at the **generation level**: takes into account the correlation with the previous and the next generation, better acceptance rate.
- ▶ at the **individual level**: does not take into account any correlation, adaptative acceptance rate.

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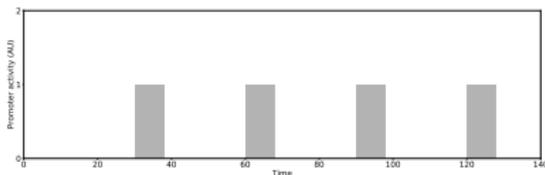
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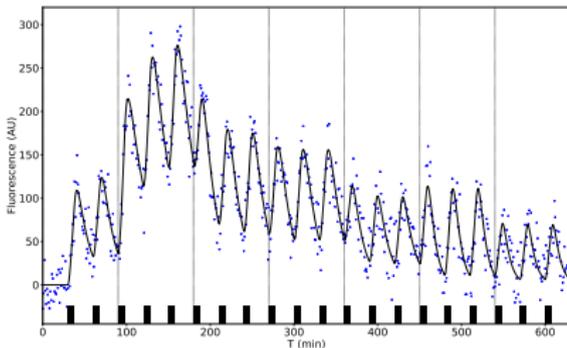
Simulation of 20 datasets with

▶ $u(t) =$

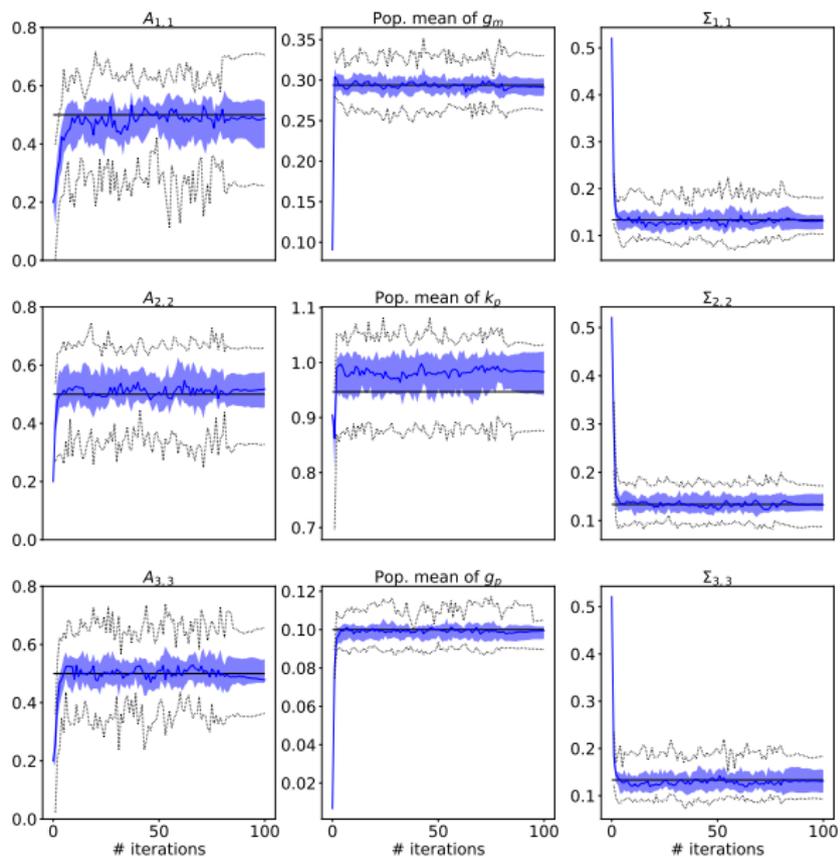


- ▶ 128 individuals (7 generations),
- ▶ a fixed value of k_m , for identifiability reasons,
- ▶ inheritance parameter $A = \text{Diag}(0.5, 0.5, 0.5)$,
- ▶ global mean parameter $b = [\log(0.294), \log(0.947), \log(0.1)]^T$,
- ▶ global covariance parameter $\Omega = \text{Diag}(0.1, 0.1, 0.1)$,

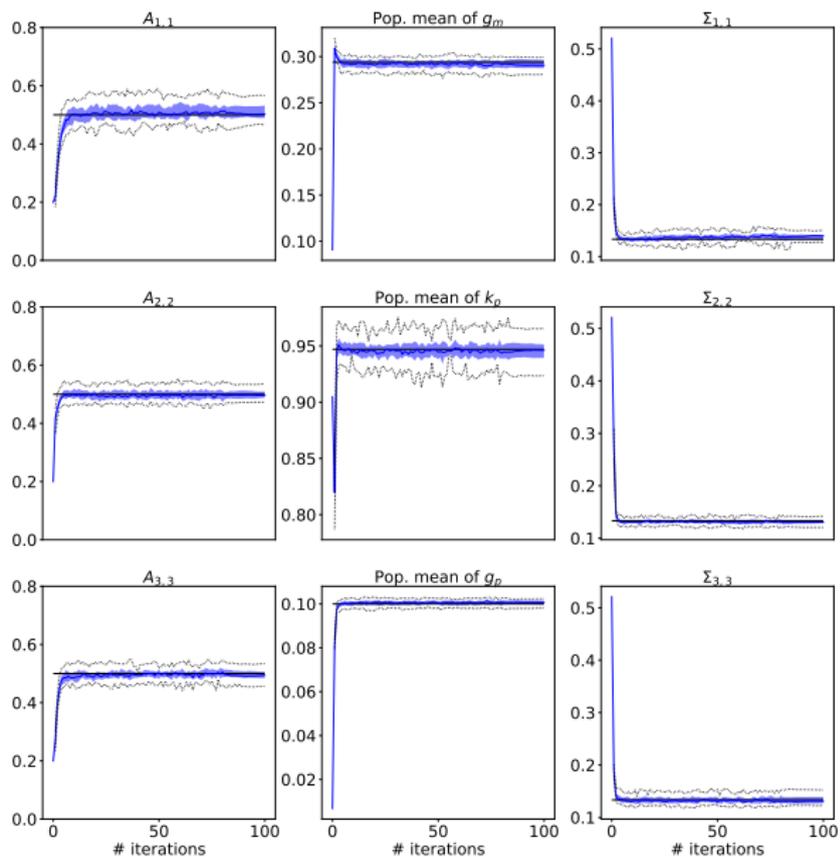
- ▶ noise of measure $h = 20$,



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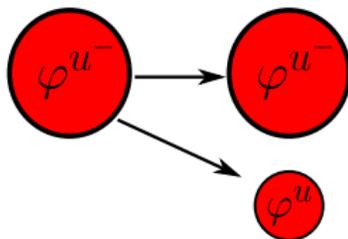
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- ▶ Extended model for gene expression (Llamosi *et al.* (2016)):

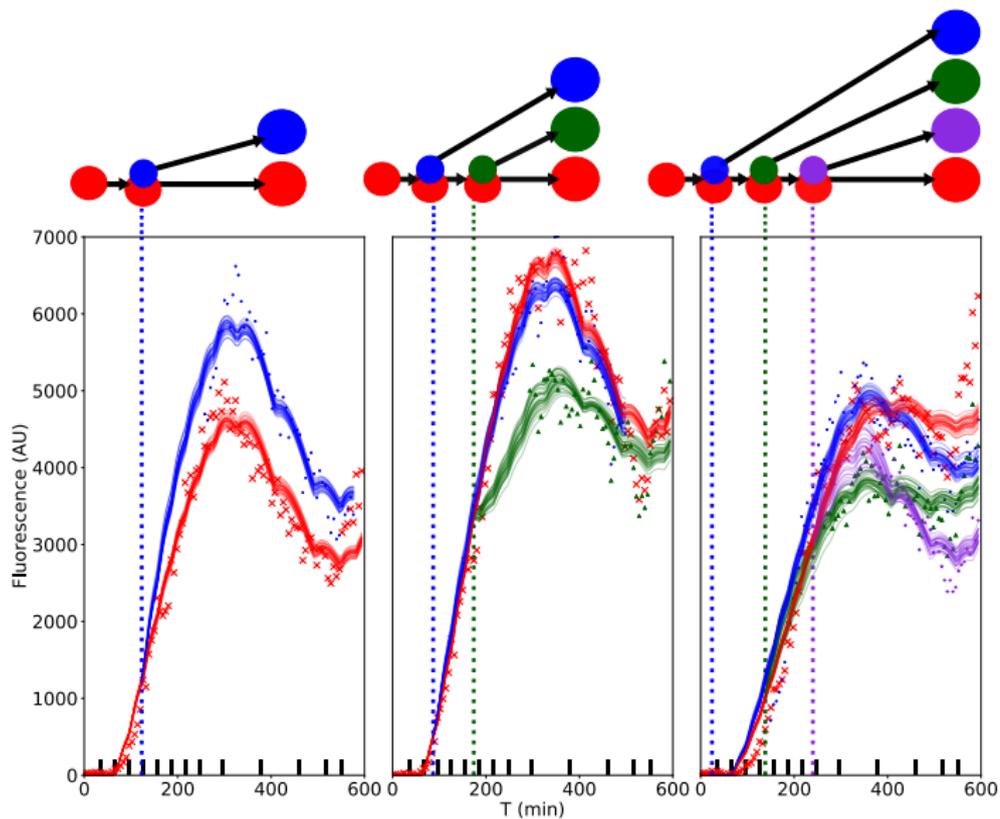
$$\begin{cases} \dot{u}(t) &= k_h u_c(t) - g_h u(t), \\ \dot{m}(t) &= k_m u(t) - g_m m(t), \\ \dot{p}(t) &= k_p m(t) - g_p p(t), \end{cases}$$

with fixed values for k_h and g_h .

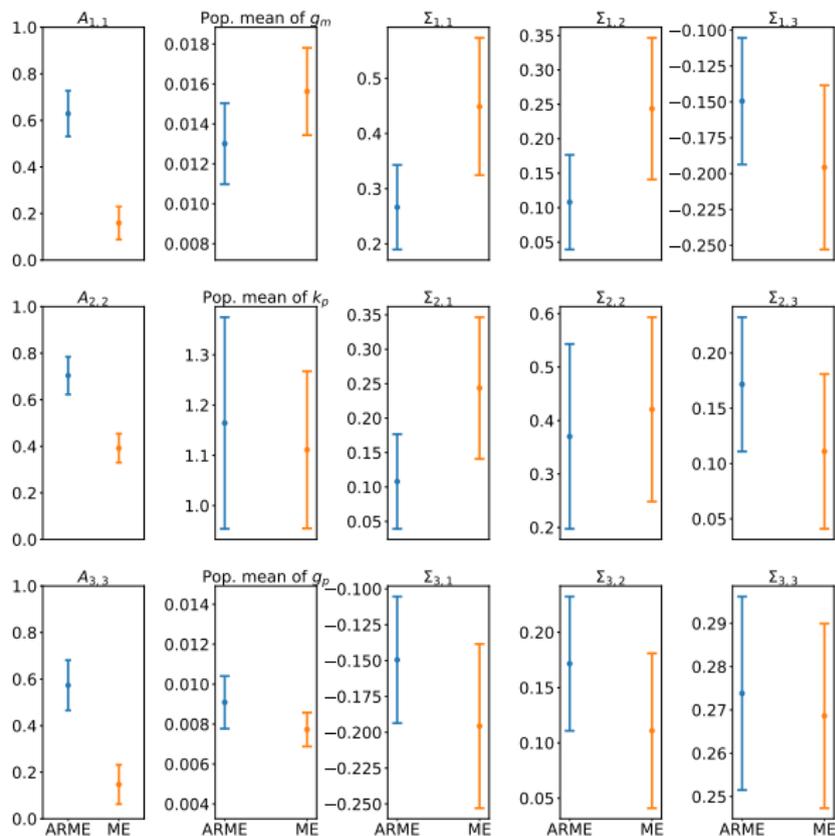
- ▶ Maturation time for reporter molecules.
- ▶ Budding yeast (*S. cerevisiae*): the mother keeps its own kinetic parameter at division.



Single-cell data fits after ARME identification



Results from identification of a ARME model



Conclusions

- ◆ Daughter cell parameters are determined by the mother to an extent as large as 60% (a state-of-the art indirect method assessed this value at 20 – 40%).
- ◆ Indirect methods underestimate inheritance
- ◆ Inheritance is equal for the different parameters: it acts at the level of global regulatory factors (at least for the system and data we examined).

Perspectives

- ◆ Consider intrinsic noise, more complex inheritance models, etc.
- ◆ Proof of the convergence of the algorithm

Conclusions

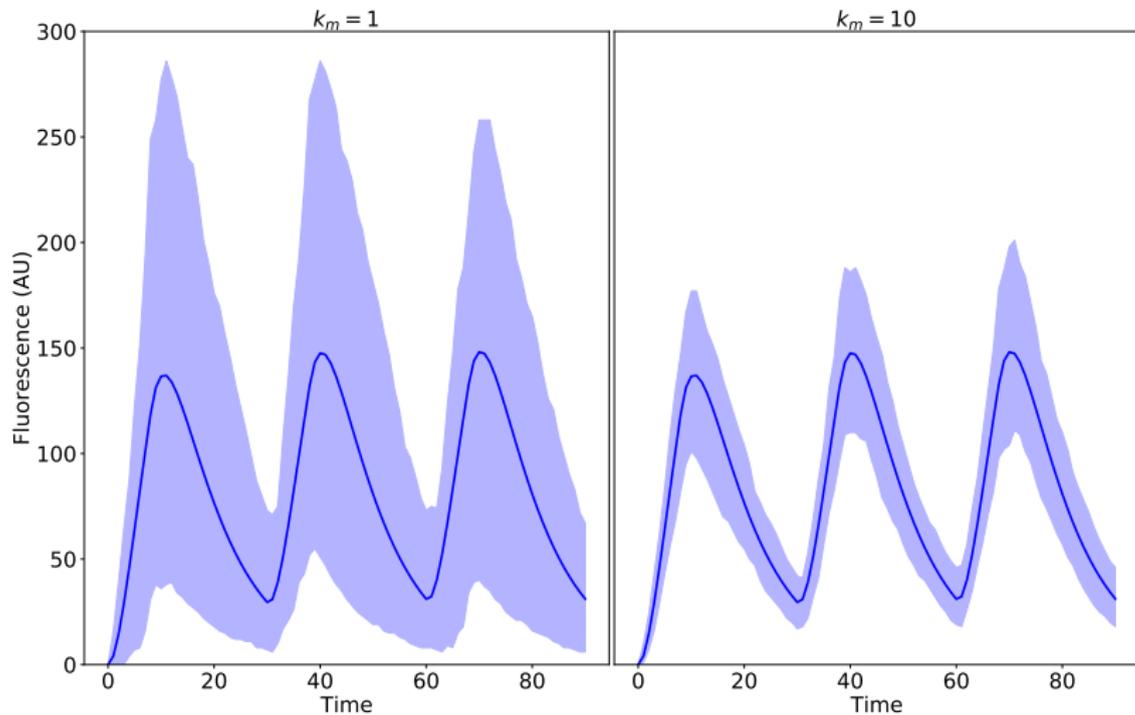
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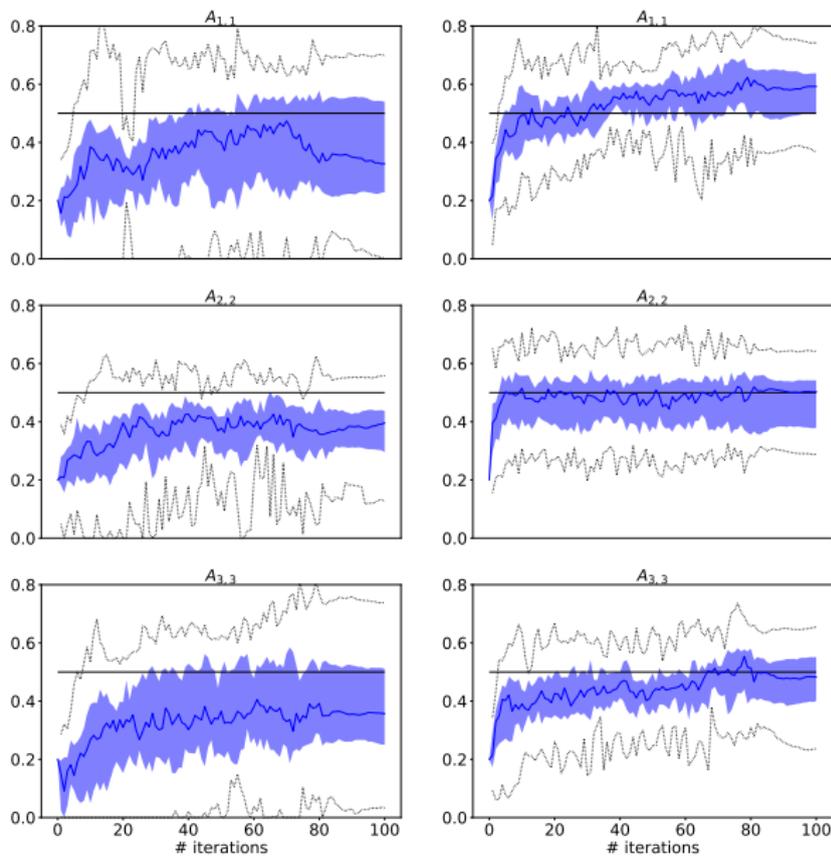
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Thank you for your attention!

Performance in presence of intrinsic noise



Performance in presence of intrinsic noise



Experimental design

