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Lévy processes

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Lévy Processes

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Some books

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Give us the tools, and we will finish the work.
Winston Churchill, February 9, 1941.

Definition and Main Properties of Lévy Processes

Definition

An \mathbb{R}^d -valued process X such that $X_0 = 0$ is a **Lévy process** if

- a) for every s, t , $X_{t+s} - X_t$ is independent of \mathcal{F}_t^X
- b) for every s, t the r.v.'s $X_{t+s} - X_t$ and X_s have the same law.
- c) X is continuous in probability, i.e., $\mathbb{P}(|X_t - X_s| > \epsilon) \rightarrow 0$ when $s \rightarrow t$ for every $\epsilon > 0$.

The sum of two independent Lévy processes is a Lévy process.

Property c) implies that a Lévy process has no jumps at fixed time.

A Lévy process admits a càdlàg modification

(A process Y is said to be a modification of X is $\mathbb{P}(X_t = Y_t) = 1, \forall t$)

Let $T > 0$ be fixed. For any $\epsilon > 0$, the set $\{t \in [0, T] : |\Delta X_t| > \epsilon\}$ is finite

The set $\{t \in [0, T] : |\Delta X_t| > 0\}$ is countable

Examples

- **Brownian motion**

The standard Brownian motion is a process W with **continuous paths** such that

- for every s, t , $W_{t+s} - W_t$ is independent of \mathcal{F}_t^W ,
- for every s, t , the r.v. $W_{t+s} - W_t$ has the same law as W_s .

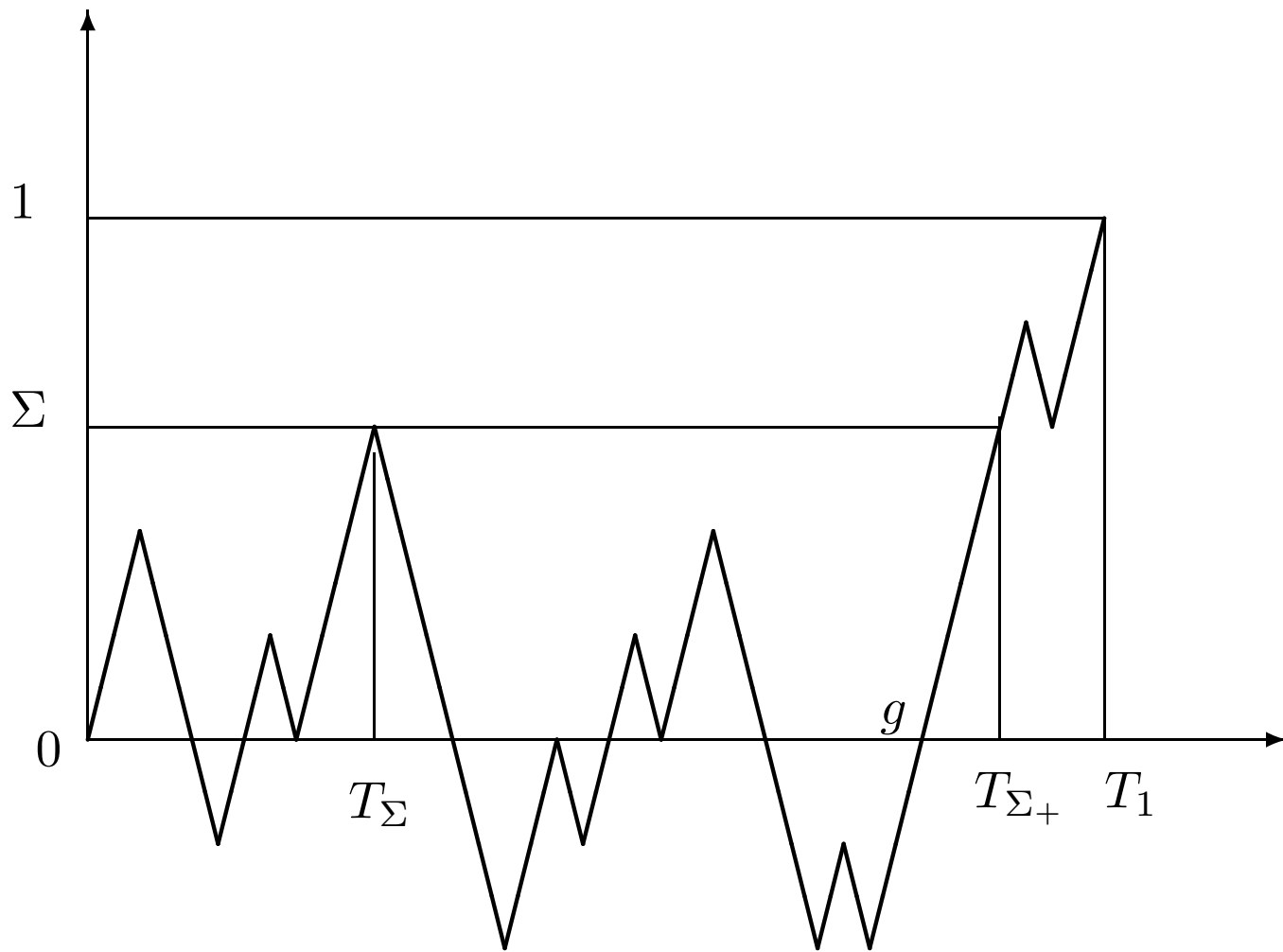
The law of W_s is $\mathcal{N}(0, s)$.

- **Brownian hitting time process**

Let W be a standard BM and for $a > 0$, define

$$T_a := \inf\{t : W_t = a\}$$

The process $(T_a, a \geq 0)$ is a Lévy process.



Non continuity of T_a

• Poisson process

The standard Poisson process is a **counting process** such that

- for every s, t , $N_{t+s} - N_t$ is independent of \mathcal{F}_t^N ,
- for every s, t , the r.v. $N_{t+s} - N_t$ has the same law as N_s .

Then, the r.v. N_t has a Poisson law with parameter λt

$$\mathbb{P}(N_t = k) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

If X has a Poisson law with parameter $\theta > 0$, then

(i) for any $s \in \mathbb{R}$, $\mathbb{E}[s^X] = e^{\theta(s-1)}$.

(ii) $\mathbb{E}[X] = \theta$, $\text{Var}(X) = \theta$.

(iii) for any $u \in \mathbb{R}$, $\mathbb{E}(e^{iuX}) = \exp(\theta(e^{iu} - 1))$

(iv) for any $\alpha \in \mathbb{R}$, $\mathbb{E}(e^{\alpha X}) = \exp(\theta(e^\alpha - 1))$

• Compound Poisson Process

Let λ be a positive number and $F(dy)$ be a probability law on \mathbb{R} (we assume that $\mathbb{P}(Y_1 = 0) = 0$). A (λ, F) -**compound Poisson process** is a process $X = (X_t, t \geq 0)$ of the form

$$X_t = \sum_{k=1}^{N_t} Y_k$$

where N is a Poisson process with intensity $\lambda > 0$ and the $(Y_k, k \in \mathbb{N})$ are i.i.d. random variables, independent of N , with law $F(dy) = \mathbb{P}(Y_1 \in dy)$.

If $\mathbb{E}(|Y_1|) < \infty$, for any t , $\mathbb{E}(X_t) = \lambda t \mathbb{E}(Y_1)$.

The characteristic function of the r.v. X_t is

$$\mathbb{E}[e^{iuX_t}] = e^{\lambda t(\mathbb{E}[e^{iuY_1}] - 1)} = \exp\left(\lambda t \int_{\mathbb{R}} (e^{iuy} - 1)F(dy)\right).$$

Assume that $\mathbb{E}[e^{\alpha Y_1}] < \infty$. Then, the Laplace transform of the r.v. X_t is

$$\mathbb{E}[e^{\alpha X_t}] = e^{\lambda t(\mathbb{E}[e^{\alpha Y_1}] - 1)} = \exp\left(\lambda t \int_{\mathbb{R}} (e^{\alpha y} - 1)F(dy)\right).$$

We shall note $\nu(dy) = \lambda F(dy)$ and say that X is a ν -compound Poisson process.

Martingales

Let X be a Lévy process.

- If $\mathbb{E}(|X_t|) < \infty$, the process $X_t - \mathbb{E}(X_t)$ is a martingale.
- For any u , the process $Z_t(u) := \frac{e^{iuX_t}}{\mathbb{E}(e^{iuX_t})}$ is a martingale.
- If $\mathbb{E}(e^{\lambda X_t}) < \infty$, the process $\frac{e^{\lambda X_t}}{\mathbb{E}(e^{\lambda X_t})}$ is a martingale

Examples

- **Brownian motion**

The standard Brownian motion is a martingale, the process $Y = \mathcal{E}(\lambda W)$ defined by $Y_t = e^{\lambda W_t - \frac{1}{2}\lambda^2 t}$ is a martingale.

The Doléans-Dade exponential $Y = \mathcal{E}(\lambda W)$ satisfies

$$dY_t = Y_t \lambda dW_t$$

- **Poisson process**

The process $M_t = N_t - \lambda t$ is a martingale.

For any α , the process

$$\exp(\alpha N_t - \lambda t(e^\alpha - 1)) = \exp(\alpha M_t - \lambda t(e^\alpha - 1 - \alpha)) = \mathcal{E}(\alpha M)_t$$

is a martingale.

For any β , the process $(1 + \beta)^{N_t} e^{-\lambda \beta t}$ is a martingale

If X is a counting process and if, for some λ the process $M_t = N_t - \lambda t$ is a martingale, then X is a Poisson process

If X is a counting process with stationary and independent increments, then X is a Poisson process.

• Compound Poisson Processes

Assume that $\mathbb{E}(|Y_1|) < \infty$. Then, the process

$(Z_t := X_t - t\lambda\mathbb{E}(Y_1), t \geq 0)$ is a martingale and in particular,

$$\mathbb{E}(X_t) = \lambda t \mathbb{E}(Y_1) = \lambda t \int_{-\infty}^{\infty} y F(dy) = t \int_{-\infty}^{\infty} y \nu(dy)$$

For any $\alpha \in \mathbb{R}$ such that $\int_{-\infty}^{\infty} |e^{\alpha x} - 1| F(dx) < \infty$, the process

$$\exp\left(\alpha X_t - t\lambda \int_{-\infty}^{\infty} (e^{\alpha x} - 1) F(dx)\right) = \exp\left(\alpha X_t - t \int_{-\infty}^{\infty} (e^{\alpha x} - 1) \nu(dx)\right)$$

is a martingale.

Random Measures

- **Counting process:** Let (T_n) be a sequence of random times, with

$$0 < T_1 < \cdots < T_n \dots$$

and $N_t = \sum_{n \geq 1} \mathbb{1}_{T_n \leq t}$. Let A be a Borel set in \mathbb{R}^+ and

$$\mathbf{N}(\omega; A) := \text{Card} \{n \geq 1 : T_n(\omega) \in A\}$$

The measure \mathbf{N} is a random measure and $N_t(\omega) = \mathbf{N}(\omega,]0, t])$.

For a Poisson process

$$\mathbb{E}(\mathbf{N}(A)) = \lambda \text{Leb}(A)$$

Let ν be a radon measure on E . A **random Poisson measure** \mathbf{N} on E with intensity ν is a measure such that

- $\mathbf{N}(A)$ is an integer valued random measure,
- $\mathbf{N}(A) < \infty$ for A bounded Borel set,
- for disjoint sets A_i , the r.v.'s $\mathbf{N}(A_i)$ are independent
- the r.v. $\mathbf{N}(A)$ is Poisson distributed with parameter $\nu(A)$

• **Compound Poisson process:** Define $\mathbf{N} = \sum_n \delta_{T_n, Y_n}$ on $\mathbb{R}^+ \times \mathbb{R}$,
i.e

$$\mathbf{N}(\omega, [0, t] \times A) = \sum_{n=1}^{N_t(\omega)} \mathbb{1}_{Y_n(\omega) \in A}.$$

We shall also write $\mathbf{N}_t(dx) = \mathbf{N}([0, t], dx)$. The measure \mathbf{N} is a random Poisson measure on $\mathbb{R}^+ \times \mathbb{R}$ with intensity $\lambda dt F(dx)$

We denote by $(f * \mathbf{N})_t$ the integral

$$\int_0^t \int_{\mathbb{R}} f(x) \mathbf{N}(ds, dx) = \int_{\mathbb{R}} f(x) \mathbf{N}_t(dx) = \sum_{k=1}^{N_t} f(Y_k) = \sum_{s \leq t} f(\Delta X_s) \mathbb{1}_{\Delta X_s \neq 0}.$$

In particular

$$X_t = \sum_{k=1}^{N_t} Y_k = \sum_{s \leq t} \Delta X_s = \int_0^t \int_{\mathbb{R}} x \mathbf{N}(ds, dx)$$

If $\nu(|f|) < \infty$, the process

$$\begin{aligned} M_t^f &: = (f * \mathbf{N})_t - t\nu(f) = \int_0^t \int_{\mathbb{R}} f(x)(\mathbf{N}(ds, dx) - ds\nu(dx)) \\ &= \sum_{s \leq t} f(\Delta X_s) \mathbb{1}_{\Delta X_s \neq 0} - t\nu(f) \end{aligned}$$

is a martingale.

PROOF: Indeed, the process $Z_t = \sum_{k=1}^{N_t} f(Y_k)$ is a $\hat{\nu}$ compound Poisson process, where $\hat{\nu}$, defined as

$$\hat{\nu}(A) = \lambda \mathbb{P}(f(Y_n) \in A)$$

is the image of ν by f . Hence, if $\mathbb{E}(|f(Y_1)|) < \infty$, the process $Z_t - t\lambda\mathbb{E}(f(Y_1)) = Z_t - t \int f(x)\nu(dx)$ is a martingale.

Using again that Z is a compound Poisson process, it follows that the process

$$\begin{aligned} & \exp \left(\sum_{k=1}^{N_t} f(Y_k) - t \int_{-\infty}^{\infty} (e^{f(x)} - 1) \nu(dx) \right) \\ &= \exp \left(\int_0^t \int_{\mathbb{R}} f(x) \mathbf{N}(ds, dx) - t \int_{-\infty}^{\infty} (e^{f(x)} - 1) \nu(dx) \right) \end{aligned}$$

is a martingale

If X is a pure jump process, if there exists λ and a probability measure σ such that $\sum_{s \leq t} f(\Delta X_s) \mathbb{1}_{\Delta X_s \neq 0} - t\lambda\sigma(f)$ is a martingale, then X is a compound Poisson process.

• Lévy Processes

The random variable $\mathbf{N}([s, t] \times A) = \sum_{s < u \leq t} \mathbb{1}_A(\Delta X_u)$ represents the number of jumps in the time interval $]s, t]$ with jump size in A .

We define ν by

$$\nu(A) = \mathbb{E}(\mathbf{N}([0, 1] \times A))$$

For A compact set such that $0 \notin A$, $\nu(A) < \infty$

The process

$$\mathbf{N}_t^A = \sum_{0 < s \leq t} \mathbb{1}_A(\Delta X_s) = \mathbf{N}([0, t] \times A)$$

is a Poisson process with intensity $\nu(A)$.

The processes \mathbf{N}^A and \mathbf{N}^C are independent if $\nu(A \cap C) = 0$, in particular if A and C are disjoint.

Let A be a Borel set of \mathbb{R}^d with $0 \notin \bar{A}$, and f a Borel function defined on A . We have

$$\int_A f(x) \mathbf{N}_t(\omega, dx) = \int_0^t \int_A f(x) \mathbf{N}(\omega, ds, dx) = \sum_{0 < s \leq t} f(\Delta X_s(\omega)) \mathbb{1}_A(\Delta X_s(\omega)).$$

The process

$$\int_A f(x) \mathbf{N}_t(\omega, dx)$$

is a Lévy process; if $\int_A |f(x)| \nu(dx) < \infty$, then

$$M_t^f = \int_A f(x) \mathbf{N}_t(\omega, dx) - t \int_A f(x) \nu(dx) = \int_0^t \int_A f(x) (\mathbf{N}(ds, dx) - \nu(dx) ds)$$

is a martingale.

If f is bounded and vanishes in a neighborhood of 0,

$$\mathbb{E} \left(\sum_{0 < s \leq t} f(\Delta X_s) \right) = t \int_{\mathbb{R}^d} f(x) \nu(dx).$$

The measure ν satisfies

$$\int (1 \wedge |x|^2) \nu(dx) < \infty$$

i.e. $\int_{|x| \geq 1} \nu(dx) < \infty$ and $\int_{|x| < 1} |x|^2 \nu(dx) < \infty$

Infinitely Divisible Random Variables

Definition

A random variable X taking values in \mathbb{R}^d is **infinitely divisible** if its characteristic function satisfies

$$\hat{\mu}(u) = \mathbb{E}(e^{i(u \cdot X)}) = (\hat{\mu}_n(u))^n$$

where $\hat{\mu}_n$ is a characteristic function.

Examples: The **Gaussian law** $\mathcal{N}(m, \sigma^2)$ has the characteristic function $\exp(ium - u^2\sigma^2/2)$.

Cauchy laws. The standard Cauchy law has the characteristic function $\exp(-c|u|)$.

The **hitting time of the level a for a Brownian motion** has Laplace transform

$$\mathbb{E}[\exp(-\frac{\lambda^2}{2}T_a)] = \exp(-|a|\lambda)$$

Poisson laws. The Poisson law with parameter λ has characteristic function

$$\exp(c(e^{iu} - 1))$$

Poisson Random Sum. Let X_i i.i.d. r.v.'s with characteristic function φ and N a r.v. independent of the X_i 's with a Poisson law. Let

$$X = X_1 + X_2 + \cdots + X_N$$

The characteristic function of X is

$$\exp(-\lambda(1 - \varphi(u)))$$

Gamma laws. The Gamma law $\Gamma(a, \nu)$ has density

$$\frac{\nu^a}{\Gamma(a)} x^{a-1} e^{-\nu x} \mathbb{1}_{x>0}$$

and characteristic function

$$(1 - iu/\nu)^{-a}$$

A **Lévy measure** ν is a positive measure on $\mathbb{R}^d \setminus \{0\}$ such that

$$\int_{\mathbb{R}^d \setminus \{0\}} \min(1, \|x\|^2) \nu(dx) < \infty.$$

Lévy-Khintchine representation.

If X is an infinitely divisible random variable, there exists a triple (m, A, ν) where $m \in \mathbb{R}^d$, A is a non-negative quadratic form and ν is a Lévy measure such that

$$\hat{\mu}(u) = \exp \left(i(u \cdot m) - \frac{1}{2}(u \cdot Au) + \int_{\mathbb{R}^d} (e^{i(u \cdot x)} - 1 - i(u \cdot x) \mathbb{1}_{\{|x| \leq 1\}}) \nu(dx) \right).$$

The triple (m, A, ν) is called the *characteristic triple*.

If $\int |x| \mathbb{1}_{\{|x| \leq 1\}} \nu(dx) < \infty$, one writes the LK representation in a *reduced form*

$$\hat{\mu}(u) = \exp \left(i(u \cdot m_0) - \frac{1}{2}(u \cdot Au) + \int_{\mathbb{R}^d} (e^{i(u \cdot x)} - 1) \nu(dx) \right).$$

A first step is to prove that any function φ such that

$$\varphi(u) = \exp \left(i(u \cdot m) - \frac{1}{2}(u \cdot Au) + \int_{\mathbb{R}^d} (e^{i(u \cdot x)} - 1 - i(u \cdot x) \mathbb{1}_{\{|x| \leq 1\}}) \nu(dx) \right) \quad (*)$$

is a characteristic function (hence, i.d.).

If φ satisfies (*), one proves that

- it is continuous at 0
- it is the limit of characteristic functions.

The result follows.

Continuity: show that

$$\psi(u) = \int (e^{iux} - 1 - iux \mathbb{1}_{\{|x| \leq 1\}}) \nu(dx)$$

is continuous.

$$\psi(u) = \int_{|x| \leq 1} (e^{iux} - 1 - iux) \nu(dx) + \int_{|x| > 1} (e^{iux} - 1) \nu(dx)$$

Then, using the fact that

$$|e^{iux} - 1 - iux| \leq \frac{1}{2} u^2 x^2$$

the result is obtained

Limit:

$$\int_{|x| \leq 1} (e^{iux} - 1 - iux)\nu(dx) = \lim \int_{|x| \geq 1/n} (e^{iux} - 1 - iux)\nu(dx)$$

The right-hand side corresponds to compound Poisson process

If $\hat{\mu}$ is i.d., then it satisfies LK.

If $\hat{\mu}$ is i.d., then $\hat{\mu}(u)$ does not vanish.

Then,

$$\hat{\mu}(u) = (\hat{\mu}_n(u))^n$$

implies that

$$\Phi_n(u) := \exp(n(\hat{\mu}_n(u) - 1)) = \exp(n(e^{\frac{1}{n} \ln \hat{\mu}(u)} - 1))$$

converges to $\hat{\mu}(u)$.

$$\Phi_n(u) = \exp\left(n \int (e^{iux} - 1) \mu_n(dx)\right)$$

is associated with a compound Poisson process.

Examples:

Gaussian laws. The Gaussian law $\mathcal{N}(m, \sigma^2)$ has the characteristic function $\exp(ium - u^2\sigma^2/2)$. Its characteristic triple is $(m, \sigma, 0)$.

Cauchy laws. The standard Cauchy law has the characteristic function

$$\exp(-c|u|) = \exp\left(\frac{c}{\pi} \int_{-\infty}^{\infty} (e^{iux} - 1)x^{-2} dx\right).$$

Its reduced form characteristic triple is $(0, 0, \pi^{-1}x^{-2}dx)$.

Gamma laws. The Gamma law $\Gamma(a, \nu)$ has characteristic function

$$(1 - iu/\nu)^{-a} = \exp \left(a \int_0^{\infty} (e^{iux} - 1) e^{-\nu x} \frac{dx}{x} \right).$$

Its reduced form characteristic triple is $(0, 0, \mathbb{1}_{\{x>0\}} ax^{-1} e^{-\nu x} dx)$.

Brownian hitting times. The first hitting time of $a > 0$ for a Brownian motion has characteristic triple (in reduced form)

$$(0, 0, \frac{a}{\sqrt{2\pi}} x^{-3/2} \mathbb{1}_{\{x>0\}} dx).$$

Indeed $\mathbb{E}(e^{-\lambda T_a}) = e^{-a\sqrt{2\lambda}}$. Moreover

$$\sqrt{2\lambda} = \frac{1}{\sqrt{2}\Gamma(1/2)} \int_0^\infty (1 - e^{-\lambda x}) x^{-3/2} dx,$$

hence, using that $\Gamma(1/2) = \sqrt{\pi}$

$$\mathbb{E}(e^{-\lambda T_a}) = \exp\left(-\frac{a}{\sqrt{2\pi}} \int_0^\infty (1 - e^{-\lambda x}) x^{-3/2} dx\right).$$

Inverse Gaussian laws. The Inverse Gaussian law has density

$$\frac{a}{\sqrt{2\pi}} e^{a\nu} x^{-3/2} \exp\left(-\frac{1}{2}(a^2 x^{-1} + \nu^2 x)\right) \mathbb{1}_{\{x>0\}}$$

This is the law of the first hitting time of a for a Brownian motion with drift ν . The Inverse Gaussian law has characteristic triple (in reduced form)

$$\left(0, 0, \frac{a}{\sqrt{2\pi x^3}} \exp\left(-\frac{1}{2}\nu^2 x\right) \mathbb{1}_{\{x>0\}} dx\right).$$

Inverse Gaussian laws. The Inverse Gaussian law has density

$$\frac{a}{\sqrt{2\pi}} e^{a\nu} x^{-3/2} \exp\left(-\frac{1}{2}(a^2 x^{-1} + \nu^2 x)\right) \mathbb{1}_{\{x>0\}}$$

This is the law of the first hitting time of a for a BM with drift ν .

The Inverse Gaussian law has characteristic triple (in reduced form)

$$\left(0, 0, \frac{a}{\sqrt{2\pi x^3}} \exp\left(-\frac{1}{2}\nu^2 x\right) \mathbb{1}_{\{x>0\}} dx\right).$$

Indeed

$$\begin{aligned} & \exp\left(-\frac{a}{\sqrt{2\pi}} \int_0^\infty \frac{dx}{x^{3/2}} (1 - e^{-\lambda x}) e^{-\nu^2 x/2}\right) \\ &= \exp\left(-\frac{a}{\sqrt{2\pi}} \int_0^\infty \frac{dx}{x^{3/2}} \left((e^{-\nu^2 x/2} - 1) + (1 - e^{-(\lambda + \nu^2/2)x})\right)\right) \\ &= \exp(-a(-\nu + \sqrt{\nu^2 + 2\lambda})) \end{aligned}$$

is the Laplace transform of the first hitting time of a for a BM with drift ν .

Stable Random Variables

A random variable is stable if for any $a > 0$, there exist $b > 0$ and $c \in \mathbb{R}$ such that $[\hat{\mu}(u)]^a = \hat{\mu}(bu) e^{icu}$.

X is stable if

$$\forall n, \exists(\beta_n, \gamma_n), \text{ such that } X_1^{(n)} + \dots + X_n^{(n)} \stackrel{\text{law}}{=} \beta_n X + \gamma_n$$

where $(X_i^{(n)}, i \leq n)$ are i.i.d. random variables with the same law as X .

A stable law is infinitely divisible.

The characteristic function of a stable law can be written

$$\hat{\mu}(u) = \begin{cases} \exp(i\beta u - \frac{1}{2}\sigma^2 u^2), & \text{for } \alpha = 2 \\ \exp(-\gamma|u|^\alpha [1 - i\beta \operatorname{sgn}(u) \tan(\pi\alpha/2)]), & \text{for } \alpha \neq 1, \neq 2 \\ \exp(\gamma|u|(1 - i\beta \operatorname{sgn}(u) \ln|u|)), & \alpha = 1 \end{cases},$$

where $\beta \in [-1, 1]$. For $\alpha \neq 2$, the Lévy measure of a stable law is absolutely continuous with respect to the Lebesgue measure, with density

$$\nu(dx) = \begin{cases} c^+ x^{-\alpha-1} dx & \text{if } x > 0 \\ c^- |x|^{-\alpha-1} dx & \text{if } x < 0. \end{cases}$$

Examples: A Gaussian variable is stable with $\alpha = 2$. The Cauchy law is stable with $\alpha = 1$.

Lévy-Khintchine Representation

Let X be a Lévy process. Then, X_1 is i.i.d.

There exists $m \in \mathbb{R}^d$, a non-negative semi-definite quadratic form A , a Lévy measure ν such that for $u \in \mathbb{R}^d$

$$\mathbb{E}(\exp(i(u \cdot X_1))) = \exp \left(i(u \cdot m) - \frac{1}{2}(u \cdot Au) + \int_{\mathbb{R}^d} (e^{i(u \cdot x)} - 1 - i(u \cdot x) \mathbb{1}_{|x| \leq 1}) \nu(dx) \right)$$

where ν is the Lévy measure.

- If $\nu(\mathbb{R} \setminus \{0\}) < \infty$, the process X has a finite number of jumps in any finite time interval. In finance, one refers to **finite activity**.

- If $\nu(\mathbb{R} \setminus \{0\}) < \infty$, the process X has a finite number of jumps in any finite time interval. In finance, one refers to **finite activity**.
- If $\nu(\mathbb{R} \setminus \{0\}) = \infty$, the process corresponds to **infinite activity**.

The complex valued continuous function Φ such that

$$\mathbb{E} [\exp(iuX_1)] = \exp(-\Phi(u))$$

is called the **characteristic exponent** (sometimes the Lévy exponent) of the Lévy process X .

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is called the **characteristic exponent** (sometimes the Lévy exponent) of the Lévy process X .

If $\mathbb{E} [e^{\lambda X_1}] < \infty$ for any $\lambda > 0$, the function Ψ defined on $[0, \infty[$, such that

$$\mathbb{E} [\exp(\lambda X_1)] = \exp(\Psi(\lambda))$$

is called the **Laplace exponent** of the Lévy process X .

The complex valued continuous function Φ such that

$$\mathbb{E} [\exp(iuX_1)] = \exp(-\Phi(u))$$

is called the **characteristic exponent** (sometimes the Lévy exponent) of the Lévy process X .

If $\mathbb{E} [e^{\lambda X_1}] < \infty$ for any $\lambda > 0$, the function Ψ defined on $[0, \infty[$, such that

$$\mathbb{E} [\exp(\lambda X_1)] = \exp(\Psi(\lambda))$$

is called the **Laplace exponent** of the Lévy process X .

It follows that, if $\Psi(\lambda)$ exists,

$$\mathbb{E} [\exp(iuX_t)] = \exp(-t\Phi(u)), \quad \mathbb{E} [\exp(\lambda X_t)] = \exp(t\Psi(\lambda))$$

and

$$\Psi(\lambda) = -\Phi(-i\lambda).$$

From LK formula, the characteristic exponent and the Laplace exponent can be computed as follows:

$$\begin{aligned}\Phi(u) &= -ium + \frac{1}{2}\sigma^2 u^2 - \int (e^{iux} - 1 - iux \mathbb{1}_{|x|\leq 1})\nu(dx) \\ \Psi(\lambda) &= \lambda m + \frac{1}{2}\sigma^2 \lambda^2 + \int (e^{\lambda x} - 1 - \lambda x \mathbb{1}_{|x|\leq 1})\nu(dx).\end{aligned}$$

Martingales

- If $\mathbb{E}(|X_t|) < \infty$, i.e., $\int_{|x| \geq 1} |x| \nu(dx) < \infty$;

$\mathbb{E}(X_t) = t(m + \int_{|x| \geq 1} |x| \nu(dx))$, the process $X_t - \mathbb{E}(X_t)$ is a martingale.

Martingales

- If $\mathbb{E}(|X_t|) < \infty$, i.e., $\int_{|x| \geq 1} |x| \nu(dx) < \infty$, the process $X_t - \mathbb{E}(X_t)$ is a martingale and $\mathbb{E}(X_t) = t(m + \int_{|x| \geq 1} |x| \nu(dx))$.
- If $\Psi(\alpha)$ exists (i.e., if $\int_{|x| > 1} e^x \nu(dx) < \infty$), the process

$$\frac{e^{\alpha X_t}}{\mathbb{E}(e^{\alpha X_t})} = e^{\alpha X_t - t\Psi(\alpha)}$$

is a martingale

More generally, for any bounded predictable process H

$$\mathbb{E} \left[\sum_{s \leq t} H_s f(\Delta X_s) \right] = \mathbb{E} \left[\int_0^t ds H_s \int f(x) d\nu(x) \right]$$

and if H is a predictable function (i.e. $H : \Omega \times \mathbb{R}^+ \times \mathbb{R}^d \rightarrow \mathbb{R}$ is $\mathcal{P} \times \mathcal{B}$ measurable)

$$\mathbb{E} \left[\sum_{s \leq t} H_s(\omega, \Delta X_s) \right] = \mathbb{E} \left[\int_0^t ds \int d\nu(x) H_s(\omega, x) \right].$$

Both sides are well defined and finite if

$$\mathbb{E} \left[\int_0^t ds \int d\nu(x) |H_s(\omega, x)| \right] < \infty$$

(Exponential formula.) Let X be a Lévy process and ν its Lévy measure. For all t and all Borel function f defined on $\mathbb{R}^+ \times \mathbb{R}^d$ such that $\int_0^t ds \int |1 - e^{f(s,x)}| \nu(dx) < \infty$, one has

$$\mathbb{E} \left[\exp \left(\sum_{s \leq t} f(s, \Delta X_s) \mathbb{1}_{\{\Delta X_s \neq 0\}} \right) \right] = \exp \left(- \int_0^t ds \int (1 - e^{f(s,x)}) \nu(dx) \right).$$

The above property does not extend to predictable functions.

Lévy-Itô's decomposition

If X is a R^d -valued Lévy process, it can be decomposed into $X = Y^{(0)} + Y^{(1)} + Y^{(2)} + Y^{(3)}$ where $Y^{(0)}$ is a affine function, $Y^{(1)}$ is a linear transform of a Brownian motion, $Y^{(2)}$ is a compound Poisson process with jump size greater than or equal to 1 and $Y^{(3)}$ is a Lévy process with jumps sizes smaller than 1. The processes $Y^{(i)}$ are independent.

More precisely

$$X_t = mt + \sigma W_t + X_t^1 + \lim_{\epsilon \rightarrow 0} \tilde{X}_t^\epsilon$$

where

$$X_t^1 = \int_0^t \int_{\{|x| \geq 1\}} x \mathbf{N}(dx, ds) = \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq 1}$$

$$\tilde{X}_t^\epsilon = \int_0^t \int_{\{\epsilon \leq |x| < 1\}} x (\mathbf{N}(dx, ds) - \nu(dx) ds)$$

The processes X^1 is a compound Poisson process, the process \tilde{X}^ϵ is a compensated compound Poisson process, it is a martingale. Note that $\int_0^t \int_{\{\epsilon \leq |x| < 1\}} x \mathbf{N}(dx, ds)$ and $\int_0^t \int_{\{\epsilon \leq |x| < 1\}} x \nu(dx) ds$ are well defined. However, these quantities do not converge as ϵ goes to 0.

Path properties

- The Lévy process X is continuous iff $\nu = 0$
- The Lévy process X with piecewise constant paths iff it is a compound Poisson process or iff $m = 0$, $\sigma = 0$ and $\int \nu(dx) < \infty$
- The Lévy process X is with finite variation path iff $\sigma = 0$ and $\int_{|x| \leq 1} |x| \nu(dx) < \infty$. In that case,

$$ium + \int (e^{iux} - 1 - iux \mathbb{1}_{|x| \leq 1}) \nu(dx)$$

can be written

$$ium_0 + \int (e^{iux} - 1) \nu(dx)$$

and

$$X_t = m_0 t + \sum_{s \leq t} \Delta X_s$$

- If $\int_{|x| \leq 1} |x| \nu(dx) = \infty$, the sum $\sum_{s \leq t} |\Delta X_s| \mathbf{1}_{|\Delta X_s| \leq \epsilon}$ diverges, however the compensated sum converges.

If X is a Lévy process with jumps bounded (by 1), it admits moments of any order, and, setting $Z_t = X_t - \mathbb{E}(X_t)$, $Z = Z^c + Z^d$ where Z^c is a continuous martingale,

$$Z_t^d = \int_{|x|<1} x (\mathbf{N}(dt, dx) - \nu(dx)dt)$$

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If X is a Lévy process, it admits a decomposition as

$$dX_t = \alpha dt + \sigma dB_t + \int_{|x|<1} x (\mathbf{N}(dt, dx) - \nu(dx)dt) + \int_{|x|\geq 1} x \mathbf{N}(dt, dx)$$

The Lévy process is a semi-martingale, hence $\sum_{0 < s \leq t} (\Delta X_s)^2 < \infty$

Some definitions on general stochastic processes

Local martingale

An adapted, right-continuous process M is an **F-local martingale** if there exists a sequence of stopping times (T_n) such that

- (i) The sequence T_n is increasing and $\lim_n T_n = \infty$, a.s.
- (ii) For every n , the stopped process $M^{T_n} \mathbb{1}_{\{T_n > 0\}}$ is an **F**-martingale.

Covariation of Martingales

- **Continuous local martingales:** Let X be a continuous local martingale. The **predictable quadratic variation** process of X is the continuous increasing process $\langle X \rangle$ such that $X^2 - \langle X \rangle$ is a local martingale.

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Let X and Y be two continuous local martingales.

- The predictable covariation process is the continuous finite variation process $\langle X, Y \rangle$ such that $XY - \langle X, Y \rangle$ is a local martingale. Note that $\langle X \rangle = \langle X, X \rangle$ and

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- Integration by parts formula

$$X_t Y_t = X_0 Y_0 + \int_0^t X_s dY_s + \int_0^t Y_s dX_s + \langle X, Y \rangle_t$$

- **General Local martingales:**

Let X and Y be two local martingales.

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► This covariation process is the limit in probability of

$\sum_{i=1}^{p(n)} (X_{t_{i+1}} - X_{t_i})(Y_{t_{i+1}} - Y_{t_i})$, for $0 < t_1 < \dots < t_{p(n)} \leq t$ when $\sup_{i \leq p(n)} (t_i - t_{i-1})$ goes to 0.

► The covariation $[X, Y]$ of both processes X and Y can be also defined by polarisation

$$[X + Y, X + Y] = [X, X] + [Y, Y] + 2[X, Y]$$

Let X and Y be two local martingales.

► The **predictable covariation process** is the finite variation process $\langle X, Y \rangle$ such that

$XY - \langle X, Y \rangle$ is a local martingale

$\langle X, Y \rangle$ is predictable.

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The existence of the predictable covariation process requests some additional conditions on the local martingales ($[X, L]$ is \mathbb{P} -locally integrable).

If W is a Brownian motion $\langle W \rangle_t = [W]_t = t$.

If M is the compensated martingale of a Poisson process, $[M]_t = N_t$ and $\langle M \rangle_t = \lambda t$, and $[W, M] = 0$.

If \mathbb{P} and \mathbb{Q} are equivalent, the covariation process under \mathbb{P} and under \mathbb{Q} are equal. This is not the case for the predictable covariation process.

Spaces of martingales

\mathbf{H}^2 is the set of **square integrable** martingales, i.e., martingales such that $\sup_{t < \infty} E(M_t^2) < \infty$

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For any martingale $M \in \mathbf{H}^2$, we denote by M^c its projection on $\mathbf{H}^{2,c}$ and by M^d its projection on $\mathbf{H}^{2,d}$. Then, $M = M^c + M^d$ is the decomposition of any martingale in \mathbf{H}^2 into its continuous and purely discontinuous parts.

Covariation of Semi-martingales

A semi-martingale is a càdlàg process X such that $X_t = M_t + A_t$, M martingale, A bounded variation process. One can write

$X_t = M_t^c + M_t^d + A_t$ where M^c is continuous. We use the (usual) notation $X^c := M^c$.

► If X and Y are semi-martingales and if X^c, Y^c are their continuous martingale parts, their quadratic covariation is

$$[X, Y]_t = \langle X^c, Y^c \rangle_t + \sum_{s \leq t} (\Delta X_s)(\Delta Y_s).$$

► In the case where X, Y are continuous semi-martingales, their predictable covariation process is the predictable covariation process of their continuous martingale parts.

Stieltjes Integral

Let U be a càdlàg process with bounded variation (i.e., the difference between two increasing processes). The **Stieltjes integral** $\int_0^\infty \theta_s dU_s$ is defined for elementary processes θ of the form $\theta_s = \vartheta_a \mathbb{1}_{]a,b]}(s)$, with ϑ_a a r.v. as $\int_0^\infty \theta_s dU_s = \vartheta_a (U(b) - U(a))$ and for θ such that $\int_0^\infty |\theta_s| |dU(s)| < \infty$ by linearity and passage to the limit. (Hence, the integral is defined path-by-path.) Then, one defines the integral

$$\int_0^t \theta_s dU_s = \int_{]0,t]} \theta_s dU_s = \int_0^\infty \mathbb{1}_{\{]0,t]\}} \theta_s dU_s .$$

Note that if U has a jump at time t_0 , then $(\Theta_t := \int_0^t \theta_s dU_s, t \geq 0)$ has also a jump at time t_0 given as $\Delta \Theta_{t_0} = \Theta_{t_0} - \Theta_{t_0^-} = \theta_{t_0} \Delta U_{t_0}$.

Integration by Parts If U and V are two finite variation processes, Stieltjes' integration by parts formula can be written as follows

$$\begin{aligned} U_t V_t &= U_0 V_0 + \int_{]0,t]} V_s dU_s + \int_{]0,t]} U_{s-} dV_s \\ &= U_0 V_0 + \int_{]0,t]} V_{s-} dU_s + \int_{]0,t]} U_{s-} dV_s + \sum_{s \leq t} \Delta U_s \Delta V_s . \end{aligned}$$

The summation $\sum_{s \leq t} \Delta U_s \Delta V_s$ is in fact a summation over a denumerable number of times s , i.e., the times where U and V admit a common jump. As a partial check, one can verify that the jumps of the left-hand side, i.e., $U_t V_t - U_{t-} V_{t-}$, are equal to the jumps of the right hand side $V_{t-} \Delta U_t + U_{t-} \Delta V_t + \Delta U_t \Delta V_t$.

Chain Rule Let $F \in C^1$ and A a finite variation process. Then,

$$F(A_t) = F(A_0) + \int_0^t F'(A_{s-})dA_s + \sum_{s \leq t} (F(A_s) - F(A_{s-}) - F'(A_{s-})\Delta A_s)$$

or,

$$F(A_t) = F(A_0) + \int_0^t F'(A_{s-})dA_s^c + \sum_{s \leq t} F(A_s) - F(A_{s-})$$

where A^c is the continuous part of A .

Stochastic Integral

Let N be a counting process. The **stochastic integral**

$$\int_0^t C_s dN_s$$

is defined pathwise as a Stieltjes integral for every bounded measurable process (not necessarily \mathbf{F}^N -adapted) $(C_t, t \geq 0)$ by

$$(C \star N)_t \stackrel{def}{=} \int_0^t C_s dN_s = \int_{]0,t]} C_s dN_s \stackrel{def}{=} \sum_{n=1}^{\infty} C_{T_n} \mathbb{1}_{\{T_n \leq t\}}.$$

We emphasize that the integral $\int_0^t C_s dN_s$ is here an integral over the time interval $]0, t]$, where the upper limit t is included and the lower limit 0 excluded. This integral is finite since there is a finite number of jumps during the time interval $]0, t]$.

We shall also write

$$\int_0^t C_s dN_s = \sum_{s \leq t} C_s \Delta N_s$$

where the right-hand side contains only a finite number of non-zero terms. The integral $\int_0^\infty C_s dN_s$ is defined as $\int_0^\infty C_s dN_s = \sum_{n=1}^\infty C_{T_n}$, when the right-hand side converges.

We shall also use the differential notation $d(C \star N)_t \stackrel{def}{=} C_t dN_t$.

Integration by parts formula for Poisson process

Let $(x_t, t \geq 0)$ and $(y_t, t \geq 0)$ be two predictable processes and let $X_t = x + \int_0^t x_s dN_s$ and $Y_t = y + \int_0^t y_s dN_s$. The jumps of X (resp. of Y) occur at the same times as the jumps of N and

$\Delta X_s = x_s \Delta N_s, \Delta Y_s = y_s \Delta N_s$. Then

$$X_t Y_t = xy + \sum_{s \leq t} \Delta(XY)_s = xy + \sum_{s \leq t} X_{s-} \Delta Y_s + \sum_{s \leq t} Y_{s-} \Delta X_s + \sum_{s \leq t} \Delta X_s \Delta Y_s$$

The first equality is obvious, the second one is easy to check.

Integration by parts formula for Poisson process

Let $(x_t, t \geq 0)$ and $(y_t, t \geq 0)$ be two predictable processes and let $X_t = x + \int_0^t x_s dN_s$ and $Y_t = y + \int_0^t y_s dN_s$. The jumps of X (resp. of Y) occur at the same times as the jumps of N and $\Delta X_s = x_s \Delta N_s, \Delta Y_s = y_s \Delta N_s$. Then

$$X_t Y_t = xy + \sum_{s \leq t} \Delta(XY)_s = xy + \sum_{s \leq t} X_{s-} \Delta Y_s + \sum_{s \leq t} Y_{s-} \Delta X_s + \sum_{s \leq t} \Delta X_s \Delta Y_s$$

The first equality is obvious, the second one is easy to check. Hence, from the definition of stochastic integrals

$$X_t Y_t = xy + \int_0^t Y_{s-} dX_s + \int_0^t X_{s-} dY_s + [X, Y]_t$$

where (note that $(\Delta N_t)^2 = \Delta N_t$)

$$[X, Y]_t := \sum_{s \leq t} \Delta X_s \Delta Y_s = \sum_{s \leq t} x_s y_s \Delta N_s = \int_0^t x_s y_s dN_s.$$

More generally, if $dX_t = \mu_t dt + x_t dN_t$ with $X_0 = x$ and $dY_t = \nu_t dt + y_t dN_t$ with $Y_0 = y$, one gets

$$X_t Y_t = xy + \int_0^t Y_{s-} dX_s + \int_0^t X_{s-} dY_s + [X, Y]_t$$

where $[X, Y]_t = \int_0^t x_s y_s dN_s$.

If x is a predictable (bounded) process, the integral

$$\int_0^t x_s dM_s$$

is a martingale.

This is no more the case if x is not predictable, even if the integral is well defined. The process $\int_0^t N_s dM_s$ is not a martingale.

In particular, from integration by parts formula, if $dX_t = x_t dM_t$ and $dY_t = y_t dM_t$, the process $X_t Y_t - [X, Y]_t$ is a local martingale.

Doláns-Dade exponential of a finite variation process

Let U be a càdlàg process with finite variation. The unique solution of

$$dY_t = Y_{t-} dU_t, \quad Y_0 = y$$

is the stochastic exponential of U (the Doléans-Dade exponential of U) equal to

$$\begin{aligned} Y_t &= y \exp(U_t^c - U_0^c) \prod_{s \leq t} (1 + \Delta U_s) \\ &= y \exp(U_t - U_0) \prod_{s \leq t} (1 + \Delta U_s) e^{-\Delta U_s}. \end{aligned}$$

PROOF: Applying the integration by parts formula shows that it is a solution to the equation $dY_t = Y_{t-}dU_t$. As for the uniqueness, if $Y^i, i = 1, 2$ are two solutions, then, setting $Z = Y^1 - Y^2$ we get $Z_t = \int_0^t Z_{s-}dU_s$. Let $M_t = \sup_{s \leq t} |Z_s|$, then, if V_t is the variation process of U_t

$$|Z_t| \leq M_t V_t$$

which implies that

$$|Z_t| \leq M_t \int_0^t V_{s-}dV_s = M_t \frac{V_t^2}{2}.$$

Iterating, we obtain $|Z_t| \leq M_t \frac{V_t^n}{n!}$ and the uniqueness follows by letting $n \rightarrow \infty$. △

Itô's formula

Itô's Formula For Poisson processes

Let N be a Poisson process and f a bounded Borel function. The decomposition

$$f(N_t) = f(N_0) + \sum_{0 < s \leq t} [f(N_s) - f(N_{s-})]$$

is trivial and is the main step to obtain Itô's formula for a Poisson process.

We can write the right-hand side as a stochastic integral:

$$\begin{aligned} \sum_{0 < s \leq t} [f(N_s) - f(N_{s-})] &= \sum_{0 < s \leq t} [f(N_{s-} + 1) - f(N_{s-})] \Delta N_s \\ &= \int_0^t [f(N_{s-} + 1) - f(N_{s-})] dN_s, \end{aligned}$$

hence, the canonical decomposition of $f(N_t)$ as the sum of a martingale and an absolute continuous adapted process is

$$f(N_t) = f(N_0) + \int_0^t [f(N_{s-} + 1) - f(N_{s-})] dM_s + \int_0^t [f(N_{s-} + 1) - f(N_{s-})] \lambda ds .$$

More generally, assume that N is an inhomogeneous Poisson process (i.e., N is a counting process and there exists a non-negative function λ such that $N_t - \int_0^t \lambda(s)ds$ is a martingale). Let h be an adapted process and g a predictable process such that $\int_0^t |h_s|ds < \infty$, $\int_0^t |g_s|\lambda_s ds < \infty$. Let $F \in C^{1,1}(\mathbb{R}^+ \times \mathbb{R})$ and

$$dX_t = h_t dt + g_t dM_t = (h_t - g_t \lambda_t) dt + g_t dN_t$$

Then

$$\begin{aligned} F(t, X_t) &= F(0, X_0) + \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) (h_s - g_s \lambda(s)) ds \\ &\quad + \sum_{s \leq t} F(s, X_s) - F(s, X_{s-}) \\ &= F(0, X_0) + \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) dX_s \\ &\quad + \sum_{s \leq t} [F(s, X_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) g_s \Delta N_s] . \end{aligned}$$

$$\begin{aligned} F(t, X_t) &= F(0, X_0) + \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_s) (h_s - g_s \lambda(s)) ds \\ &\quad + \int_0^t [F(s, X_s) - F(s, X_{s-})] dN_s \end{aligned}$$

PROOF: Indeed, between two jumps, $dX_t = (h_t - \lambda_t g_t)dt$, and for $T_n < s < t < T_{n+1}$,

$$F(t, X_t) = F(s, X_s) + \int_s^t \partial_t F(u, X_u) du + \int_s^t \partial_x F(u, X_u) (h_u - g_u \lambda_u) du.$$

At jump times, $F(T_n, X_{T_n}) = F(T_n, X_{T_n-}) + \Delta F(\cdot, X)_{T_n}$. \triangle

Remark that, in the “ ds ” integrals, we can write X_{s-} or X_s , since, for any bounded Borel function f ,

$$\int_0^t f(X_{s-})ds = \int_0^t f(X_s)ds.$$

Note that since dN_s a.s. $N_s = N_{s-} + 1$, one has

$$\int_0^t f(N_{s-})dN_s = \int_0^t f(N_s - 1)dN_s.$$

We shall use systematically use the form $\int_0^t f(N_{s-})dN_s$, even if the $\int_0^t f(N_s - 1)dN_s$ has a meaning.

The reason is that $\int_0^t f(N_{s-})dM_s = \int_0^t f(N_{s-})dN_s + \lambda \int_0^t f(N_{s-})ds$ is a martingale, whereas $\int_0^t f(N_s - 1)dM_s$ is not.

Check that the above formula can be written as

$$\begin{aligned}
& F(t, X_t) - F(0, X_0) \\
= & \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_s) (h_s - g_s \lambda(s)) ds \\
& + \int_0^t [F(s, X_s) - F(s, X_{s-})] dN_s \\
= & \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) dX_s \\
& + \int_0^t [F(s, X_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) g_s] dN_s
\end{aligned}$$

$$\begin{aligned}
& F(t, X_t) - F(0, X_0) \\
= & \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) dX_s \\
& + \int_0^t [F(s, X_{s-} + g_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) g_s] dN_s \\
= & \int_0^t (\partial_t F(s, X_s) + [F(s, X_{s-} + g_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) g_s] \lambda) ds \\
& + \int_0^t [F(s, X_{s-} + g_s) - F(s, X_{s-})] dM_s
\end{aligned}$$

Let X be a ν -compound Poisson process, and $Z_t = Z_0 + bt + X_t$. Then,

using that $\mathbf{N} = \sum_{n=1}^{\infty} \delta_{T_n, Y_n}$, Itô's formula

$$\begin{aligned}
 f(Z_t) - f(Z_0) &= b \int_0^t f'(Z_s) ds + \sum_{k, T_k \leq t} f(Z_{T_k}) - f(Z_{T_k-}) \\
 &= b \int_0^t f'(Z_s) ds + \int_0^t \int_{\mathbb{R}} [f(Z_{s-} + y) - f(Z_{s-})] \mathbf{N}(ds, dy) \\
 &= \int_0^t ds (\mathcal{L}f)(Z_s) + M(f)_t
 \end{aligned}$$

can be written as where $\mathcal{L}f(x) = bf'(x) + \int_{\mathbb{R}} (f(x+y) - f(x)) \nu(dy)$ is the infinitesimal generator of Z and

$$M(f)_t = \int_0^t \int_{\mathbb{R}} [f(Z_{s-} + y) - f(Z_{s-})] (\mathbf{N}(ds, dy) - ds \nu(dy))$$

is a local martingale.

Let \mathbb{Q} be equivalent to \mathbb{P} on \mathcal{F}_t , for any t and $\mathbb{Q}|_{\mathcal{F}_t} = L_t \mathbb{P}|_{\mathcal{F}_t}$ where L is a strictly positive \mathbb{P} -martingale. Any \mathbb{P} -local martingale X is a \mathbb{Q} semi-martingale and its semi-martingale decompositions are given by the following theorem:

(i)

$$X_t - \int_0^t \frac{d[X, L]_s}{L_s} \text{ is a } \mathbb{Q}\text{-local martingale}$$

(ii) If $[X, L]$ is \mathbb{P} -locally integrable, the process

$$X_t - \int_0^t \frac{d\langle X, L \rangle_s}{L_{s-}} \text{ is a } \mathbb{Q}\text{-local martingale}$$

General case

Let X be a semi-martingale and $f \in C^{1,2}$ Then

$$\begin{aligned} df(t, X_t) &= \partial_t f(t, X_t)dt + \partial_x f(t, X_{t-})dX_t + \frac{1}{2}\partial_{xx} f(t, X_{t-})d[X^c]_t \\ &\quad + f(t, X_t) - f(t, X_{t-}) - \Delta X_t \partial_x f(t, X_{t-}) \end{aligned}$$

Back to Lévy processes

Begin at the beginning, and go on till you come to the end. Then, stop.

L. Carroll, Alice's Adventures in Wonderland

Covariation processes

Let X be a (m, σ^2, ν) real valued Lévy process. Then, $X_t^c = \sigma W_t$ and

$$[X]_t = \sigma^2 t + \int_0^t \int x^2 \mathbf{N}(ds, dx)$$

If $\int x^2 \nu(dx) < \infty$,

$$\langle X \rangle_t = \sigma^2 t + t \int x^2 \nu(dx)$$

Itô's formula

If X is a Lévy process, it admits a decomposition as

$$dX_t = \alpha dt + \sigma dB_t + \int_{|x|<1} x (\mathbf{N}(dt, dx) - \nu(dx)dt) + \int_{|x|\geq 1} x \mathbf{N}(dt, dx)$$

The Lévy process is a semi-martingale, hence $\sum_{0 < s \leq t} (\Delta X_s)^2 < \infty$

$$\begin{aligned}
f(X_t) &= f(X_0) + \frac{\sigma^2}{2} \int_0^t f''(X_s) ds + \int_0^t f'(X_{s-}) dX_s \\
&\quad + \sum_{s \leq t} (f(X_{s-} + \Delta X_s) - f(X_{s-}) - \Delta X_s f'(X_{s-}))
\end{aligned}$$

As a consequence of the semi-martingale property, if F is a C^2 function, then, the series

$$\sum_{s \leq t} f(X_{s-} + \Delta X_s) - f(X_{s-}) - \Delta X_s f'(X_{s-})$$

converges, since

$$|f(X_{s-} + \Delta X_s) - f(X_{s-}) - \Delta X_s f'(X_{s-})| \leq c(\Delta X_s)^2$$

.

Let $Y_t = f(t, X_t)$, with f bounded with bounded derivatives. Then, Y is a semi-martingale

Its martingale part is

$$\partial_x f(t, X_t) \sigma dB_t + \int (f(t, X_{t-} + x) - f(t, X_{t-})) (\mathbf{N}(dt, dx) - \nu(dx)dt)$$

Its finite variation part is

$$\begin{aligned} \partial_t f(t, X_t) &+ \alpha \partial_x f(t, X_t) + \frac{1}{2} \sigma^2 \partial_{xx} f(t, X_t) \\ &+ \int (f(t, X_{t-} + x) - f(t, X_{t-}) - x \partial_x f(t, X_{t-}) \mathbb{1}_{x \leq 1}) \nu(dx) \end{aligned}$$

Representation Theorem

Let X be a \mathbb{R}^d -valued Lévy process and \mathbf{F}^X its natural filtration. Let M be a locally square integrable martingale with $M_0 = m$. Then, there exists a family (φ, ψ) of predictable processes such that

$$\int_0^t |\varphi_s^i|^2 ds < \infty, \text{ a.s.}$$

$$\int_0^t \int_{\mathbb{R}^d} |\psi_s(x)|^2 ds \nu(dx) < \infty, \text{ a.s.}$$

and

$$M_t = m + \sum_{i=1}^d \int_0^t \varphi_s^i dW_s^i + \int_0^t \int_{\mathbb{R}^d} \psi_s(x) (\mathbf{N}(ds, dx) - ds \nu(dx)).$$

Change of measure

Poisson Process

Let N be a Poisson process with intensity λ , and \mathbb{Q} be the probability defined by (with $\beta > -1$)

$$\frac{d\mathbb{Q}}{d\mathbb{P}} \Big|_{\mathcal{F}_t} = (1 + \beta)^{N_t} e^{-\lambda\beta t}$$

Then, the process N is a \mathbb{Q} -Poisson process with intensity equal to $(1 + \beta)\lambda$.

The process L defined as

$$L_t = (1 + \beta)^{N_t} e^{-\lambda\beta t}$$

is a strictly positive martingale with expectation equal to 1. Then, from the definition of \mathbb{Q} , for any sequence $0 = t_1 < t_2 < \dots < t_{n+1} = t$,

$$\mathbb{E}_{\mathbb{Q}} \left(\prod_{i=1}^n x_i^{N_{t_{i+1}} - N_{t_i}} \right) = \mathbb{E}_{\mathbb{P}} \left(e^{-\lambda\beta t} \prod_{i=1}^n ((1 + \beta)x_i)^{N_{t_{i+1}} - N_{t_i}} \right)$$

The right-hand side is computed using that, under \mathbb{P} , the process N is a Poisson process (hence with independent increments) and is equal to

$$\begin{aligned} e^{-\lambda\beta t} \prod_{i=1}^n \mathbb{E}_{\mathbb{P}} \left(((1 + \beta)x_i)^{N_{t_{i+1}} - t_i} \right) &= e^{-\lambda\beta t} \prod_{i=1}^n e^{-\lambda(t_{i+1} - t_i)} e^{\lambda(t_{i+1} - t_i)(1 + \beta)x_i} \\ &= \prod_{i=1}^n e^{(1 + \beta)\lambda(t_{i+1} - t_i)(x_i - 1)}. \end{aligned}$$

$$\mathbb{E}_{\mathbb{Q}} \left(\prod_{i=1}^n x_i^{N_{t_{i+1}} - N_{t_i}} \right) = \prod_{i=1}^n e^{(1+\beta)\lambda(t_{i+1}-t_i)(x_i-1)} .$$

In particular, for any j (take all the x_i 's, except the j th one, equal to 1)

$$\mathbb{E}_{\mathbb{Q}} \left(x_j^{N_{t_{j+1}} - N_{t_j}} \right) = e^{(1+\beta)\lambda(t_{j+1}-t_j)(x_j-1)} ,$$

which establishes that, under \mathbb{Q} , the r.v. $N_{t_{j+1}} - N_{t_j}$ has a Poisson law with parameter $(1 + \beta)\lambda$, then that

$$\mathbb{E}_{\mathbb{Q}} \left(\prod_{i=1}^n x_i^{N_{t_{i+1}} - N_{t_i}} \right) = \prod_{i=1}^n \mathbb{E}_{\mathbb{Q}} \left(x_i^{N_{t_{i+1}} - N_{t_i}} \right)$$

which is equivalent to the independence of the increments.

Compound Poisson process

Let X be a ν -compound Poisson process under \mathbb{P} , we present some particular probability measures \mathbb{Q} equivalent to \mathbb{P} such that, under \mathbb{Q} , X is still a compound Poisson process.

Let $\tilde{\nu}$ a positive finite measure on \mathbb{R} absolutely continuous w.r.t. ν , and $\tilde{\lambda} = \tilde{\nu}(\mathbb{R}) > 0$. Let

$$L_t = \exp \left(t(\lambda - \tilde{\lambda}) + \sum_{s \leq t} \ln \left(\frac{d\tilde{\nu}}{d\nu} \right) (\Delta X_s) \right).$$

Recall that

$$\exp \left(\int_0^t \int_{\mathbb{R}} f(x) \mathbf{N}(ds, dx) - t \int_{-\infty}^{\infty} (e^{f(x)} - 1) \nu(dx) \right)$$

is a martingale

Applying this martingale property for $f = \ln \left(\frac{d\tilde{\nu}}{d\nu} \right)$, the process L is a martingale. Set $\mathbb{Q}|_{\mathcal{F}_t} = L_t \mathbb{P}|_{\mathcal{F}_t}$. **Under \mathbb{Q} , the process X is a $\tilde{\nu}$ -compound Poisson process.**

PROOF: First we find the law of the r.v. X_t under \mathbb{Q} . From the definition of \mathbb{Q}

$$\begin{aligned}
\mathbb{E}_{\mathbb{Q}}(e^{iuX_t}) &= \mathbb{E}_P(e^{iuX_t} \exp \left(t(\lambda - \hat{\lambda}) + \sum_{k=1}^{N_t} f(Y_k) \right)) \\
&= \sum_{n=0}^{\infty} e^{-\lambda t} \frac{(\lambda t)^n}{n!} e^{t(\lambda - \hat{\lambda})} \left(\mathbb{E}_P(e^{iuY_1 + f(Y_1)}) \right)^n \\
&= \sum_{n=0}^{\infty} e^{-\lambda t} \frac{(\lambda t)^n}{n!} e^{t(\lambda - \hat{\lambda})} \left(\mathbb{E}_P \left(\frac{d\hat{\nu}}{d\nu}(Y_1) e^{iuY_1} \right) \right)^n \\
&= \sum_{n=0}^{\infty} \frac{(\lambda t)^n}{n!} e^{-t\hat{\lambda}} \left(\frac{1}{\lambda} \int e^{iuy} d\hat{\nu}(y) \right)^n = \exp t \int (e^{iuy} - 1) d\hat{\nu}(y)
\end{aligned}$$

It remains to check that X is with independent and stationary increments under \mathbb{Q} .

By Bayes formula, for $t > s$

$$\begin{aligned}\mathbb{E}_Q(e^{iu(X_t - X_s)} | \mathcal{F}_s) &= \frac{1}{L_s} \mathbb{E}_P(L_t e^{iu(X_t - X_s)} | \mathcal{F}_s) \\ &= \exp\left((t - s) \int (e^{iux} - 1) \tilde{\nu}(dx)\right).\end{aligned}$$

△

Esscher transform

We assume that $\mathbb{E}(e^{(\theta \cdot X_t)}) < \infty$. We define a probability measure \mathbb{Q} , equivalent to \mathbb{P} by the formula

$$\mathbb{Q}|_{\mathcal{F}_t} = \frac{e^{(\theta \cdot X_t)}}{\mathbb{E}(e^{(\theta \cdot X_t)})} \mathbb{P}|_{\mathcal{F}_t} .(*)$$

This particular choice of measure transformation, (called an Esscher transform) preserves the Lévy process property.

Let X be a \mathbb{P} -Lévy process with parameters (m, A, ν) where $A = R^T R$. Let θ be such that $\mathbb{E}(e^{(\theta \cdot X_t)}) < \infty$ and suppose \mathbb{Q} is defined by (*).

Then X is a Lévy process under \mathbb{Q}

It is not difficult to prove that X has independent and stationary increments under \mathbb{Q} . The characteristic exponent of X under \mathbb{Q} is $\Phi^{(\theta)}$ such that

$$\begin{aligned} e^{-t\Phi^{(\theta)}(u)} &= \mathbb{E}_{\mathbb{Q}}(e^{i(u \cdot X_t)}) = \mathbb{E}(e^{i(u \cdot X_t) + (\theta \cdot X_t)})e^{t\Phi(-i\theta)} \\ &= e^{-t(\Phi(u - i\theta) - \Phi(-i\theta))}. \end{aligned}$$

The characteristic exponent of X under \mathbb{Q} is

$$\Phi^{(\theta)}(u) = \Phi(u - i\theta) - \Phi(-i\theta).$$

If $\Psi(\theta) < \infty$, $\Psi^{(\theta)}(u) = \Psi(u + \theta) - \Psi(\theta)$ for $u \geq \min(-\theta, 0)$.

A simple computation leads to

$$\begin{aligned}
\Phi(u - i\theta) - \Phi(-i\theta) &= -iu \cdot m + \frac{1}{2}u \cdot Au - \frac{1}{2}iu \cdot A\theta - \frac{1}{2}i\theta \cdot Au \\
&\quad - \int (e^{\theta \cdot x} (e^{iu \cdot x} - 1) - iu \cdot x \mathbb{1}_{\{|x| \leq 1\}}) \nu(dx) \\
&= -iu \cdot \left(m + \frac{1}{2}(A + A^T)\theta + \int (e^{\theta \cdot x} - 1)x \mathbb{1}_{\{|x| \leq 1\}} \nu(dx) \right) \\
&\quad + \frac{1}{2}u \cdot Au + \int e^{\theta \cdot x} (e^{iu \cdot x} - 1 - iu \cdot x \mathbb{1}_{\{|x| \leq 1\}}) \nu(dx).
\end{aligned}$$

Hence, X_1 has the required Lévy-Khintchine representation under \mathbb{Q} .

$$\begin{aligned} \mathbb{E}_Q(\exp(i(u \cdot X_1))) &= \exp \left(i(u \cdot m^{(\theta)}) - \frac{1}{2}(u \cdot Au) \right. \\ &\quad \left. + \int_{\mathbb{R}^d} (e^{i(u \cdot x)} - 1 - i(u \cdot x) \mathbb{1}_{|x| \leq 1}) \nu^{(\theta)}(dx) \right) \end{aligned}$$

with

$$\begin{aligned} m^{(\theta)} &= m + \frac{1}{2}(A + A^T)\theta + \int_{|x| \leq 1} x(e^{\theta x} - 1)\nu(dx) \\ \nu^{(\theta)}(dx) &= e^{\theta x} \nu(dx). \end{aligned}$$

General case

More generally, any density $(L_t, t \geq 0)$ which is a positive martingale can be used.

$$dL_t = \sum_{i=1}^d \tilde{\varphi}_t^i dW_t^i + \int \tilde{\psi}_t(x) [\mathbf{N}(dt, dx) - dt\nu(dx)].$$

From the strict positivity of L , there exists φ, ψ such that

$\tilde{\varphi}_t = L_{t-} \varphi_t$, $\tilde{\psi}_t = L_{t-} (e^{\psi(t,x)} - 1)$, hence the process L satisfies

$$dL_t = L_{t-} \left(\sum_{i=1}^d \varphi_t^i dW_t^i + \int (e^{\psi(t,x)} - 1) [\mathbf{N}(dt, dx) - dt\nu(dx)] \right) (**)$$

Let $\mathbb{Q}|\mathcal{F}_t = L_t \mathbb{P}|\mathcal{F}_t$ where L is defined in (**). With respect to \mathbb{Q} ,

(i) $W_t^\varphi \stackrel{\text{def}}{=} W_t - \int_0^t \varphi_s ds$ is a Brownian motion

(ii) The process N is compensated by $e^{\psi(s,x)} ds\nu(dx)$ meaning that for any Borel function h such that

$$\int_0^T \int_{\mathbb{R}} |h(s,x)| e^{\psi(s,x)} ds\nu(dx) < \infty,$$

the process

$$\int_0^t \int_{\mathbb{R}} h(s,x) \left(\mathbf{N}(ds, dx) - e^{\psi(s,x)} ds\nu(dx) \right)$$

is a local martingale.

Fluctuation theory

Let $M_t = \sup_{s \leq t} X_s$ be the running maximum of the Lévy process X . The reflected process $M - X$ enjoys the strong Markov property.

Let θ be an exponential variable with parameter q , independent of X . Note that

$$\mathbb{E}(e^{iuX_\theta}) = q \int \mathbb{E}(e^{iuX_t})e^{-qt} dt = q \int e^{-t\Phi(u)}e^{-qt} dt.$$

Using excursion theory, **the random variables M_θ and $X_\theta - M_\theta$ can be proved to be independent**, hence

$$\mathbb{E}(e^{iuM_\theta})\mathbb{E}(e^{iu(X_\theta - M_\theta)}) = \frac{q}{q + \Phi(u)}.$$

This equality is known as the **Wiener-Hopf factorization**.

Let $m_t = \min_{s \leq t}(X_s)$. Then

$$m_\theta \stackrel{law}{=} X_\theta - M_\theta.$$

If $\mathbb{E}(e^{X_1}) < \infty$, using Wiener-Hopf factorization, Mordecki proves that

the boundaries for perpetual American options are given by

$$b_p = K\mathbb{E}(e^{m_\theta}), b_c = K\mathbb{E}(e^{M_\theta})$$

where $m_t = \inf_{s \leq t} X_s$ and θ is an exponential r.v. independent of X with parameter r , hence $b_c b_p = \frac{rK^2}{1 - \ln \mathbb{E}(e^{X_1})}$.

Pecherskii-Rogozin Identity

For $x > 0$, denote by T_x the first passage time above x defined as

$$T_x = \inf\{t > 0 : X_t > x\}$$

and by $K_x = X_{T_x} - x$ the so-called overshoot.

For every triple of positive numbers (α, β, q) ,

$$\int_0^\infty e^{-qx} \mathbb{E}(e^{-\alpha T_x - \beta K_x}) dx = \frac{\kappa(\alpha, q) - \kappa(\alpha, \beta)}{(q - \beta)\kappa(\alpha, q)}$$

where κ is the Laplace exponent of the ladder process defined as

$$e^{-\ell\kappa(\alpha, \beta)} = E(\exp(-\alpha\tau_\ell - \beta H_\ell)),$$

where H is defined in terms of M and the local time of $M - X$.

Exponential Lévy Processes

Let X be a Lévy process.

The process X is a martingale iff $\int_{|x|\geq 1} |x|\nu(dx) < \infty$ and $b + \int_{|x|\geq 1} x\nu(dx) = 0$.

The process e^X is a martingale iff $\int_{|x|\geq 1} e^x\nu(dx)$ and $b + \frac{1}{2}\sigma^2 + \int(e^x - 1 - x\mathbb{1}_{|x|\leq 1})\nu(dx) = 0$

Let $C(t, S)$ be a $C^{1,2}$ function and $S_t = S_0 e^{rt+X_t}$ where $\int_{|x|\geq 1} e^{2x}\nu(dx) < \infty$.

The process $e^{-rt}C(t, S_t)$ is a martingale iff

$$\begin{aligned} \frac{\partial C}{\partial t} + rS \frac{\partial C}{\partial S} + \frac{\sigma^2}{2} S^2 \frac{\partial^2 C}{\partial S^2} - rC \\ + \int \nu(dx) (C(t, Se^x) - C(t, S) - S(e^x - 1) \frac{\partial C}{\partial S}) = 0 \end{aligned}$$

Exponential and stochastic exponential of Lévy Processes

Doléans-Dade Exponential Let X be a real-valued (m, σ^2, ν) -Lévy process. The solution of

$$dZ_t = Z_{t-}dX_t, Z_0 = 1$$

is

$$Z_t = e^{X_t - \frac{1}{2}\sigma^2 t} \prod_{s \leq t} (1 + \Delta X_s) e^{-\Delta X_s}$$

Proof: in a first step, we prove that

$$V_t = \prod_{s \leq t} (1 + \Delta X_s) e^{-\Delta X_s}$$

is well-defined and is a finite variation process.

$$V_t = \underbrace{\prod_{s \leq t, |\Delta X_s| \leq 1/2} (1 + \Delta X_s) e^{-\Delta X_s}}_{V_t^1} \underbrace{\prod_{s \leq t, |\Delta X_s| > 1/2} (1 + \Delta X_s) e^{-\Delta X_s}}_{V_t^2}$$

The product in V^2 contains a finite number of terms

The process V^1 is non-negative and

$$\ln(V_t^1) = \sum_{s \leq t, |\Delta X_s| \leq 1/2} (\ln(1 + \Delta X_s) - \Delta X_s)$$

Using

$$0 \geq (\ln(1 + \Delta X_s) - \Delta X_s) \geq -(\Delta X_s)^2$$

we check that V^1 is well defined and with bounded variation.

Then, we apply Itô's formula. Let $Z_t = e^{X_t - \frac{1}{2}\sigma^2 t} V_t$. Then

$$\begin{aligned}
 dZ_t &= -\frac{\sigma^2}{2} Z_{t-} dt + Z_{t-} dX_t + e^{X_{t-} - \frac{1}{2}\sigma^2 t} dV_t + \frac{\sigma^2}{2} Z_{t-} dt \\
 &\quad + (Z_t - Z_{t-}) - Z_{t-} \Delta X_t - e^{X_{t-} - \frac{1}{2}\sigma^2 t} \Delta V_t \quad (\dagger) \\
 &= Z_{t-} dX_t + e^{X_{t-} - \frac{1}{2}\sigma^2 t} (V_t e^{\Delta X_t} - V_{t-} - V_{t-} \Delta X_t) \\
 &= Z_{t-} dX_t
 \end{aligned}$$

More generally, the solution of the SDE

$$dS_t = S_{t-}(b(t)dt + \sigma(t)dX_t)$$

is

$$S_t = S_0 \exp \left(\int_0^t \sigma(s)dX_s + \int_0^t \left(b(s) - \frac{\sigma^2(s)}{2} \right) ds \right) \\ \prod_{0 < s \leq t} (1 + \sigma(s)\Delta X_s) \exp(-\sigma(s)\Delta X_s).$$

Exponentials of Lévy Processes

Let X be a real-valued (m, σ^2, ν) -Lévy process.

Let $S_t = e^{X_t}$ be the ordinary exponential of the process X . The stochastic logarithm of S (i.e., the process Y which satisfies $S_t = \mathcal{E}(Y)_t$) is a Lévy process and is given by

$$Y_t := \mathcal{L}(S)_t = X_t + \frac{1}{2}\sigma^2 t - \sum_{0 < s \leq t} (1 + \Delta X_s - e^{\Delta X_s}) .$$

The Lévy characteristics of Y are

$$m_Y = m + \frac{1}{2}\sigma^2 + \int ((e^x - 1)\mathbb{1}_{\{|e^x - 1| \leq 1\}} - x\mathbb{1}_{\{|x| \leq 1\}}) \nu(dx)$$

$$\sigma_Y^2 = \sigma^2$$

$$\nu_Y(A) = \nu(\{x : e^x - 1 \in A\}) = \int \mathbb{1}_A(e^x - 1) \nu(dx) .$$

The process $Y_t = X_t + \frac{1}{2}\sigma^2 t - \sum_{0 < s \leq t} (1 + \Delta X_s - e^{\Delta X_s})$ is a Lévy process, $\sigma_Y^2 = \sigma^2$, and $\Delta Y_t = e^{\Delta X_t} - 1$.

This implies the form of $\nu_Y(dx)$. (†)

we obtain that the Lévy-Itô decomposition of Y is

$$\begin{aligned}
Y_t &= X_t + \frac{1}{2}\sigma^2 t - \sum_{0 < s \leq t} (1 + \Delta X_s - e^{\Delta X_s}) \\
&= mt + \sigma B_t + \int_0^t \int_{\{|x| \leq 1\}} x \tilde{\mathbf{N}}(ds, dx) + \int_0^t \int_{\{|x| > 1\}} x \mathbf{N}(ds, dx) + \frac{1}{2}\sigma^2 t \\
&\quad - \int_0^t \int (1 + x - e^x) \mathbf{N}(ds, dx) \\
&= m_Y t + \sigma B_t + \int_0^t \int (e^x - 1) \mathbb{1}_{\{|e^x - 1| \leq 1\}} \tilde{\mathbf{N}}(ds, dx) \\
&\quad + \int_0^t \int (e^x - 1) \mathbb{1}_{\{|e^x - 1| > 1\}} \mathbf{N}(ds, dx) \\
&= m_Y t + \sigma B_t + \int_0^t \int y \mathbb{1}_{\{|y| \leq 1\}} \tilde{\mathbf{N}}_Y(ds, dy) + \int_0^t \int y \mathbb{1}_{\{|y| > 1\}} \mathbf{N}_Y(ds, dy).
\end{aligned}$$

The result follows.

Let $Z_t = \mathcal{E}(X)_t$ the Doléans-Dade exponential of X . If $Z > 0$, the ordinary logarithm of Z is a Lévy process L given by

$$L_t := \ln(Z_t) = X_t - \frac{1}{2}\sigma^2 t + \sum_{0 < s \leq t} (\ln(1 + \Delta X_s) - \Delta X_s) .$$

Its Lévy characteristics are

$$m_L = m - \frac{1}{2}\sigma^2 + \int (\ln(1 + x) \mathbb{1}_{\{|\ln(1+x)| \leq 1\}} - x \mathbb{1}_{\{|x| \leq 1\}}) \nu(dx)$$

$$\sigma_L^2 = \sigma^2$$

$$\nu_L(A) = \nu(\{x : \ln(1 + x) \in A\}) = \int \mathbb{1}_A(\ln(1 + x)) \nu(dx)$$

Exponentials of Lévy Processes

Let $S_t = xe^{X_t}$ where X is a (m, σ^2, ν) real valued Lévy process.

Let us assume that $\mathbb{E}(e^{-\alpha X_1}) < \infty$, for $\alpha \in [-\epsilon, \epsilon]$. This implies that X has finite moments of all orders.

A particular case

Let W be a Brownian motion and Z be a ν -compound Poisson process independent of W of the form $Z_t = \sum_{n=1}^{N_t} Y_n$. Let

$$dS_t = S_{t-}(\mu dt + \sigma dW_t + dZ_t), \quad (0.1)$$

where μ and σ are constants. The process $(S_t e^{-rt}, t \geq 0)$ is a martingale if and only if $E(|Y_1|) < \infty$ and $\mu + \lambda E(Y_1) = r$.

If $Y_1 \geq -1$ a.s., the process S can be written in an exponential form as

$$S_t = S_0 e^{X_t}, \quad X_t = bt + \sigma W_t + V_t$$

where $b = \mu - \frac{1}{2}\sigma^2$, V is the (λ, \tilde{F}) -compound Poisson process

$$V_t = \sum_{n=1}^{N_t} \ln(1 + Y_n) = \sum_{n=1}^{N_t} U_n$$

with $\tilde{F}(u) = F(e^u - 1)$.

Option pricing with Esscher Transform

Let $S_t = S_0 e^{rt+X_t}$ where X is a Lévy process under the historical probability \mathbb{P} . Assume that $\Psi(\alpha) = \mathbb{E}(e^{\alpha X_1}) < \infty$ on some open interval (a, b) with $b - a > 1$ and that there exists a real number θ such that $\Psi(\theta) = \Psi(\theta + 1)$.

The process $e^{-rt} S_t = S_0 e^{X_t}$ is a martingale under the probability \mathbb{Q} defined as $\mathbb{Q} = Z_t \mathbb{P}$ where $Z_t = \frac{e^{\theta X_t}}{\Psi(\theta)}$

Hence, the value of a contingent claim $h(S_T)$ can be obtained, assuming that the emm chosen by the market is \mathbb{Q} as

$$V_t = e^{-r(T-t)} \mathbb{E}_{\mathbb{Q}}(h(S_T) | \mathcal{F}_t) = e^{-r(T-t)} \frac{1}{\Psi(\theta)} \mathbb{E}_{\mathbb{P}}(h(y e^{r(T-t)+X_{T-t}} e^{\theta X_{T-t}}) |_{y=S_t})$$

A Differential Equation for Option Pricing

Let $S_t = \exp(rt + X_t)$ where X is a (m, σ^2, ν) Lévy process with $\int e^{2x} \nu(dx) < \infty$. Then

$$dS_t = rS_t dt + \sigma S_t dW_t + \int (e^x - 1) S_{t-} (\mathbf{N}(dt, dx) - dt\nu(dx))$$

$$\partial_t C(t, x) + rx \partial_x C + \frac{1}{2} \sigma^2 \partial_{xx} C + \int \nu(dy) (C(t, xe^y) - C(t, x) - x(e^y - 1) \partial_x C) = 0$$

and a terminal condition $C(T, x) = h(x)$

Put-call Symmetry

Let us study a financial market with a riskless asset with constant interest rate r , and a price process (a currency) $S_t = S_0 e^{X_t}$ where X is a Q -Lévy process such that $(Z_t = e^{-(r-\delta)t} S_t / S_0, t \geq 0)$ is a Q -strictly positive martingale with initial value equal to 1. The Q -characteristic triple (m, σ^2, ν) of X is such that

$$m = r - \delta - \sigma^2 / 2 - \int (e^y - 1 - y \mathbb{1}_{\{|y| \leq 1\}}) \nu(dy).$$

Then,

$$\begin{aligned} E_Q(e^{-rT}(S_T - K)^+) &= E_Q(e^{-\delta T} Z_T (S_0 - K S_0/S_T)^+) \\ &= E_{\hat{Q}}(e^{-\delta T} (S_0 - K S_0/S_T)^+) \end{aligned}$$

with $\hat{Q}|_{\mathcal{F}_t} = Z_t Q|_{\mathcal{F}_t}$. The process X is a \hat{Q} -Lévy process, with characteristic exponent $\Psi(\lambda + 1) - \Psi(1)$. The process $S_0/S_t = e^{-X_t}$ is the exponential of the Lévy process, $Y = -X$ which is the dual of the Lévy process X and the characteristic exponent of Y is $\tilde{\Psi}(\lambda) = \Psi(1 - \lambda) - \Psi(1)$. Hence, the following symmetry between call and put prices holds:

$$C_E(S_0, K, r, \delta, T, \Psi) = P_E(K, S_0, \delta, r, T, \tilde{\Psi})$$

Subordinators

A Lévy process which takes values in $[0, \infty[$ (i.e. with increasing paths) is a subordinator.

$$X_t = bt + \int_0^t \int x \mathbf{N}(dx, ds)$$

In this case, the parameters in the Lévy-Khintchine decomposition are $m \geq 0$, $\sigma = 0$ and the Lévy measure ν is a measure on $]0, \infty[$ with

$$\int_{]0, \infty[} (1 \wedge x) \nu(dx) < \infty.$$

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Let Z be a subordinator and X an independent Lévy process. Then, $\tilde{X}_t = X_{Z_t}$ is a Lévy process, called subordinated Lévy process.

Compound Poisson process. A compound Poisson process with $Y_k \geq 0$ is a subordinator.

Gamma process. The Gamma process $G(t; \gamma)$ is a subordinator which satisfies

$$G(t + h; \gamma) - G(t; \gamma) \stackrel{law}{=} \Gamma(h; \gamma).$$

Here $\Gamma(h; \gamma)$ follows the Gamma law. The Gamma process is an increasing Lévy process, hence a subordinator, with one sided Lévy measure

$$\frac{1}{x} \exp\left(-\frac{x}{\gamma}\right) \mathbb{1}_{x>0}.$$

Hitting times Let W be a BM, and

$$T_r = \inf\{t \geq 0 : W_t \geq r\}.$$

The process $(T_r, r \geq 0)$ is a stable $(1/2)$ subordinator, its Lévy measure is $\frac{1}{\sqrt{2\pi} x^{3/2}} \mathbb{1}_{x>0} dx$. Let B be a BM independent of W . The process B_{T_t} is a Cauchy process, its Lévy measure is $dx/(\pi x^2)$.

Changes of Lévy characteristics under subordination

Let X be a (a^X, A^X, ν^X) Lévy process and Z be a subordinator with drift β and Lévy measure ν^Z , independent of X . The process $\tilde{X}_t = X_{Z_t}$ is a Lévy process with characteristic exponent

$$\Phi(u) = i(\tilde{a} \cdot u) + \frac{1}{2} \tilde{A}(u) - \int (e^{i(u \cdot x)} - 1 - i(u \cdot x) \mathbb{1}_{|x| \leq 1}) \tilde{\nu}(dx)$$

with

$$\begin{aligned} \tilde{a} &= \beta a^X + \int \nu^Z(ds) \mathbb{1}_{|x| \leq 1} x \mathbb{P}(X_s \in dx) \\ \tilde{A} &= \beta A^X \\ \tilde{\nu}(dx) &= \beta \nu^X dx + \int \nu^Z(ds) \mathbb{P}(X_s \in dx). \end{aligned}$$

“Vous leur conseillerez donc de faire le calcul. Elles [les grandes personnes] adorent les chiffres: ça leur plaira. Mais ne perdez pas votre temps à ce pensum. C’est inutile. Vous avez confiance en moi.”

Le petit prince, A. de St Exupéry. Gallimard. 1946. p. 59.

Variance-Gamma Model

The **variance Gamma process** is a Lévy process where X_t has a Variance Gamma law $\text{VG}(\sigma, \nu, \theta)$. Its characteristic function is

$$\mathbb{E}(\exp(iuX_t)) = \left(1 - iu\theta\nu + \frac{1}{2}\sigma^2\nu u^2\right)^{-t/\nu}.$$

The Variance Gamma process can be characterized as a **time changed BM** with drift as follows: let W be a BM, $\gamma(t)$ a $G(1/\nu, 1/\nu)$ process. Then

$$X_t = \theta\gamma(t) + \sigma W_{\gamma(t)}$$

is a $\text{VG}(\sigma, \nu, \theta)$ process.

The variance Gamma process is a finite variation process. Hence it is the difference of two increasing processes. Madan et al. showed that it is the difference of two independent Gamma processes

$$X_t = G(t; \mu_1, \gamma_1) - G(t; \mu_2, \gamma_2).$$

Indeed, the characteristic function can be factorized

$$\mathbb{E}(\exp(iuX_t)) = \left(1 - \frac{iu}{\nu_1}\right)^{-t/\gamma} \left(1 + \frac{iu}{\nu_2}\right)^{-t/\gamma}$$

with

$$\begin{aligned} \nu_1^{-1} &= \frac{1}{2} \left(\theta\nu + \sqrt{\theta^2\nu^2 + 2\nu\sigma^2} \right) \\ \nu_2^{-1} &= \frac{1}{2} \left(\theta\nu - \sqrt{\theta^2\nu^2 + 2\nu\sigma^2} \right) \end{aligned}$$

The Lévy density of X is

$$\begin{aligned} \frac{1}{\gamma} \frac{1}{|x|} \exp(-\nu_1|x|) & \quad \text{for } x < 0 \\ \frac{1}{\gamma} \frac{1}{x} \exp(-\nu_2x) & \quad \text{for } x > 0. \end{aligned}$$

The density of X_1 is

$$\frac{2e^{\frac{\theta x}{\sigma^2}}}{\gamma^{1/\gamma} \sqrt{2\pi} \sigma \Gamma(1/2)} \left(\frac{x^2}{\theta^2 + 2\sigma^2/\gamma} \right)^{\frac{1}{2\gamma} - \frac{1}{4}} K_{\frac{1}{\gamma} - \frac{1}{2}} \left(\frac{1}{\sigma^2} \sqrt{x^2(\theta^2 + 2\sigma^2/\gamma)} \right)$$

where K_α is the modified Bessel function.

Stock prices driven by a Variance-Gamma process have dynamics

$$S_t = S_0 \exp \left(rt + X(t; \sigma, \nu, \theta) + \frac{t}{\nu} \ln \left(1 - \theta\nu - \frac{\sigma^2\nu}{2} \right) \right)$$

From $\mathbb{E}(e^{X_t}) = \exp \left(-\frac{t}{\nu} \ln \left(1 - \theta\nu - \frac{\sigma^2\nu}{2} \right) \right)$, we get that $S_t e^{-rt}$ is a martingale. The parameters ν and θ give control on skewness and kurtosis.

The CGMY model, introduced by Carr et al. is an extension of the Variance-Gamma model. The Lévy density is

$$\begin{cases} \frac{C}{x^{Y+1}} e^{-Mx} & x > 0 \\ \frac{C}{|x|^{Y+1}} e^{Gx} & x < 0 \end{cases}$$

with $C > 0$, $M \geq 0$, $G \geq 0$ and $Y < 2$, $Y \notin \mathbb{Z}$.

If $Y < 0$, there is a finite number of jumps in any finite interval, if not, the process has infinite activity. If $Y \in [1, 2[$, the process is with infinite variation. This process is also called KoBol.

Double Exponential Model

The Model A particular Lévy model is the double exponential jumps model, introduced by Kou and Wang. In this model

$$X_t = \mu t + \sigma W_t + \sum_{i=1}^{N_t} Y_i,$$

where W is a Brownian motion independent of N and $\sum_{i=1}^{N_t} Y_i$ is a compound Poisson process. The r.v.'s Y_i are i.i.d., independent of N and W and the density of the law of Y_1 is

$$f(x) = p\eta_1 e^{-\eta_1 x} \mathbb{1}_{\{x>0\}} + (1-p)\eta_2 e^{\eta_2 x} \mathbb{1}_{\{x<0\}}.$$

The Lévy measure of X is $\nu(dx) = \lambda f(x)dx$.

Here, η_i are positive real numbers, and $p \in [0, 1]$. With probability p (resp. $(1-p)$), the jump size is positive (resp. negative) with exponential law with parameter η_1 (resp η_2).

It is easy to prove that

$$\mathbb{E}(Y_1) = \frac{p}{\eta_1} - \frac{1-p}{\eta_2}, \quad \text{var}(Y_1) = \frac{p}{\eta_1^2} + \frac{1-p}{\eta_2^2} + p(1-p) \left(\frac{1}{\eta_1} + \frac{1}{\eta_2} \right)^2$$

and that, for $\eta_1 > 1$, $\mathbb{E}(e^{Y_1}) = p \frac{\eta_1}{\eta_1-1} + (1-p) \frac{\eta_2}{1+\eta_2}$. Moreover

$$\mathbb{E}(e^{iuX_t}) = \exp \left(t \left\{ -\frac{1}{2} \sigma^2 u^2 + ibu + \lambda \left(\frac{p\eta_1}{\eta_1 - iu} + \frac{(1-p)\eta_2}{\eta_2 + iu} - 1 \right) \right\} \right),$$

where $b = \mu + \lambda \mathbb{E}(Y_1 \mathbb{1}_{|Y_1| \leq 1}) =$

$\mu + \lambda p \left(\frac{1-e^{-\eta_1}}{\eta_1} - e^{-\eta_1} \right) - \lambda(1-p) \left(\frac{1-e^{-\eta_2}}{\eta_2} - e^{-\eta_2} \right)$. The Laplace exponent of X , i.e., the function Ψ such that $\mathbb{E}(e^{\beta X_t}) = \exp(\Psi(\beta)t)$ is defined for $-\eta_2 < \beta < \eta_1$ as

$$\Psi(\beta) = \beta\mu + \frac{1}{2}\beta^2\sigma^2 + \lambda \left(\frac{p\eta_1}{\eta_1 - \beta} + \frac{(1-p)\eta_2}{\beta + \eta_2} - 1 \right).$$

Change of probability Let $S_t = S_0 e^{rt+X_t}$ where $X_t = (\mu - \frac{1}{2}\sigma^2)t + \sigma W_t + \sum_{i=1}^{N_t} Y_i$. Then, setting $V_i = e^{Y_i}$, using an Escher transform, the process $S_t e^{-rt}$ will be a \mathbb{Q} martingale with $\mathbb{Q}|_{\mathcal{F}_t} = L_t d\mathbb{P}|_{\mathcal{F}_t}$ and $L_t = \frac{e^{\alpha X_t}}{\mathbb{E}(e^{\alpha X_t})}$, for α such that $\Psi(\alpha) = \Psi(\alpha + 1)$. Under \mathbb{Q} , the Lévy measure of X is $\hat{\nu}(dx) = e^{\alpha x} \nu(dx) = e^{\alpha x} \lambda f(x) dx = \hat{\lambda} \hat{f}(x) dx$ where, after some standard computations

$$\begin{aligned} \hat{f}(x) &= \left(\hat{p} \hat{\eta}_1 e^{-\hat{\eta}_1 x} \mathbf{1}_{\{x>0\}} + (1 - \hat{p}) \hat{\eta}_2 e^{\hat{\eta}_2 x} \mathbf{1}_{\{x<0\}} \right) . \\ \hat{\eta}_1 &= \eta_1 - \alpha, \quad \hat{\eta}_2 = \eta_2 + \alpha \\ \hat{\lambda} &= \lambda \left(\frac{p\eta_1}{\eta_1 - \alpha} + \frac{(1-p)\eta_2}{\eta_2 + \alpha} \right) \\ \hat{p} &= p\eta_1 \frac{\eta_2 + \alpha}{\alpha p\eta_1 + \eta_2(\eta_1 - \alpha + \alpha p\eta_1)} \end{aligned}$$

In particular, the process X is a double exponential process under \mathbb{Q} .

Hitting times For any $x > 0$

$$\mathbb{P}(\tau_b \leq t, X_{\tau_b} - b \geq x) = e^{-\eta_1 x} \mathbb{P}(\tau_b \leq t, X_{\tau_b} - b \geq 0)$$

PROOF:

The infinitesimal generator of X is

$$\mathcal{L}f = \frac{1}{2}\sigma^2\partial_{xx}f + \mu\partial_xf + \lambda \int_{\mathbb{R}} (f(x+y) - f(x))\nu(dx)$$

Let $T_x = \inf\{t : X_t \geq x\}$. Then Kou and Wang establish that, for $r > 0$ and $x > 0$,

$$\mathbb{E}(e^{-rT_x}) = \frac{\eta_1 - \beta_1}{\eta_1} \frac{\beta_2}{\beta_2 - \beta_1} e^{-x\beta_1} + \frac{\beta_2 - \eta_1}{\eta_1} \frac{\beta_1}{\beta_2 - \beta_1} e^{-x\beta_2}$$

$$\mathbb{E}(e^{-rT_x} \mathbf{1}_{X_{T_x} - x > y}) = e^{\eta_1 y} \frac{\eta_1 - \beta_1}{\eta_1} \frac{\beta_2 - \eta_1}{\beta_2 - \beta_1} (e^{-x\beta_1} - e^{-x\beta_2})$$

$$\mathbb{E}(e^{\theta X_{T_x} - rT_x}) = e^{\theta x} \left(\frac{\eta_1 - \beta_1}{\beta_2 - \beta_1} \frac{\beta_2 - \theta}{\eta_1 - \theta} e^{-x\beta_1} + \left(\frac{\beta_2 - \eta_1}{\beta_2 - \beta_1} \frac{\beta_1 - \theta}{\eta_1 - \theta} e^{-x\beta_2} \right) \right)$$

where $0 < \beta_1 < \eta_1 < \beta_2$ are roots of $G(\beta) = r$. The method is based on

finding an explicit solution of $\mathcal{L}u = ru$ where \mathcal{L} is the infinitesimal generator of the process X .

Thank you for your attention.