



Computational Complexity and Evolutionary Computation

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Theory... Why should you care?

- foundations — **firm ground**
- Proofs provide insights and understanding.
- generality — **wide applicability**
- **knowledge** vs. **beliefs**
- fundamental limitations — **saves time**
- much improved teaching
- “*There is nothing more practical than a good theory.*”

Topics and Structure

- Introduction and Motivation
- (an extremely short) introduction to evolutionary algorithms
- overview of topics in theory (as presented here today)
- analytical tools and methods – and how to apply them
 - fitness-based partitions
 - expectations and deviations
 - simple general lower bounds
 - expected multiplicative decrease in distance
 - drift analysis
 - random walks and cover times
 - typical runs
 - instructive example functions
- general limitations
 - NFL
 - black box complexity

Aims and Goals of this Tutorial

- **provide an overview** of
 - goals and topics
 - methods and their applications
- **enhance your ability** to
 - read, understand, and appreciate such papers
 - make use of the results obtained this way
- **enable you** to
 - apply the methods to your problems
 - produce such results yourself
- **explain**
 - what is doable with the currently known methods
 - where there is need for more advanced methods
- **entertain**

Evolution Strategies (Bienert, Rechenberg, Schwefel)

- developed in the '60s / '70s of the last century.
- continuous optimization problems, rely on mutation.

Genetic Algorithms (Holland)

- developed in the '60s / '70s.
- binary problems, rely on crossover.

Genetic Programming (Koza)

- developed in the '90s.
- try to build good “computer programs”.

Nowadays

- more general view \Rightarrow evolutionary algorithms.

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Principle

- follow Darwin's principle (survival of the fittest).
- work with a set of solutions called population.
- parent population produces offspring population by variation operators (mutation, crossover).
- select individuals from the parents and children to create new parent population.

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Bionics/Engineering

- evolution is a “natural” enhancing process.
- bionics: algorithmic simulation \Rightarrow “enhancing” algorithm.
- used for optimization.

Biology

- **evolutionary** algorithms.
- understanding model of natural evolution.

Computer Science

- evolutionary **algorithms**.
- successful applications.
- theoretical understanding.

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Basic EA

- 1 compute an initial population $P = \{X_1, \dots, X_\mu\}$.
- 2 while (not termination condition)
 - produce an offspring population $P' = \{Y_1, \dots, Y_\lambda\}$ by crossover and/or mutation.
 - create new parent population P by selecting μ individuals from P and P' .

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Design

Important issues

- representation
- crossover operator
- mutation operator
- selection method

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Crossover operator

Aim

- two individuals x and y should produce a new solution z .

1-point Crossover

- choose a position $p \in \{1, \dots, n\}$ uniformly at random
- set $z_i = x_i$ for $1 \leq i \leq p$
- set $z_i = y_i$ for $p < i \leq n$

Uniform Crossover

- set z_i equally likely to x_i or y_i
- if $x_i = y_i$ then $z_i = x_i = y_i$
- if $x_i \neq y_i$ then $\text{Prob}(z_i = x_i) = \text{Prob}(z_i = y_i) = 1/2$

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Representation

Properties

- representation has to fit to the considered problem.
- small change in the representation \implies small change in the solution (locality).
- often direct representation works fine.

Mainly in this talk

- search space $\{0, 1\}^n$.
- individuals are bitstrings of length n .

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Mutation

Aim

- produce from a current solution x a new solution z .

Some Possibilities

- flip one randomly chosen bit of x to obtain z .
- flip each bit of x with probability p to obtain z (often $p = 1/n$).

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Selection methods

Fitness-proportional selection

- choose new population from a set of r individuals $\{x_1, \dots, x_r\}$.
- probability to choose x_i in the next selection step is $f(x_i) / (\sum_{j=1}^r f(x_j))$.
- μ individuals are selected in this way.

(μ, λ) -selection

- μ parents produce λ children.
- select μ best individuals from the children.

$(\mu + \lambda)$ -selection

- μ parents produce λ children.
- select μ best individuals from the parents and children.

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Simple algorithms

$(1+1)$ EA

- 1 Choose $s \in \{0, 1\}^n$ randomly.
- 2 Produce s' by flipping each bit of s with probability $1/n$.
- 3 Replace s by s' if $f(s') \geq f(s)$.
- 4 Repeat Steps 2 and 3 forever.

RLS

- 1 Choose $s \in \{0, 1\}^n$ randomly.
- 2 Produce s' from s by flipping **one** randomly chosen bit.
- 3 Replace s by s' if $f(s') \geq f(s)$.
- 4 Repeat Steps 2 and 3 forever.

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$(\mu + \lambda)$ -EA

$(\mu + \lambda)$ -EA

- 1 Choose μ individuals uniformly at random from $\{0, 1\}^n$.
- 2 Produce λ children by mutation.
- 3 Apply $(\mu + \lambda)$ -selection to parents and children.
- 4 Go to 2.)

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Topics in Theory

The most pressing open question depends very much on what you are interested in.

What you are interested in depends very much on who you are.

You may be

- **biologist** What is evolution and how does it work?
- **engineer** How do I solve my problem with an EA?
- **computer scientist** What can evolutionary algorithms do?

Evolutionary algorithms are

- **a model of natural evolution**
- **a robust general purpose problem solver**
- **randomized algorithms**

here and today computer scientist's point of view

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Algorithms in Computer Science

Two branches

- ① design and analysis of algorithms
"How long does it take to solve this problem?"
- ② complexity theory
"How much time is needed to solve this problem?"

For evolutionary algorithms

- ① analysis (and design) of evolutionary algorithms
"What's the expected optimization time of this EA for this problem?"
- ② general limitations — NFL and black box complexity "How much time is needed to solve this problem?"

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Fitness-Based Partitions

very simple, yet often powerful method for upper bounds

first for (1+1)-EA only

Observation due to plus-selection fitness is monotone increasing

Idea for each fitness value v , find probability p_v to increase fitness

Observation $E(\text{time to increase fitness from } v) = \frac{1}{p_v}$

Observation $E(T) = \sum_v \frac{1}{p_v}$

a bit more general group fitness values

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"Time" and Evolutionary Algorithms

At the end of the day, time is wall clock time.

in computer science more convenient: #computation steps
requires formal model of computation (Turing machine, ...)

typical for evolutionary algorithms black box optimization

fitness function not known to algorithm

gathers knowledge only by means of function evaluations

often

- evolutionary algorithm's core rather simple and fast
- evaluation of fitness function costly and slow

thus "time" = #fitness function evaluations often appropriate

Definition

Optimization Time T = #fitness function evaluations until an optimal search point is sampled for the first time

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Method of Fitness-Based Partitions

Definition

For $f: \{0, 1\}^n \rightarrow \mathbb{R}$, $L_0, L_1, \dots, L_k \subseteq \{0, 1\}^n$ with

- ① $\forall i \neq j \in \{0, 1, \dots, k\}: L_i \cap L_j = \emptyset$
- ② $\bigcup_{i=0}^k L_i = \{0, 1\}^n$
- ③ $\forall i < j \in \{0, 1, \dots, k\}: \forall x \in L_i: \forall y \in L_j: f(x) < f(y)$
- ④ $L_k = \{x \in \{0, 1\}^n \mid f(x) = \max \{f(y) \mid y \in \{0, 1\}^n\}\}$

is called an f -based partition.

Remember An f -based partition partitions the search space in accordance to fitness values grouping fitness values arbitrarily.

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Upper Bounds with f -Based Partitions

Theorem

Consider $(1+1)$ -EA on $f: \{0,1\}^n \rightarrow \mathbb{R}$ and an f -based partition L_0, L_1, \dots, L_k .

Let $s_i := \min_{x \in L_i} \sum_{j=i+1}^k \sum_{y \in L_j} \left(\frac{1}{n}\right)^{H(x,y)} \left(1 - \frac{1}{n}\right)^{n-H(x,y)}$
for all $i \in \{0, 1, \dots, k-1\}$.

$$E(T_{(1+1)\text{-EA},f}) \leq \sum_{i=0}^{k-1} \frac{1}{s_i}$$

Hint most often, very simple lower bounds for s_i suffice

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Example: Result for a Class of Functions

Definition

$f: \{0,1\}^n \rightarrow \mathbb{R}$ is called **linear**

$$\Leftrightarrow \exists w_0, w_1, \dots, w_n \in \mathbb{R}: \forall x \in \{0,1\}^n: f(x) = w_0 + \sum_{i=1}^n w_i \cdot x[i]$$

Consider $(1+1)$ -EA on linear function $f: \{0,1\}^n \rightarrow \mathbb{R}$.

For $(1+1)$ -EA, w.l.o.g. $w_0 = 0, w_1 \geq w_2 \geq \dots \geq w_n \geq 0$

First Step define f -based partition

$$L_i := \left\{ x \in \{0,1\}^n \mid \sum_{j=1}^i w_j \leq f(x) < \sum_{j=1}^{i+1} w_j \right\}, 0 \leq i \leq n$$

Second Step find lower bounds for s_i

Observation There is always at least 1-bit-mutation for leaving L_i .

$$s_i \geq \frac{1}{n} \left(1 - \frac{1}{n}\right)^{n-1} \geq \frac{1}{en}$$

Third Step $E(T_{(1+1)\text{-EA},f}) \leq \sum_{i=0}^{n-1} en = en^2 = O(n^2)$

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Very Simple Example

$(1+1)$ -EA on ONEMAX

$$\left(\text{ONEMAX}(x) = \sum_{i=1}^n x[i] \right)$$

First Step define f -based partition

trivial for each fitness value one L_i

$$L_i := \{x \in \{0,1\}^n \mid \text{ONEMAX}(x) = i\}, 0 \leq i \leq n$$

Second Step find lower bounds for s_i

Observation It suffices to flip any 0-bit from the $n-i$ 0-bits.

$$s_i \geq \binom{n-i}{1} \frac{1}{n} \left(1 - \frac{1}{n}\right)^{n-1} \geq \frac{n-i}{en}$$

$$\left(\left(1 - \frac{1}{n}\right)^{n-1} \geq \frac{1}{e} \geq \left(1 - \frac{1}{n}\right)^n \right)$$

Third Step compute upper bound

$$E(T_{(1+1)\text{-EA}, \text{ONEMAX}}) \leq \sum_{i=0}^{n-1} \frac{en}{n-i} = en \cdot \sum_{i=1}^n \frac{1}{i} = O(n \log n)$$

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Generalizing the Method

Idea not restricted to $(1+1)$ -EA, only.

Consider $(1+\lambda)$ -EA on LEADINGONES.

$$\left(\text{LEADINGONES}(x) = \sum_{i=1}^n \prod_{j=1}^i x[j] \right)$$

First Step define f -based partition

trivial for each fitness value one L_i

$$L_i := \{x \in \{0,1\}^n \mid \text{LEADINGONES}(x) = i\}, 0 \leq i \leq n$$

For the $(1+\lambda)$ -EA, we **re-define** the s_i .

$s_i := \text{Prob}(\text{leave } L_i \text{ in one generation})$

Observation $E(T_{(1+\lambda)\text{-EA},f}) \leq \lambda \cdot \sum_{i=0}^{k-1} \frac{1}{s_i}$

Jansen/De Jong/Wegener (2005): On the choice of the offspring population size in evolutionary algorithms. *Evolutionary Computation* 13(4):413-440

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$(1 + \lambda)$ -ES on LEADINGONES

Second Step find lower bounds for s_i

Observation It suffices to flip exactly the leftmost 0-bit.

$$s_i \geq 1 - \left(1 - \frac{1}{en}\right)^\lambda \geq 1 - e^{-\lambda/(en)}$$

Case Inspection Case 1 $\lambda \geq en$

$$s_i \geq 1 - \frac{1}{e}$$

Case Inspection Case 2 $\lambda < en$

$$s_i \geq \frac{\lambda}{2en}$$

Third Step compute upper bound

$$\begin{aligned} E(T_{(1+\lambda)\text{-EA, LEADINGONES}}) &\leq \lambda \cdot \left(\sum_{i=0}^{n-1} \frac{1}{1-e^{-1}} \right) + \left(\sum_{i=0}^{n-1} \frac{2en}{\lambda} \right) \\ &= O\left(\lambda \cdot \left(n + \frac{n^2}{\lambda}\right)\right) = O(\lambda \cdot n + n^2) \end{aligned}$$

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Markov Inequality and Chernoff Bounds

Theorem (Markov Inequality)

$X \geq 0$ random variable, $s > 0$

$$\text{Prob}(X \geq s \cdot E(X)) \leq \frac{1}{s}$$

Theorem (Chernoff Bounds)

Let $X_1, X_2, \dots, X_n: \Omega \rightarrow \{0, 1\}$ independent random variables with

$$\forall i \in \{1, 2, \dots, n\}: 0 < \text{Prob}(X_i = 1) < 1.$$

$$\text{Let } X := \sum_{i=1}^n X_i.$$

$$\forall \delta > 0: \text{Prob}(X > (1 + \delta) \cdot E(X)) < \left(\frac{e^\delta}{(1+\delta)^{1+\delta}}\right)^{E(X)}$$

$$\forall 0 < \delta < 1: \text{Prob}(X < (1 - \delta) \cdot E(X)) < e^{-E(X)\delta^2/2}$$

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Some Useful Background Knowledge

a short detour into very basic probability theory

We already know, we care for $E(T)$ — an expected value.

Often, we care for the probability to deviate from an expected value.

A lot is known about this, we should make use of this.

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A Very Simple Application

Consider $x \in \{0, 1\}^{100}$ selected uniformly at random

more formal for $i \in \{1, 2, \dots, 100\}$: $B_i := \begin{cases} 1 & i\text{-th bit is 1} \\ 0 & \text{otherwise} \end{cases}$

with $\text{Prob}(B_i = 0) = \text{Prob}(B_i = 1) = \frac{1}{2}$

$$B := \sum_{i=1}^{100} B_i \quad \text{clearly} \quad E(B) = 50$$

What is the probability to have at least 75 1-bits?

Markov $\text{Prob}(B \geq 75) = \text{Prob}(B \geq \frac{3}{2} \cdot 50) \leq \frac{2}{3}$

Chernoff $\text{Prob}(B \geq 75) = \text{Prob}(B \geq (1 + \frac{1}{2}) \cdot 50)$
 $\leq \left(\frac{\sqrt{e}}{(3/2)^{3/2}}\right)^{50} < 0.0045$

Truth $\text{Prob}(B \geq 75) = \sum_{i=75}^{100} \binom{100}{i} 2^{-100}$
 $= \frac{89,310,453,796,450,805,935,325}{316,912,650,057,057,350,374,175,801,344} < 0.000000282$

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The Law of Total Probability

Theorem (Law of Total Probability)

Let B_i with $i \in I$ be a partition of some probability space Ω .
 $\forall A \subseteq \Omega: \text{Prob}(A) = \sum_{i \in I} \text{Prob}(A \mid B_i) \cdot \text{Prob}(B_i)$

immediate consequence $\text{Prob}(A) \geq \text{Prob}(A \mid B) \cdot \text{Prob}(B)$

Useful for lower bounds

when some event “determines” expected optimization time

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Lower bound for OneMax

Chernoff bounds

- Expected number of 1-bits in initial solution is $n/2$.
- At least $n/3$ 0-bits with probability $1 - e^{-\Omega(n)}$ (Chernoff).

Lower Bound

- Probability that at least one 0-bit has not been flipped during $t = (n-1) \ln n$ steps is

$$1 - (1 - (1 - 1/n)^{(n-1) \ln n})^{n/3} \geq 1 - e^{-1/3} = \Omega(1).$$

- Expected optimization time for ONEMAX is $\Omega(n \log n)$

Generalization

- $\Omega(n \log n)$ for each function with poly. number of optima.

A Very Simple Example

Consider $(1+1)$ -EA on $f: \{0, 1\}^n \rightarrow \mathbb{R}$
 with $f(x) := \begin{cases} n - \frac{1}{2} & \text{if } x = 0^n \\ \text{ONEMAX}(x) & \text{otherwise} \end{cases}$.

Theorem

$$\mathbb{E}(T_{(1+1)\text{-EA}}, f) = \Omega\left(\left(\frac{n}{2}\right)^n\right)$$

Proof.

Define event B : $(1+1)$ -EA initializes with $x = 0^n$
 clearly $\text{Prob } B = 2^{-n}$

Observation $\mathbb{E}(T_{(1+1)\text{-EA}}, f \mid B) = n^n$
 since all bits have to flip simultaneously

Law of Total Probability

$$\mathbb{E}(T_{(1+1)\text{-EA}}, f) \geq n^n \cdot 2^{-n} = \left(\frac{n}{2}\right)^n$$



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Coupon Collector's Theorem

Proposition

Given n different coupons. Choose at each trial a coupon uniformly at random. Let X be a random variable describing the number of trials required to choose each coupon at least once. Then

$$\mathbb{E}(X) = nH_n$$

holds, where H_n denotes the n th Harmonic number, and

$$\lim_{n \rightarrow \infty} \text{Prob}(X \leq n(\ln n - c)) = e^{-e^c}$$

holds for each constant $c \in \mathbb{R}$.

Expected multiplicative distance decrease

Basic idea

- Assumption: Function values are integers.
- Define a set O of l operations to obtain an optimal solution.
- Average gain of these l operations is $\frac{f(x_{opt}) - f(x)}{l}$.

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Expected multiplicative distance decrease

Upper bound

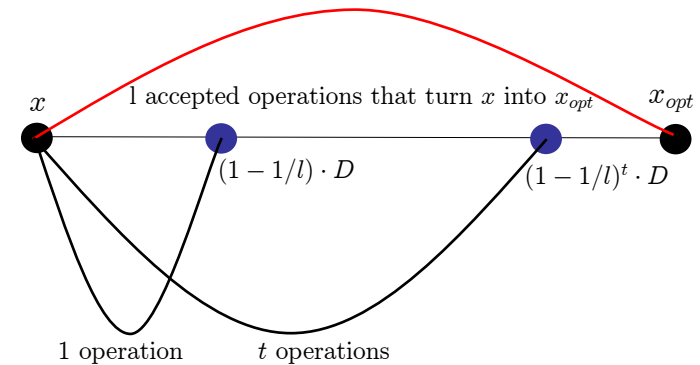
- Let $d_{max} = \max_{x \in \{0,1\}^n} f(x_{opt}) - f(x)$.
- 1 operation: expected distance at most $(1 - 1/l) \cdot d_{max}$.
- t operations: expected distance at most $(1 - 1/l)^t \cdot d_{max}$.
- Expected number of $O(l \cdot \log d_{max})$ operations to reach optimum.
- Assume: expected time for each operation is at most r .
- Upper bound $O(r \cdot l \cdot \log d_{max})$ to obtain an optimal solution.

F. Neumann, I. Wegener: Randomized local search, evolutionary algorithms, and the minimum spanning tree problem, GECCO 2004.

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Figure: Distance Decrease

$$D = f(x_{opt}) - f(x)$$



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Example

Linear Functions

- $f(x) = w_1x_1 + w_2x_2 + \dots + w_nx_n$.
- $w_i \in \mathbb{Z}$.
- $w_{max} = \max_i w_i$.

Upper bound

- Consider all operations that flip a single bit.
- Each necessary operation is accepted.
- $d_{max} = n \cdot w_{max}$.
- Expected number of operations $O(n \log d_{max})$.
- Waiting time for a single bit flip $O(1)$.
- Upper bound $O(n(\log n + \log w_{max}))$.
- If $w_{max} = \text{poly}(n)$, upper bound $O(n \log n)$.

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A More Flexibel Proof Method

Sad Facts

- f -based partitions restricted to “well behaving” functions
- direct lower bound often too difficult

How can we find a more flexibel method?

Observation f -based partition measure **progress** by $f(x_{t+1}) - f(x_t)$

Idea consider a more general **measure of progress**

Define distance $d: Z \rightarrow \mathbb{R}_0^+$, (Z set of all populations)
with $d(P) = 0 \Leftrightarrow P$ contains optimal solution

Caution “Distance” need **not** be a metric!

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Upper Bound Drift Theorem

Drift Theorem (Upper Bound)

Let A be some evolutionary algorithm, P_t its t -th population, f some function, Z the set of all possible populations, $d: Z \rightarrow \mathbb{R}_0^+$ some distance measure with
 $d(P) = 0 \Leftrightarrow P$ contains an optimum of f ,
 $M = \max\{d(P) \mid P \in Z\}$, $D_t := d(P_{t-1}) - d(P_t)$,
 $\Delta := \min\{E(D_t \mid T \geq t) \mid t \in \mathbb{N}_0\}$.
 $\Delta > 0 \Rightarrow E(T_{A,f}) \leq M/\Delta$

Proof

Observe $M \geq E\left(\sum_{t=1}^T D_t\right)$

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Drift

Define distance $d: Z \rightarrow \mathbb{R}_0^+$, (Z set of all populations)
with $d(P) = 0 \Leftrightarrow P$ contains optimal solution

Observation $T = \min\{t \mid d(P_t) = 0\}$

Consider maximum distance $M := \max\{d(P) \mid P \in Z\}$,
decrease in distance $D_t := d(P_{t-1}) - d(P_t)$

Definition $E(D_t \mid T \geq t)$ is called **drift**.

Pessimistic point of view $\Delta := \min\{E(D_t \mid T \geq t) \mid t \in \mathbb{N}_0\}$

Drift Theorem (Upper Bound) $\Delta > 0 \Rightarrow E(T) \leq M/\Delta$

He/Yao (2004): A study of drift analysis for estimating computation time of evolutionary algorithms. *Natural Computing* 3(1):21–35

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Proof of the Drift Theorem (Upper Bound)

$$\begin{aligned}
 M &\geq E\left(\sum_{t=1}^T D_t\right) = \sum_{t=1}^{\infty} \text{Prob}(T = t) \cdot E\left(\sum_{i=1}^T D_i \mid T = t\right) \\
 &= \sum_{t=1}^{\infty} \text{Prob}(T = t) \cdot \sum_{i=1}^t E(D_i \mid T = t) \\
 &= \sum_{t=1}^{\infty} \sum_{i=1}^t \text{Prob}(T = t) \cdot E(D_i \mid T = t) \\
 &= \sum_{i=1}^{\infty} \sum_{t=i}^{\infty} \text{Prob}(T = t) \cdot E(D_i \mid T = t)
 \end{aligned}$$

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$$\begin{aligned}
 &\geq \sum_{i=1}^{\infty} \sum_{t=i}^{\infty} \text{Prob}(T=t) \cdot \mathbb{E}(D_i \mid T=t) \\
 &= \sum_{i=1}^{\infty} \sum_{t=i}^{\infty} \text{Prob}(T \geq i) \cdot \text{Prob}(T=t \mid T \geq i) \cdot \mathbb{E}(D_i \mid T=t) \\
 &= \sum_{i=1}^{\infty} \text{Prob}(T \geq i) \sum_{t=i}^{\infty} \text{Prob}(T=t \mid T \geq i) \cdot \mathbb{E}(D_i \mid T=t \wedge T \geq i) \\
 &= \sum_{i=1}^{\infty} \text{Prob}(T \geq i) \sum_{t=1}^{\infty} \text{Prob}(T=t \mid T \geq i) \cdot \mathbb{E}(D_i \mid T=t \wedge T \geq i) \\
 &= \sum_{i=1}^{\infty} \text{Prob}(T \geq i) \mathbb{E}(D_i \mid T \geq i) \geq \Delta \cdot \sum_{i=1}^{\infty} \text{Prob}(T \geq i) = \Delta \cdot \mathbb{E}(T) \\
 &\text{thus } \mathbb{E}(T) \leq \frac{M}{\Delta} \quad \square
 \end{aligned}$$

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Consider $(1+1)$ -EA on linear function $f: \{0,1\}^n \rightarrow \mathbb{R}$

now with drift analysis

remember $f(x) = \sum_{i=1}^n w_i \cdot x[i]$
with $w_1 \geq w_2 \geq \dots \geq w_n > 0$

Define $d(x) := \ln \left(1 + 2 \sum_{i=1}^{n/2} (1 - x[i]) + \sum_{i=(n/2)+1}^n (1 - x[i]) \right)$

Observe

$$M = \max \{d(x) \mid x \in \{0,1\}^n\} = \ln \left(1 + \frac{3}{2}n \right) = \Theta(\ln n)$$

He/Yao (2004): A study of drift analysis for estimating computation time of evolutionary algorithms. *Natural Computing* 3(1):21–35

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Consider $(1, n)$ -EA on LEADINGONES

Theorem

$$\mathbb{E}(T_{(1, n)\text{-EA, LEADINGONES}}) = O(n^2)$$

Proof.

$$d(x) := n - \text{LEADINGONES}(x) \rightsquigarrow M = n$$

$$\begin{aligned}
 &\mathbb{E}(d(x_{t-1}) - d(x_t) \mid T > t) \\
 &\geq 1 \cdot \left(1 - \left(1 - \frac{1}{en} \right)^n \right) - n \cdot \left(1 - \left(1 - \frac{1}{n} \right)^n \right) \\
 &= \Omega(1)
 \end{aligned}$$

thus $\mathbb{E}(T) = O(n)$

thus $\mathbb{E}(T_{(1, n)\text{EA, LEADINGONES}}) = n \cdot \mathbb{E}(T) = O(n^2)$ ◻

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$$d(x) := \ln \left(1 + 2 \sum_{i=1}^{n/2} (1 - x[i]) + \sum_{i=(n/2)+1}^n (1 - x[i]) \right)$$

Need lower bound for $\mathbb{E}(d(x_{t-1}) - d(x_t) \mid T \geq t)$

Observe minimal for $x_{t-1} = 011 \dots 1$ or $\underbrace{11 \dots 1}_{\text{left}} \underbrace{01 \dots 1}_{\text{right}}$

Consider separately and do tedious calculations...

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Calculation for $011 \dots 1$

$$\begin{aligned}
 & \mathbb{E}(d(x_{t-1}) - d(x_t) \mid T \geq t) \\
 &= \frac{1}{n} \left(1 - \frac{1}{n}\right)^{n-1} (\ln(3) - \ln(1)) \\
 &\quad + \binom{n/2}{1} \left(\frac{1}{n}\right)^2 \left(1 - \frac{1}{n}\right)^{n-2} (\ln(3) - \ln(1+1)) \\
 &\quad - \sum_{b_r=3}^{n/2} \binom{n/2}{b_r} \left(\frac{1}{n}\right)^{1+b_r} \left(1 - \frac{1}{n}\right)^{n-b_r-1} (\ln(1+b_r) - \ln(3)) \\
 &\quad - \sum_{b_l=1}^{(n/2)-1} \sum_{b_r=0}^{n/2} \binom{(n/2)-1}{b_l} \binom{n/2}{b_r} \left(\frac{1}{n}\right)^{1+b_l+b_r} \left(1 - \frac{1}{n}\right)^{n-b_l-b_r-1} \\
 &\quad \quad (\ln(1+2b_l+b_r) - \ln(3)) \\
 &= \Omega\left(\frac{1}{n}\right)
 \end{aligned}$$

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Result for $(1+1)$ -EA on General Linear Functions

We have

- $d(x) := \ln \left(1 + 2 \sum_{i=1}^{n/2} (1 - x[i]) + \sum_{i=(n/2)+1}^n (1 - x[i]) \right)$
- $d(x) \leq \ln(1 + (3/2)n) = O(\log n)$
- $\mathbb{E}(d(x_{t-1}) - d(x_t) \mid T \geq t) = \Omega(1/n)$

together $\mathbb{E}(T_{(1+1)\text{-EA},f}) = O(n \log n)$ for any linear f

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Calculation for $1^{n/2}01^{(n/2)-1}$

$$\begin{aligned}
 & \mathbb{E}(d(x_{t-1}) - d(x_t) \mid T \geq t) \\
 &= \frac{1}{n} \left(1 - \frac{1}{n}\right)^{n-1} (\ln(2) - \ln(1)) \\
 &\quad - \binom{n/2}{1} \left(\frac{1}{n}\right)^2 \left(1 - \frac{1}{n}\right)^{n-2} (\ln(1+2) - \ln(2)) \\
 &\quad - \sum_{b_r=2}^{(n/2)-1} \binom{(n/2)-1}{b_r} \left(\frac{1}{n}\right)^{1+b_r} \left(1 - \frac{1}{n}\right)^{n-b_r-1} (\ln(1+b_r) - \ln(2)) \\
 &= \Omega\left(\frac{1}{n}\right)
 \end{aligned}$$

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Drift Analysis of Lower Bounds

We have drift analysis for upper bounds

How can we obtain lower bounds when analyzing drift?

Idea Check proof of drift theorem (upper bound).

Can inequalities be reversed?

Remember
$$\begin{aligned}
 M &\geq \mathbb{E} \left(\sum_{t=1}^T D_t \right) = \dots = \sum_{i=1}^{\infty} \text{Prob}(T \geq i) \mathbb{E}(D_i \mid T \geq i) \\
 &\geq \Delta \cdot \sum_{i=1}^{\infty} \text{Prob}(T \geq i) = \Delta \cdot \mathbb{E}(T)
 \end{aligned}$$

with

- $M = \max\{d(P) \mid P \in Z\}$
- $\Delta = \min\{\mathbb{E}(d(P_{t-1}) - d(P_t) \mid T \geq t)\}$

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observation only two inequalities need to be reversed

- ① $M \geq \sum \dots$ with $M = \max\{d(P) \mid P \in Z\}$
- ② $\sum \dots \geq \Delta_l \cdot \sum \dots$ with $\Delta_l = \min\{E(d(P_{t-1}) - d(P_t) \mid T \geq t)\}$

clearly for lower bound $\Delta_u = \max\{E(d(P_{t-1}) - d(P_t) \mid T \geq t)\}$
 sensible and sufficient for “ \leq ”

clearly for lower bound instead of $M \min\{d(P) \mid P \in Z\}$
 possible and sufficient for “ \leq ”,
 but pointless, since $\min\{d(P) \mid P \in Z\} = 0$

He/Yao (2004): A study of drift analysis for estimating computation time of evolutionary algorithms. *Natural Computing* 3(1):21–35

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Define trivial distance
 $d(x) := n - \text{LEADINGONES}(x)$

Observation necessary for decrease of distance
 left-most 0-bit flips

thus $\text{Prob}(\text{decrease distance}) \leq \frac{1}{n}$

How can we bound the decrease in distance?

Observation trivially, by n — not useful

better question How can we bound the expected decrease in distance?

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clearly $E\left(\sum_{t=1}^T D_t\right)$ fixed, if initial population is known

thus lower bound on $d(P_0)$ yields lower bound on $E(T)$

making this concrete

- $E(T \mid d(P_0) \geq M_u) \geq M_u / \Delta_u$
- $E(T) \geq \text{Prob}(d(P_0) \geq M_u) \cdot E(T \mid d(P_0) \geq M_u) \geq \text{Prob}(d(P_0) \geq M_u) \cdot M_u / \Delta_u$
- $E(T) \geq \sum \text{Prob}(d(P_0) \geq d) \cdot d / \Delta_u \geq E(d(P_0)) / \Delta_u$

thus drift analysis suitable as method for upper and lower bounds

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Note decrease in distance $\hat{=}$ increase in fitness

Observation two sources for increase in fitness

- ① the left-most 0-bit
- ② bits to the right of this bits that happen to be 1-bits

Observation bits to the right of the left-most 0-bit
 have no influence on selection and
 never had influence on selection

Claim These bits are uniformly distributed.

obvious holds after random initialization

Claim standard bit mutations do not change this

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Standard Bit Mutations on Uniformly Distributed Bits

Claim $\forall t \in \mathbb{N}_0: \forall x \in \{0, 1\}^n: \text{Prob}(x_t = x) = 2^{-n}$

clearly holds for $t = 0$

$$\begin{aligned}
 \text{Prob}(x_t = x) &= \sum_{x' \in \{0, 1\}^n} \text{Prob}((x_{t-1} = x') \wedge (\text{mut}(x') = x)) \\
 &= \sum_{x' \in \{0, 1\}^n} \text{Prob}(x_{t-1} = x') \cdot \text{Prob}(\text{mut}(x') = x) \\
 &= \sum_{x' \in \{0, 1\}^n} 2^{-n} \cdot \text{Prob}(\text{mut}(x') = x) \\
 &= 2^{-n} \sum_{x' \in \{0, 1\}^n} \text{Prob}(\text{mut}(x') = x) \\
 &= 2^{-n} \cdot 1 = 2^{-n} \quad \square
 \end{aligned}$$

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Expected Increase in Fitness and Expected Initial Distance

$$\begin{aligned}
 &E(\text{increase in fitness}) \\
 &= \sum_{i=1}^n i \cdot \text{Prob}(\text{fitness increase} = i) \\
 &\leq \sum_{i=1}^n i \cdot \frac{1}{n} \cdot 2^{-i} \leq \frac{1}{n} \sum_{i=1}^{\infty} \frac{i}{2^i} = \frac{2}{n}
 \end{aligned}$$

$$\begin{aligned}
 E(d(x_0)) &= n - \sum_{i=1}^n i \cdot \text{Prob}(\text{LEADINGONES}(x_0) = i) \\
 &= n - \sum_{i=1}^n \frac{i}{2^{i+1}} \geq n - \frac{1}{2} \sum_{i=1}^{\infty} \frac{i}{2^i} = n - 1
 \end{aligned}$$

thus $E(T_{(1+1)} \text{EA, LEADINGONES}) \geq \frac{(n-1)n}{2} = \Omega(n^2)$
thus $E(T_{(1+1)} \text{EA, LEADINGONES}) = \Theta(n^2)$

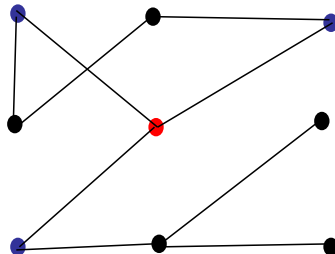
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Random Walks

Random Walks on Graphs

Given: An undirected connected graph.

- A **random walk** starts at a vertex v .
- Whenever it reaches a vertex w , it chooses in the next step a random neighbor of w .



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Result Cover Time

Theorem (Upper bound for Cover Time)

Given an undirected connected graph with n vertices and m edges, the expected number of steps until a random walk has visited all vertices is at most $2m(n-1)$.

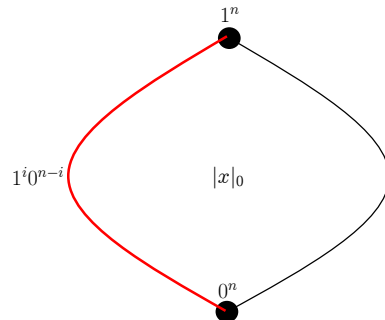
R. Aleliunas et al.: Random walks, universal traversal sequences, and the complexity of maze problems, FOCS 1979.

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Example: Plateaus

Definition

$$\text{PLATEAU}(x) := \begin{cases} |x|_0 & : x \notin \{1^i 0^{n-i}, 0 \leq i \leq n\} \\ n+1 & : x \in \{1^i 0^{n-i}, 0 \leq i < n\} \\ n+2 & : x = 1^n. \end{cases}$$



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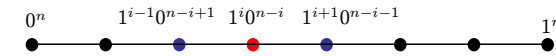
Method of Typical Runs

- Phase 1:** Given EA starts with random initialization, with probability at least $1 - p_1$, it reaches a population satisfying condition C_1 in at most T_1 steps.
- Phase 2:** Given EA starts with a population satisfying condition C_1 , with probability at least $1 - p_2$, it reaches a population satisfying condition C_2 in at most T_2 steps.
- ...
- Phase k :** Given EA starts with a population satisfying condition C_{k-1} , with probability at least $1 - p_k$, it reaches a population containing a global optimum in at most T_k steps.

This yields: $\text{Prob} \left(T_{\text{EA},f} \leq \sum_{i=1}^k T_i \right) \geq 1 - \sum_{i=1}^k p_i$

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Result: PLATEAU



Upper bound (RLS)

- Solution with fitness $\geq n+1$ in expected time $O(n \log n)$.
- Random walk on the plateau of fitness $n+1$.
- Probability $1/2$ to increase (reduce) the number of ones.
- Expected waiting time for an accepted step $\Theta(n)$.
- Optimum reached within $O(n^2)$ expected accepted steps.
- Upper bound $O(n^3)$ (same holds for $(1+1)$ -EA).

T. Jansen, I. Wegener: Evolutionary algorithms - how to cope with plateaus of constant fitness and when to reject strings of the same fitness, IEEE Trans. Evolutionary Computation 6, 58-71 (2002)

From Success Probability to Expected Optimization Time

Sometimes

“Phase 1: Given EA starts with random initialization”

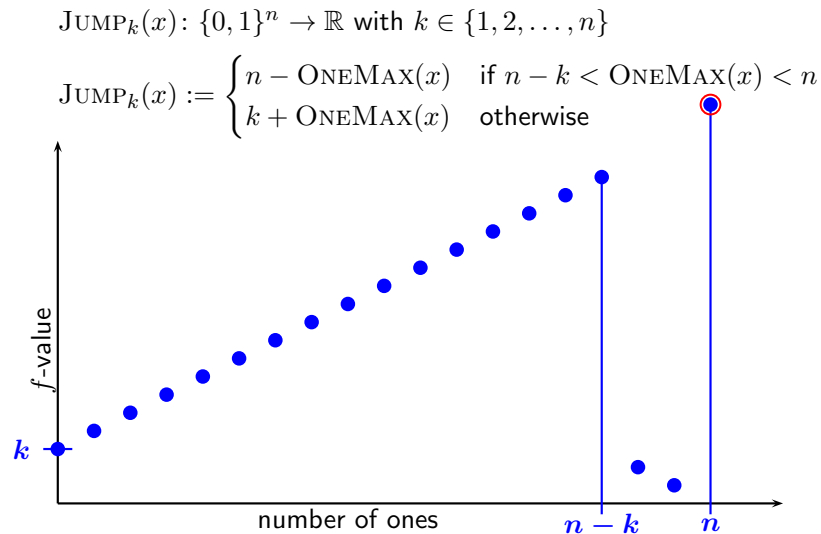
can be replaced by

“Phase 1: EA may start with an arbitrary population”

In this case, a failure in any phase can be described as a restart.

This yields: $\mathbf{E}(T_{\text{EA},f}) \leq \frac{\sum_{i=1}^k T_i}{1 - \sum_{i=1}^k p_i}$

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Theorem

Let $k = O(\log n)$, $c \in \mathbb{R}^+$ a sufficiently large constant, $\mu = n^{O(1)}$, $\mu \geq k \log^2 n$, $0 < p_c \leq 1/(ckn)$.
 $E(T_{\text{GA}(\mu, p_c)}) = O(\mu n^2 k + 2^{2k}/p_c)$

Method of Proof: Typical Run

Jansen/Wegener (2002): On the analysis of evolutionary algorithms - a proof that crossover really can help. *Algorithmica* 34(1):47-66

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$(\mu+1)$ -EA with prob. p_c for uniform crossover

- Initialization**
Choose $x_1, \dots, x_\mu \in \{0, 1\}^n$ uniformly at random.
- Selection and Variation**
With probability p_c :
 Select z_1 and z_2 independently from x_1, \dots, x_μ .
 $z := \text{uniform crossover}(z_1, z_2)$
 $y := \text{standard } 1/n \text{ bit mutation}(z)$
 Otherwise:
 Select z from x_1, \dots, x_μ .
 $y := \text{standard } 1/n \text{ bit mutation}(z)$
- Selection for Replacement**
If $f(y) \geq \min\{f(x_1), \dots, f(x_\mu)\}$
 Then Replace some x_i with min. f -value by y .
- "Stopping Criterion"**
Continue at 2.

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Notation:

$x_i[j]$ is the j -th bit of x_i

$\text{OPT} : n + k \in \{\text{JUMP}_k(x_1), \dots, \text{JUMP}_k(x_\mu)\}$

i	C_{i-1}	C_i	T_i
1	\emptyset	$\min\{\text{JUMP}_k(x_1), \dots, \text{JUMP}_k(x_\mu)\} \geq n$	$O(\mu n \log n)$
2	C_1	$\left(\forall j \in \{1, \dots, n\} : \sum_{h=1}^{\mu} (1 - x_h[j]) \leq \frac{\mu}{4k} \right) \vee \text{OPT}$	$O(\mu n^2 k)$
3	C_2	OPT	$O(2^{2k}/p_c)$

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Phase 1: Towards the Gap

Reaching some point x with $\text{JUMP}_k(x) \geq n$
is not more difficult than optimizing ONEMAX.

For $\mu = 1$, $O(n \log n)$ follows.

For larger μ , observe:

With probability at least $(1 - p_c) \cdot (1 - 1/n)^n = \Omega(1)$
a copy of a parent is produced.

Making a copy of some x_j with $\text{JUMP}_k(x_j) \geq \text{JUMP}_k(x_i)$
is not worse than choosing x_i .

This implies $O(\mu n \log n)$ as expected length.

Markov's inequality: failure probability $p_1 \leq \varepsilon$
for any constant $\varepsilon > 0$

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Zero-Bits at the First Position

Consider one generation.

Let z be the current number of zero-bits in first position.

The value of z can change by at most 1.

event A_z^+ : z changes to $z + 1$

event A_z^- : z changes to $z - 1$

Goal: Estimate $\text{Prob}(A_z^+)$ and $\text{Prob}(A_z^-)$.

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Phase 2: At the Gap

We are going to prove:

After $c' \mu n^2 k$ generations (c' const. suff. large)
with probability at most p'_2
there are at most $\mu/(4k)$ zero-bits at the first position.

This implies:

After $c' \mu n^2 k$ generations (c' const. suff. large)
there are at most $\mu/(4k)$ zero-bits at any position
with probability at most $p_2 := n \cdot p'_2$.

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A Closer Look at A_z^+

“Smaller/Simpler” Events:

event	description	probability
B_z	do crossover	p_c
C_z	at selection for replacement, select x with 1 at first position	$(\mu - z)/\mu$
D_z	at selection for reproduction, select parent with 0 at first position	z/μ
E_z	no mutation at first position	$1 - \frac{1}{n}$
$F_{z,i}^+$	out of $k - 1$ 0-bits i mutate and out of $n - k$ 1-bits i mutate	$\binom{k-1}{i} \binom{n-k}{i} \left(\frac{1}{n}\right)^{2i} \left(1 - \frac{1}{n}\right)^{n-2i}$
$G_{z,i}^+$	out of k 0-bits i mutate and out of $n - k - 1$ 1-bits $i - 1$ mutate	$\binom{k}{i} \binom{n-k-1}{i-1} \left(\frac{1}{n}\right)^{2i-1} \left(1 - \frac{1}{n}\right)^{n-2}$

Observe:

$$A_z^+ \subseteq B_z \cup \left(\overline{B_z} \cap C_z \cap \left[\left(D_z \cap E_z \cap \bigcup_{i=0}^{k-1} F_{z,i}^+ \right) \cup \left(\overline{D_z} \cap \overline{E_z} \cap \bigcup_{i=1}^k G_{z,i}^+ \right) \right] \right)$$

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A Still Closer Look at A_z^+

Using

$$A_z^+ \subseteq B_z \cup \left(\overline{B_z} \cap C_z \cap \left[\left(D_z \cap E_z \cap \bigcup_{i=0}^{k-1} F_{z,i}^+ \right) \cup \left(\overline{D_z} \cap \overline{E_z} \cap \bigcup_{i=1}^k G_{z,i}^+ \right) \right] \right)$$

together with

$$\text{Prob}(B_z) = p_c$$

$$\text{Prob}(C_z) = \frac{\mu - z}{\mu}$$

$$\text{Prob}(D_z) = \frac{z}{\mu}$$

$$\text{Prob}(E_z) = 1 - \frac{1}{n}$$

$$\text{Prob}(F_{z,i}^+) = \binom{k-1}{i} \binom{n-k}{i} \left(\frac{1}{n}\right)^{2i} \left(1 - \frac{1}{n}\right)^{n-2i}$$

$$\text{Prob}(G_{z,i}^+) = \binom{k}{i} \binom{n-k-1}{i-1} \left(\frac{1}{n}\right)^{2i-1} \left(1 - \frac{1}{n}\right)^{n-2i}$$

yields some bound on $\text{Prob}(A_z^+)$.

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A Still Closer Look at A_z^-

Using

$$A_z^- \supseteq \overline{B_z} \cap C_z \cap \left[\left(D_z \cap \overline{E_z} \cap \bigcup_{i=1}^k F_{z,i}^- \right) \cup \left(\overline{D_z} \cap E_z \cap \bigcup_{i=0}^k G_{z,i}^- \right) \right]$$

together with the known probabilities

yields again some bound.

Instead of considering the two bounds directly,
we consider their difference:

If z is large, say $z \geq \frac{\mu}{8k}$:

$$\text{Prob}(A_z^-) - \text{Prob}(A_z^+) = \Omega\left(\frac{1}{nk}\right)$$

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A Closer Look at A_z^-

“Smaller/Simpler” Events:

event	description	probability
B_z	do crossover	p_c
C_z	at selection for replacement, select x with 1 at first position	$(\mu - z)/\mu$
D_z	at selection for reproduction, select parent with 0 at first position	z/μ
E_z	no mutation at first position	$1 - \frac{1}{n}$
$F_{z,i}^-$	out of $k-1$ 0-bits $i-1$ mutate and out of $n-k$ 1-bits i mutate	$\binom{k-1}{i-1} \binom{n-k}{i} \left(\frac{1}{n}\right)^{2i-1} \left(1 - \frac{1}{n}\right)^{n-2}$
$G_{z,i}^-$	out of k 0-bits i mutate and out of $n-k-1$ 1-bits i mutate	$\binom{k}{i} \binom{n-k-1}{i} \left(\frac{1}{n}\right)^{2i} \left(1 - \frac{1}{n}\right)^{n-2}$

Observe:

$$A_z^- \supseteq \overline{B_z} \cap C_z \cap \left[\left(D_z \cap \overline{E_z} \cap \bigcup_{i=1}^k F_{z,i}^- \right) \cup \left(\overline{D_z} \cap E_z \cap \bigcup_{i=0}^k G_{z,i}^- \right) \right]$$

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Bias Towards 1-Bits

We know: $z \geq \frac{\mu}{8k} \Rightarrow \text{Prob}(A_z^-) - \text{Prob}(A_z^+) = \Omega\left(\frac{1}{nk}\right)$

Consider $c^* \mu n^2 k$ generations; c^* sufficiently large constant

$$\mathbb{E}(\text{difference in 0-bits}) = \Omega\left(\frac{n^2 k}{nk}\right) = \Omega(nk)$$

Having c^* sufficiently large implies $< \mu/(4k)$ 0-bits at the end of the phase.

Really?

Only if $z \geq \mu/(8k)$ holds all the time!

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Coping with Our Assumption

As long as $z \geq \mu/(8k)$ holds, things work out nicely.

Consider last point of time, when $z < \mu/(8k)$ holds in the c^*n^2k generations.

Case 1: at most $\mu/(8k)$ generations left

number of 0-bits $< \mu/(8k) + \mu/(8k) = \mu/(4k)$
no problem

Case 2: more than $\mu/(8k)$ generations left

Observation: $\mu/(8k) = \Omega(\log^2 n)$

For $\Omega(\log^2 n)$ generations, our assumption holds.

Apply Chernoff's bound for these generations.

Yields $p'_2 = e^{-\Omega(\log^2 n)}$.

Together: $p_2 = n \cdot p'_2 = e^{-\Omega(\log^2 n) + \ln n} = e^{-\Omega(\log^2 n)}$

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Concluding Phase 3

We have

Prob (find optimum in current generation) $= \Omega(p_c \cdot 2^{-2k})$

Prob (find optimum in $c_3 2^{2k}/p_c$ generations) $\geq 1 - \varepsilon(c_3)$

failure probability $p_3 \leq \varepsilon'$ for any constant $\varepsilon' > 0$

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Phase 3: Finding the Optimum

In the beginning, we have at most $\mu/(4k)$ 0-bits at each position.

In the same way as for Phase 2, we make sure that we always have at most $\mu/(2k)$ 0-bits at each position.

Prob (find optimum in current generation)
 \geq Prob(crossover and select two parents without common 0-bit and create 1^n with uniform crossover and no mutation)

Prob (crossover) $= p_c$

Prob (create 1^n with uniform crossover) $= (1/2)^{2k}$

Prob (no mutation) $= (1 - 1/n)^n$

Prob (select two parent without common 0-bit) $\leq k \cdot \frac{\mu/(2k)}{\mu} = \frac{1}{2}$

Together:

Prob (find optimum in current generation) $= \Omega(p_c \cdot 2^{-2k})$

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Concluding the Proof

Length of the three phases:

$O(\mu n \log n) + O(\mu n^2 k) + O(2^{2k}/p_c) = O(\mu n^2 k + 2^{2k}/p_c)$

Sum of Failure Probabilities: $\varepsilon + e^{-\Omega(\log^2 n)} + \varepsilon' \leq \varepsilon^* < 1$

$E(T_{\text{GA}(\mu, p_c)}) = O(\mu n^2 k + 2^{2k}/p_c)$

□

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Black Box Optimization

Setting

- Given two finite spaces S and R .
- Find for a given function $f: S \rightarrow R$ an optimal solution.
- Count number of fitness evaluations.
- No search point is evaluated more than once.

Definition (Black Box Algorithm)

An algorithm A is called **black box algorithm** if it finds for each $f: S \rightarrow R$ an optimal solution after a finite number of fitness evaluations.

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What Follows from NFL?

Implications

- Considering all functions, each black box algorithm has the same performance.
- Considering all functions, each algorithm is as good as **random search**.
- **Hill climbing** is as good as **Hill descending**.

Questions

- Is the result surprising? **Perhaps**
- Is it interesting? **No!!!**

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NFL

Theorem (NFL)

Given two finite spaces R and S and two arbitrary black box algorithms A and A' . The average number of fitness evaluations among all functions $f: S \rightarrow R$ is the same for A and A' .

D.H. Wolpert, W. G. Macready: No Free Lunch Theorems for Optimization, IEEE Transactions on Evolutionary Computation, 1997.

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What Does Not Follow from NFL?

Drawbacks

- **No one wants to consider all functions!!!**
- More realistic is to consider a class of functions or problems.
- NFL Theorem does not hold in this case.
- **NFL Theorem useless for understanding realistic szenarios.**

Implication

- **Restrict** considerations to class of functions/problems.
- Are there general results for such cases where NFL does not hold?
- \Rightarrow **black box complexity**.

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Motivation for Complexity Theory

If our evolutionary algorithm performs **poorly** is it **our fault** or is the **problem intrinsically hard**?

Example $\text{NEEDLE}(x) := \prod_{i=1}^n x[i]$

Such questions are answered by complexity theory.

Typically one concentrates on computational complexity with respect to run time.

Is this really fair when looking at evolutionary algorithms?

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Notation

Let $\mathcal{F} \subseteq \{f: S \rightarrow W\}$ be a class of functions, A a black box algorithm for \mathcal{F} , x_t the t -th search point sampled by A .

optimization time of A on $f \in \mathcal{F}$:

$$T_{A,f} = \min \{t \mid f(x_t) = \max\{f(x) \in S\}\}$$

worst case expected optimization time of A on \mathcal{F} :

$$T_{A,\mathcal{F}} = \max \{E(T_{A,f}) \mid f \in \mathcal{F}\}$$

black box complexity of \mathcal{F} :

$$B_{\mathcal{F}} = \min \{T_{A,\mathcal{F}} \mid A \text{ is black box algorithm for } \mathcal{F}\}$$

Droste/Jansen/Wegener (2006): Upper and lower bounds for randomized search heuristics in black-box optimization. *Theory of Computing Systems* 39(4):525–544

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Black Box Optimization

When talking about NFL we have realized classical algorithms and black box algorithms work in different scenarios.

classical algorithms	black box algorithms
problem class known	problem class known
problem instance known	problem instance unknown

This different optimization scenario requires a different complexity theory.

We consider **Black Box Complexity**.

We hope for general lower bounds for all black box algorithms.

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Comparison With Computational Complexity

$$\mathcal{F} := \left\{ f: \{0,1\}^n \rightarrow \mathbb{R} \mid f(x) = w_0 + \sum_{i=1}^n w_i x_i + \sum_{1 \leq i < j \leq n} w_{i,j} x_i x_j \right\}$$

with $w_i, w_{i,j} \in \mathbb{R}$

known: Optimization of \mathcal{F} is NP-hard since MAX-2-SAT is contained in \mathcal{F} .

Theorem: $B_{\mathcal{F}} = O(n^2)$

Proof

$w_0 = f(0^n)$ (**1 search point**)

$w_i = f(0^{i-1}10^{n-i}) - w_0$ (**n search points**)

$w_{i,j} = f(0^{i-1}10^{j-i-1}10^{n-j}) - w_i - w_j - w_0$ (**$\binom{n}{2}$ search points**)

Compute optimal solution x^* without access to the oracle.

$f(x^*)$ (**1 search point**)

together: $\binom{n}{2} + n + 2 = O(n^2)$ search points

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Observation: $\forall \mathcal{F}: B_{\mathcal{F}} \leq |\mathcal{F}|$

Consequence: $B_f = 1$ for any f — pointless

Can we still have meaningful results for our example functions?

Evolutionary algorithms are often symmetric with respect to 0s and 1s.

Definition: For $f: \{0, 1\}^n \rightarrow \mathbb{R}$, we define $f^* := \{f_a \mid a \in \{0, 1\}^n\}$ where $f_a(x) := f(a \oplus x)$.

Clearly, such EAs perform equal on all $f' \in f^*$.

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very powerful general tool for lower bounds known

Theorem (Yao's Minimax Principle)

For all distributions p over \mathcal{I} and all distributions q over \mathcal{A} :
 $\min_{\mathcal{A}} \mathbb{E}(T_{A, I_p}) \leq \max_{\mathcal{I}} \mathbb{E}(T_{A_q, I})$

in words:

We get a lower bound for the worst-case performance of a randomized algorithm by proving a lower bound on the worst-case performance of an optimal deterministic algorithm for an arbitrary probability distribution over the inputs.

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Theorem

For any $\mathcal{F} \subseteq \{f: \{0, 1\}^n \rightarrow \mathbb{R}\}$, $B_{\mathcal{F}} \leq 2^{n-1} + 1/2$ holds.

Proof

Consider pure random search without re-sampling of search points. For each step t , $\text{Prob}(\text{find global optimum}) \geq 2^{-n}$.

$$B_{\mathcal{F}} \leq \sum_{i=1}^{2^n} i \cdot 2^{-n} = \frac{2^n(2^n+1)}{2^{n+1}} = 2^{n-1} + \frac{1}{2}$$

□

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Theorem

$$B_{\text{NEEDLE}^*} = 2^{n-1} + 1/2$$

Proof by application of Yao's Minimax Principle

The upper bound coincides with the general upper bound.

We consider each NEEDLE_a as possible input.

We choose the uniform distribution.

Deterministic algorithms sample the search space in a pre-defined order without re-sampling.

Since the position of the unique global optimum is chosen uniformly at random,

we have $\text{Prob}(T = t) = 2^{-n}$ for all $t \in \{1, \dots, 2^n\}$.

This implies $\mathbb{E}(T) = \sum_{i=1}^{2^n} i \cdot 2^{-n} = \frac{2^n(2^n+1)}{2^{n+1}} = 2^{n-1} + \frac{1}{2}$.

□

Remark We already knew this from NFL.

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Theorem

$$B_{\text{ONEMAX}^*} = \Omega(n/\log n)$$

Proof by application of Yao's Minimax Principle:

We choose the uniform distribution.

A deterministic algorithm is a tree with at least 2^n nodes: otherwise at least one $f \in \text{ONEMAX}^*$ cannot be optimized.

The degree of the nodes is bounded by $n + 1$: this is the number of different function values.

Therefore, the average depth of the tree is bounded below by

$$(\log_{n+1} 2^n) - 1 = \frac{n}{\log_2(n+1)} = \Omega(n/\log n). \quad \square$$

Remark: $B_{\text{ONEMAX}^*} = O(n)$ is easy to see.

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class of unimodal functions:

$$\mathcal{U} := \{f: \{0, 1\}^n \rightarrow \mathbb{R} \mid f \text{ unimodal}\}$$

What is $B_{\mathcal{U}}$?

We want to find a **lower bound** on $B_{\mathcal{U}}$.

Remember: For any point not optimal under a unimodal function, there exists a **path** to the global optimum

Definition: l points p_1, p_2, \dots, p_l with $H(p_i, p_{i+1}) = 1$ for all $1 \leq i < l$ form a **path of length l** .

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Consider $f: \{0, 1\}^n \rightarrow \mathbb{R}$.

We call $x \in \{0, 1\}^n$ a **local maximum** of f , iff for all $x' \in \{0, 1\}^n$ with $H(x, x') = 1$ $f(x) \geq f(x')$ holds.

We call f **unimodal**, iff f has exactly one local optimum.

We call f **weakly unimodal**, iff all local optima are global optima, too.

Observation: (Weakly) Unimodal functions can be optimized by hill-climbers.

Does this mean unimodal functions are easy to optimize?

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Consider the following functions:

$P := (p_1, p_2, \dots, p_{l(n)})$ with $p_1 = 1^n$ is a path — **not necessarily a simple path**.

$$f_P(x) := \begin{cases} n + i & \text{if } x = p_i \text{ and } x \neq p_j \text{ for all } j > i, \\ \text{ONEMAX}(x) & \text{if } x \notin P \end{cases}$$

Observation: f_P is unimodal.

$$\mathcal{P}_{l(n)} := \{f_P \mid P \text{ has length } l(n)\}$$

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Construct P with length $l(n)$ randomly:

1. $p_1 := 1^n; i := 2$
2. While $i \leq l(n)$ do
3. Choose $p_i \in \{x \mid H(x, p_{i-1}) = 1\}$ uniformly at random.
4. $i := i + 1$

For each path P with length $l(n)$,
we can calculate the probability to construct P randomly this way.

Remark: Paths P constructed this way are likely to contain circles.

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We need to prove that
an **optimal deterministic** algorithm
needs on average **more than 2^{n^δ} steps**
to find a global optimum.

We strengthen the position of the deterministic algorithm by

- ① letting it know which functions have probability 0.
- ② giving away for free the knowledge about any p_i with $f(p_i) \leq f(p_j)$ once p_j is sampled,
- ③ giving away for free the knowledge about p_{j+1}, \dots, p_{j+n} if p_j is the current known best path point and some point not on the path is sampled,
- ④ giving away for free the knowledge about $p_{l(n)}$ (the global optimum) once p_{j+n} is sampled while p_j is the current known best path point.

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Theorem: $\forall \delta$ with $0 < \delta < 1$ constant: $B_{\mathcal{U}} > 2^{n^\delta}$.

For a proof, we want to apply **Yao's Minimax Principle**.

We define a probability distribution in the following way:

$\delta < \varepsilon < 1$ constant; $l(n) := 2^{n^\varepsilon}$

For all $f \in \mathcal{U}$ we define

$$\text{Prob}(f) := \begin{cases} p & \text{if } f \in \mathcal{P}_{l(n)} \text{ and } P \text{ is constructed with prob. } p, \\ 0 & \text{otherwise.} \end{cases}$$

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Omit all circles from P .
The remaining length $l'(n)$ is called the **true length of P** .

What lower bound can be proven this way?

at best: $(l'(n) - n + 1)/n$

Observation: We need a good lower bound on $l'(n)$.

How likely is it to return to old path points?

alternatively: What is the probability distribution for the Hamming distance points on the path?

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Lemma

$\forall \beta > 0$ constant: $\exists \alpha(\beta) > 0$ constant: $\forall i \leq l(n) - \beta n$:
 $\forall j \geq \beta n$: $\text{Prob}(H(p_i, p_{i+j}) \leq \alpha(\beta)n) = 2^{-\Omega(n)}$

Proof: Due to symmetry:

Considering $i = 1$ and some $j \geq \beta n$ suffices.

$H_t := H(p_1, p_t)$

We want to prove: $\text{Prob}(H_j \leq \alpha(\beta)n) = 2^{-\Omega(n)}$

We choose $\alpha(\beta) := \min\{1/50, \beta/5\}$.

Due to the random path construction:

- $H_{t+1} \in \{H_t - 1, H_t + 1\}$
- $\text{Prob}(H_{t+1} = H_t + 1) = 1 - H_t/n$
- $\text{Prob}(H_{t+1} = H_t - 1) = H_t/n$

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Case 2: $H_t < 2\gamma n$

Clearly, $H_i < 3\gamma n$ for all $i \in \{t, \dots, j\}$.

Therefore, $\text{Prob}(H_i = H_{i-1} + 1) \geq 1 - 3\gamma \geq 7/10$,

$\text{Prob}(H_i = H_{i-1} - 1) \leq 3/10$.

Define independent random variable $S_t, S_{t+1}, \dots, S_j \in \{0, 1\}$ with $\text{Prob}(S_k = 1) = 7/10$.

Define $S := \sum_{k=t}^j S_k$.

Observation: $\text{Prob}(S \geq (3/5)\gamma n) \leq \text{Prob}(H_j \geq (1/5)\gamma n)$

Since

- 1 $H_t \geq 0$
- 2 $\text{Prob}(H_i = H_{i-1} + 1) \geq \text{Prob}(S_i = 1)$
- 3 $\geq (3/5)\gamma n$ increasing steps $\Rightarrow \leq (2/5)\gamma n$ decreasing steps
- 4 $H_j \geq (3/5)\gamma n - (2/5)\gamma n$

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Define $\gamma := \min\{1/10, j/n\}$.

Observations:

- $\gamma \leq 1/10$
- $\gamma \geq 5\alpha(\beta)$
- γ bounded below and above by positive constants

Consider the last γn steps towards p_j .

Let t be the first of these steps.

Note: $(\gamma \leq j/n) \Rightarrow (\gamma n \leq j)$

Case 1: $H_t \geq 2\gamma n$

Clearly, $H_j \geq \underbrace{2\gamma n}_{\text{in the beginning}} - \underbrace{\gamma n}_{\text{number of steps}} = \gamma n > \alpha(\beta)n$.

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We have γn independent random variable $S_t, S_{t+1}, \dots, S_j \in \{0, 1\}$

with $\text{Prob}(S_k = 1) = 7/10$ and $S := \sum_{k=t}^j S_k$.

Apply Chernoff Bounds:

$E(S) = (7/10)\gamma n$

$\text{Prob}(S < \frac{3}{5}\gamma n)$
 $= \text{Prob}(S < (1 - \frac{1}{7}) \frac{7}{10}\gamma n)$
 $< e^{-(7/10)\gamma n(1/7)^2/2} = e^{-(1/140)\gamma n} = 2^{-\Omega(n)}$

□

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True Path Length

Lemma with $\beta = 1$ yields:

$$\text{Prob}(\text{return to path after } n \text{ steps}) = 2^{-\Omega(n)}$$

$$\begin{aligned} &\text{Prob}(\text{return to path after } \geq n \text{ steps happens anywhere}) \\ &= 2^{n^\varepsilon} \cdot 2^{-\Omega(n)} = 2^{-\Omega(n)} \end{aligned}$$

$$\text{Prob}(l'(n) \geq l(n)/n) = 1 - 2^{-\Omega(n)}$$

We can prove **at best** lower bound of $\frac{l'(n)-n+1}{n} > \frac{l(n)}{n^2} - 1 > 2^{n^\delta}$.

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Later steps

$$N \neq \emptyset$$

Partition N :

$$N_{\text{far}} := \{y \in N \mid H(y, p_i) \geq \alpha(1/2)n\}$$

$$N_{\text{near}} := N \setminus N_{\text{far}}$$

Case 1: $N_{\text{near}} = \emptyset$

Consider random path construction starting in p_i .

A : path hits x

E : path hits no point in N_{far}

Clearly, optimal deterministic algorithm avoid N_{far} .

Thus, we are interested in $\text{Prob}(A \mid E)$

$$= \frac{\text{Prob}(A \cap E)}{\text{Prob}(E)} \leq \frac{\text{Prob}(A)}{\text{Prob}(E)}.$$

Clearly, $\text{Prob}(E) = 1 - 2^{-\Omega(n)}$.

Thus, $\text{Prob}(A \mid E) \leq (1 + 2^{-\Omega(n)}) \text{Prob}(A) = 2^{-\Omega(n)}$.

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An Optimal Deterministic Algorithm

Let N denote the points known not to belong to P .

Let p_i denote the best currently known point on the path.

Initially, $N = \emptyset$, $i \geq 1$.

Algorithm decides to sample x as next point.

Case 1: $H(p_i, x) \leq \alpha(1)n$

$$\text{Prob}(x = p_j \text{ with } j \geq n) = 2^{-\Omega(n)}$$

Case 2: $H(p_i, x) > \alpha(1)n$

Consider random path construction starting in p_i .

Similar to Lemma:

$$\text{Prob}(\text{hit } x) = 2^{-\Omega(n)}$$

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Later Steps With Close Known Points

Case 2: $N_{\text{near}} \neq \emptyset$

Knowing points near by can increase $\text{Prob}(A)$.

Ignore the first $n/2$ steps of path construction; consider $p_{i+n/2}$.

$$\text{Prob}(N_{\text{near}} = \emptyset \text{ now}) = 1 - 2^{-\Omega(n)}$$

Repeat Case 1.

◻

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... and that was it from us for today.

There is more,
but you have a good idea of what can be done.

Reminder — What we have just seen:

- analysis of the expected optimization time of some evolutionary algorithms by means of
 - fitness-based partitions
 - Markov's inequality and Chernoff bounds
 - coupon collector's theorem
 - expected multiplicative distance decrease
 - drift analysis
 - random walks and cover times
 - typical runs
 - example functions
- general limitations for evolutionary algorithms by means of
 - NFL
 - black box complexity

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Are there just these methods and results for toy examples?
Is there nothing really cool, interesting, and useful?

By these and other methods there are results for evolutionary algorithms for

- “real” combinatorial optimization problems
 - Euler circuits, Ising model, longest common subsequences
 - maximum cliques, maximum matchings, minimum spanning trees
 - shortest paths, sorting, partition
- “advanced” evolutionary algorithms
 - coevolutionary algorithms, memetic algorithms
 - with crossover, different (offspring) population sizes, problem-specific variation operators
- other randomized search heuristics
 - ant colony optimization
 - artificial immune systems
 - estimation of distribution algorithms

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Tutorials Today all in this room

- 10:40–12:30 **Theory of Randomized Search Heuristics**
Carsten Witt
- 14:00–15:50 **Theory of Swarm Intelligence**
Dirk Sudholt
- 16:10–18:00 **Drift Analysis**
Benjamin Doerr

Books

- Anne Auger, Benjamin Doerr (Eds.)
Theory of Randomized Search Heuristics: Foundations and Recent Developments.
World Scientific. To appear
- Frank Neumann, Carsten Witt
Bioinspired Computation in Combinatorial Optimization: Algorithms and Their Computational Complexity.
Springer, 2010

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